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2	Relationship between structure of macrobenthic assemblages and
3	environmental variables in shallow sublittoral soft bottoms
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5	Puri Veiga <sup>*1,2</sup> , Waldo Redondo <sup>3</sup> , Isabel Sousa-Pinto <sup>1,2</sup> , Marcos Rubal <sup>1,2</sup>
6	
7	<sup>1</sup> Laboratory of Coastal Biodiversity, Interdisciplinary Centre of Marine and
8	Environmental Research (CIIMAR/CIMAR), University of Porto, Rua dos Bragas 289,
9	P 4050-123 Porto, Portugal
10	<sup>2</sup> Department of Biology, Faculty of Sciences, University of Porto, Rua do Campo
11	Alegre s/n 4150-181 Porto, Portugal
12	<sup>3</sup> Instituto de Investigaciones Marinas, Rúa de Eduardo Cabello, 6, E-36208 Vigo
13	(Pontevedra), Spain
14	*Corresponding-author. e-mail: <u>puri.sanchez@fc.up.pt</u>
15	Tel.: +351 223401800; fax: +351 223390608
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#### 18 ABSTRACT

We establish baseline knowledge of abundance, diversity and multivariate structure of 19 20 macrobenthos from shallow sublitoral soft bottoms in the North Portuguese coast and 21 elucidate main environmental factors that shape their spatial patterns. In this area 22 distribution of soft bottoms is patchy, surrounded by boulders and rocky substrates. This particular landscape and the lack of significant antropogenic disturbances are values for 23 24 the conservation of this habitat. Sediment and physicochemical properties of the water column were studied to provide models for each studied macrobenthic variable. Our 25 26 models highlighted that most of variation (59%-72%) in macrobenthic spatial patterns 27 was explained by the studied environmental variables. Sedimentary variables were more relevant that those of the water column. Therefore, disturbances affecting sedimentary 28 environment could cause dramatic changes in macrobenthic assemblages because of the 29 limited availability of soft bottoms in the area. In this way, results are useful to adopt 30 right management and conservation strategies. 31

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Keywords: Benthic ecology; Sediments; Macrobenthos; Community composition;
Environmental variables; Coastal waters; Shallow sublittoral; Northeast Atlantic

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#### 37 **1. Introduction**

38 Coastal ecosystems provide valuable goods and services to humans but anthropogenic 39 use has also altered the oceans through direct and indirect means (Halpern et al., 2008). 40 Particularly, in recent decades, worldwide marine ecosystems are suffering the synergistic effects of multiple stressors derived from anthropogenic activities such as 41 42 overfishing, invasive species or pollution (Claudet and Fraschetti, 2010). These 43 stressors act as major drivers of ecosystems altering the structure and functioning of their assemblages with consequences to human well-being (Worm et al., 2006). In this 44 45 scenario, there is an imperative need for adopting management and conservation 46 strategies in marine systems that will be crucial for the sustainable use of resources (Desroy et al., 2002; Claudet and Fraschetti, 2010). However, the major constrains to 47 implement conservation strategies in marine ecosystems are the general lack of baseline 48 data prior to impacts and substantial gaps in the current knowledge of natural patterns of 49 variability of their assemblages, which are intrinsically variable (Claudet and Fraschetti, 50 2010; Schückel et al., 2015). 51

Soft bottom macrobenthos plays an important role in marine ecosystem processes such 52 as nutrient cycling, pollutant metabolism or secondary production (Snelgrove, 1998; 53 54 Pratt et al., 2014). Most of macrobenthic species display a sedentary lifestyle, intermediate trophic level positions, relatively long life-span and varying responses to 55 56 changes in environmental stress that make macrobenthos an effective and useful indicator for the assessment of coastal system quality (Dauvin, 2007). Over the past few 57 decades, macrobenthos has been a key element of many monitoring programmes; in this 58 59 way, upgrading our knowledge about its biodiversity is useful, particularly in marine soft-bottoms (Ellingsen, 2002; Veiga et al., 2016). Although soft-bottoms are the largest 60 ecosystem on Earth in terms of area coverage, only a small percentage of their 61

62 macrobenthos has been studied and most of its species are still undescribed (Snelgrove, 1998). Within soft-bottom ecosystems, sublittoral macrobenthic assemblages have been 63 less studied that those from the intertidal and remaining largely unknown (Desroy et al., 64 2002; Schückel et al., 2015). Spatial distribution of these assemblages is heterogeneous 65 (Mann and Lazier, 2006) and sediment features (e.g. grain size, organic matter content 66 and food availability) have been identified as responsible for spatial patterns of 67 macrobenthos (Ellingsen 2002; Van Hoey et al., 2004; Hily et al., 2008; Ramey and 68 Bodnar, 2008). Moreover, at greater spatial scales, physicochemical characteristics of 69 the water column and hydrodynamics, seem to control directly or indirectly abundance 70 and distribution of macrobenthos by influencing food availability, bottom-water 71 oxygenation and larval dispersion (Dauvin et al., 2004; Blanchet et al., 2005; Schückel 72 et al., 2015) 73

Spatial models such as multiple regression or canonical correspondence analyses have 74 revealed that the percentage of the variation in assemblage structure from soft-bottom 75 habitats explained by environmental factors is very variable (i.e. between 10% and 90 76 %) (e.g. Veiga et al., 2009; Olsson et al., 2013; Schückel et al., 2015). These previous 77 studies were done in subtidal areas where soft bottoms are the dominant habitat and 78 79 cover wide extensions. However, in some regions like north Portuguese coast, soft bottoms at shallow sublittoral are restricted to patches surrounded by large extensions of 80 boulders and rocky substrates, the latter being the predominant habitat (Rodríguez et al., 81 2011). Moreover, on the one hand, the north Portuguese coast is still an area 82 characterized by relatively low levels of anthropogenic pressure, deserving attention for 83 its conservation. Previous investigations showed that concentrations of nutrients, PAHs 84 and trace metals were near background values (e.g. Reis, 2012; Reis et al., 2014; Rubal 85 et al., 2014), indicating that North Portuguese coast is not subjected to severe 86

87	eutrophication or pollution by industrialization and urbanization of the surrounding
88	areas. On the other hand, benthic studies done in this area have been focused on
89	intertidal assemblages from rocky shores (e.g. Araújo et al., 2006; Rubal et al., 2011;
90	Veiga et al., 2013) and soft bottoms (Veiga et al., 2014). However, there is a gap in
91	knowledge about the structure of assemblages from subtidal soft-bottoms. As proof of
92	this, new species of macro- and meiobenthos have been recently described from shallow
93	subtidal sediments of the North Portuguese coast (Esquete et al., 2015; Rubal et al., In
94	press), indicating that this system may be also of high value for conservation.
95	The study of spatial patterns in macrobenthic assemblages in this area will let us
96	establish baseline knowledge, mandatory to detect future potential changes in species
97	distribution and helpful for monitoring and management issues(Desroy et al., 2002;
98	Claudet and Fraschetti, 2010; Dutertre et al., 2013; Schückel et al., 2015). Moreover,
99	elucidating main natural environmental factors that shape spatial patterns of
100	macrobenthic assemblages from subtidal soft-bottom will help to discriminate between
101	natural and anthropogenic changes (Glockzin and Zettler, 2008; Dutertre et al., 2013).
102	Therefore, the aims of this study were to determine the natural environmental variables
103	that shape the structure of macrobenthic assemblages in shallow sublittoral soft bottoms
104	in the North Portuguese coast and providing baseline information for assessing the
105	quality of this system in the future, which will be crucial for adopting right management
106	and conservation strategies. To achieve these aims first, spatial patterns of sediment
107	features, physicochemical properties of the water column and macrobenthic
108	assemblages were described. Then, the relationship between spatial distribution patterns
109	of macrobenthos and those of environmental factors were investigated using
110	multivariate statistical approaches. This will allow identifying useful predictor variables

- and generating simple models to explain natural spatial variability in macrobenthicassemblages.
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#### 114 **2. Material and methods**

- 115 *2.1 Study area*
- 116 The study was carried out on shallow subtidal soft bottoms in the North of Portugal,
- encompassing over 22 km of coast between 41°51′10.01″N; 8°52′54.00″W and
- $41^{\circ}39'39.72''$ N;  $8^{\circ}50'24.42''$ W (Table 1 and Figure S1). This subtidal area is
- 119 predominantly covered by rocky shores that constituting the 69%, whereas soft bottoms
- are the second most abundant habitat (21%) followed by boulders (10%) (Rodríguez et
- al., 2011). The coast in this area is north-to-south oriented, exposed to prevailing
- 122 northwest oceanic swell. Moreover, this coastal area is subjected to the influence of
- 123 river plumes, being Miño and Lima the most important rivers regarding flow and to
- upwelling events (Lemos and Pires, 2004).
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## 126 2.2 Sampling design

Sampling was conducted in May 2012 at four shallow subtidal soft bottom localities 127 128 (Table 1, Figure S1). A two-factor sampling design was used to assess the spatial 129 patterns of macrobenthic assemblages and their relationship with sedimentary and water 130 column environment. The largest spatial scale was that of locality, which included four levels: Moledo, Âncora, Gelfa and Lima, spaced kms from one another. At each 131 132 locality, three sites, approximately 100s of ms apart, were randomly established within 133 each soft bottom patch. Localities and sites were selected considering the availability of soft bottoms (Figure S1) based on a previous work that had characterised main habitats 134

135 of the study area including bathymetric and geomorphologic analyses (Rodríguez et al., 2011). Sediment samples within each site were randomly collected, about 10s of ms 136 apart, using a Van Veen grab (sampling surface of  $0.12 \text{ m}^2$ ) to a mean depth of 12 m 137 (between 9.5 and 15 m) (Table 1). At each site, a total of seven grabs were collected, 138 five to the study of macrobenthos and the remaining two to study the sedimentary 139 environment (i.e. grain size and organic matter). Macrobenthic samples were 140 immediately washed on board over a 0.5-mm mesh sieve. The retained macrofauna was 141 142 then preserved in 4% neutralised formaldehyde solution with Rose Bengal in labelled plastic bags until its posterior study. Samples to sedimentary study were frozen. To 143 characterise the water column environment, three independent measures of oxygen 144 concentration, salinity and temperature were obtained at each locality by means of a 145 CTD SBE25. Moreover, three independent water column samples of 250 ml were 146 147 collected at each site and locality for nutrient analyses: nitrate (NO<sub>3</sub>), phosphorus (PO<sub>4</sub>) and ammonium (NH<sub>3</sub>) as close to the bottom as possible avoiding sediment 148 149 resuspension.

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151 2.3 Sampling processing

Macrobenthos was sorted, identified to the lowest possible taxon (usually species level) 152 and counted. The organic matter content was calculated by measuring the loss of weight 153 on ignition in a furnace at 450°C for 4 h. In order to study the sediment grain size, 154 samples were dried and then sieved. The following sedimentary fractions were 155 considered: coarse gravel (> 4 mm), fine gravel (2-4 mm), very coarse sand (1-2 mm), 156 coarse sand (0.5-1 mm), medium sand (0.25-0.5 mm), fine sand (0.125-0.25 mm), very 157 fine sand (0.063-0.125 mm) and silt/clay (< 0.063 mm). Then, the median particle size 158 (Md; Bale and Kenny, 2005) and sorting coefficient of the sediment (QD $\Phi$ ; Yamanaka 159

et al., 2012) were calculated. Nutrient analyses were done directly in filtered seawater
samples by Molecular Absorption Spectrometry using a segmented flux autoanalyser
(San Plus System, Skalar). The concentrations of NO<sub>3</sub>, PO<sub>4</sub> and NH<sub>3</sub> were determined
according to Skalar methods M461-318 (EPA 353.2), M155-008R (EPA 350.1) and
M503-555R (Standard Method 450-P I) (Skalar, 2004).

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166 *2.4 Data analyses* 

Data were analysed by means of univariate and multivariate techniques to test the
hypothesis that variables of the water column and sediment beside total number of
individuals (N), taxon richness (S), Shannon's diversity index (H') and the multivariate
structure of macrobenthic assemblages will differ among localities.

Analyses of variance (ANOVA) were done to test for differences in the water column 171 variables (i.e. oxygen concentration, salinity and temperature) among localities. These 172 analyses were based on a one-way model, including locality as fixed factor with four 173 levels and three replicates. ANOVA was also used to test for differences in the nutrient 174 175 content (i.e. NO<sub>3</sub>, PO<sub>4</sub> and NH<sub>3</sub>), sediment variables (i.e. different sedimentary grain sizes and organic matter content), N, S and H' of macrobenthos among localities. These 176 177 analyses were based on a two-way model, including Locality as fixed factor with four levels and Site as random factor nested in Locality with three levels, considering five 178 179 replicates for macrobenthic data, three for nutrients and two for sedimentary variables. Cochran's C tests were previously done to check for homogeneity of variances and, 180 181 when test was significant (p < 0.05), data were transformed to remove the heterogeneity 182 of variances. When this was not possible, untransformed data were analysed and results were considered robust if significant at p < 0.01, to compensate for the increased 183 probability of type I error (Underwood, 1997). Whenever ANOVA showed significant 184

185	differences (p $< 0.05$ ), a post hoc Student-Newman-Keuls (SNK) test was done to
186	explore differences among all pairs of levels of the selected factor (i.e. locality).
187	A distance-based multivariate analysis of variance (PERMANOVA, Anderson 2001a),
188	according to the two-way aforementioned design, was used to test differences among
189	localities on the multivariate structure of macrobenthic assemblages. For each replicate,
190	values of abundance for each species were square-root transformed to downweight the
191	influence of numerically dominant taxa, which were used for the calculation of the
192	Bray–Curtis similarity matrix by permutation of residuals under a reduced model (999).
193	When PERMANOVA showed significant differences ( $p < 0.05$ ), a pair-wise
194	comparison (999 permutations) was done to explore differences among all pairs of
195	levels of the selected factor. As the number of unique permutations for the pair-wise
196	comparison was lower than 30, Monte Carlo P-values were considered (Terlizzi et al.,
197	2005).
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<ol> <li>197</li> <li>198</li> <li>199</li> <li>200</li> <li>201</li> <li>202</li> <li>203</li> <li>204</li> <li>205</li> <li>206</li> </ol>	2005). In order to test whether differences in the multivariate structure of assemblages between localities were due to varying multivariate dispersion, the PERMDISP procedure was done (Anderson, 2006). Multivariate patterns were illustrated by non-metric multidimensional scaling (nMDS) ordination of replicates for each locality. The SIMPER procedure (Clarke, 1993) was used to determine the percentage of contribution (δi%) of each taxon to the Bray-Curtis dissimilarity between assemblages sampled at each locality (δi). A taxon was considered important if its contribution to total percentage dissimilarity was ≥3%. The ratio δi/SD(δi) was used to quantify the consistency of the contribution of a particular taxon to the average dissimilarity in all
<ol> <li>197</li> <li>198</li> <li>199</li> <li>200</li> <li>201</li> <li>202</li> <li>203</li> <li>204</li> <li>205</li> <li>206</li> <li>207</li> </ol>	2005). In order to test whether differences in the multivariate structure of assemblages between localities were due to varying multivariate dispersion, the PERMDISP procedure was done (Anderson, 2006). Multivariate patterns were illustrated by non-metric multidimensional scaling (nMDS) ordination of replicates for each locality. The SIMPER procedure (Clarke, 1993) was used to determine the percentage of contribution (δi%) of each taxon to the Bray-Curtis dissimilarity between assemblages sampled at each locality (δi). A taxon was considered important if its contribution to total percentage dissimilarity was ≥3%. The ratio δi/SD(δi) was used to quantify the consistency of the contribution of a particular taxon to the average dissimilarity in all pair-wise comparisons of samples among localities. Values ≥1 indicated a high degree

209	The relationship between the uni- and multivariate macrobenthic data and the
210	environmental variables of the water column and sedimentary environment was
211	analysed using nonparametric multivariate multiple regression (McArdle and Anderson,
212	2001). Environmental variables were subjected to a stepwise forward-selection
213	procedure to develop a model of the macrobenthic data (i.e. N, S, H <sup>'</sup> and the
214	multivariate structure of assemblage). Analyses were based on Euclidean similarity
215	matrices for N, S and H <sup><math>\prime</math></sup> and Bray-Curtis similarity matrix for multivariate data that
216	were square-root transformed. P-values were done using 9999 permutations of residuals
217	under the reduced model (Anderson, 2001b). All non-parametric multivariate multiple
218	regressions were done using the computer program DISTLM (Anderson, 2002).
219	Draftsman plots were done previously to check the skewness of environmental variables
220	and data were $log(x+1)$ transformed. Some pairs of variables (i.e. temperature – oxygen
221	concentration; $PO_4 - NO_2$ and coarse sand – very coarse sand) showed strong
222	correlations (r > 0.90). As these pointed out that they are redundant variables,
223	temperature, $PO_4$ and coarse sand were removed from the analyses whereas oxygen
224	concentration, NO <sub>2</sub> and very coarse sand were maintained. Constrained ordination, a
225	distance-based redundancy analysis (dbRDA, Legendre and Anderson, 1999), was done
226	to explicitly investigate the relationship between environmental variables and
227	macrobenthic assemblages.

These analyses were also completed on sets of environmental variables that formed two natural groups, those associated with the water column environment and the associated with the sediment, to test the hypothesis that sedimentary variables will explain a greater percentage of variability in spatial patterns of macrobenthos than that explained by variables of the water column environment. Water and sediment sets were analysed separately for their relationship with the macrobenthic data (ignoring the other set), and

both sets were then subjected to a stepwise forward-selection procedure to develop amodel of macrobenthic data.

236

237 **3. Results** 

238 3.1 Environmental variables

Depth ranged from 9.5 m (Site 1 of Lima) to 15 m (Sites 1 of Âncora and Gelfa) (Table 239 1). Regarding water column environment (Table S1), results of ANOVA showed 240 significant differences among localities for oxygen concentration and temperature 241 (Table 2). Post hoc analyses indicated that oxygen concentration and temperature in 242 Gelfa were significantly higher than in the remaining localities (Fig. 1A and 1C). 243 However, salinity did not show significant differences among localities (Table 2, Fig. 244 245 1B). Moreover, results of ANOVA indicated no significant differences among localities for nutrient content (NO<sub>3</sub>, NH<sub>3</sub> and PO<sub>4</sub>) despite some variability was evident among 246 sites (Table 2; Figure 2A-C). 247 Sediments were predominantly composed by fine sand with the only exception of Miño 248 that showed coarse and medium sand (Table 1). Moreover, sediments were well sorted 249 and moderately well sorted, except on Site 3 of Âncora where they were very well 250 sorted (Table 1). Concerning sedimentary environment, results of ANOVA showed 251 significant differences among localities for organic matter and most of the sedimentary 252 grain sizes, except for coarse and fine gravel (Table 3). Post hoc analyses indicated that 253 very coarse, coarse and medium sand content in Miño were significantly higher than in 254 the remaining localities (Figure 3C-E). However, organic matter, fine, very fine sand 255 and silt/clay content were significantly lower in Miño (Figure 3F-I). 256

257

258 3.2 Macrobenthic assemblages

A total of 9669 individuals belonging to 65 taxa were identified throughout the study 259 (Table S2). Results of ANOVA showed significant differences among localities for N 260 (Table 4). Post hoc analyses indicated that N in Moledo and Âncora was significantly 261 higher than in the remaining localities (Figure 4A). However, results of ANOVA 262 263 indicated no significant differences among localities for S and H' (Table 4, Figure 4B-264 C). PERMANOVA analysis showed that the multivariate structure of macrobenthic assemblages differed significantly among localities (Table 5). Pair-wise comparisons 265 indicated significant differences between Miño and the remaining localities (Table 5). 266 267 The documented multivariate pattern was visualized as a clear separation between replicates from Miño and those of the remaining localities in the nMDS ordination 268 (Figure 5). Moreover, the PERMDISP analysis for locality (F = 1.58, p = 0.274) 269 indicated that the dispersion of replicates did not provide a significant contribution to 270 observed differences among localities. 271 SIMPER analysis identified 34 taxa as the main responsible for differences between 272 273 Miño and the remaining localities. Collectively, these taxa contributed more than 90% to the total dissimilarity, although only the contribution by nine of them was  $\geq 3\%$ 274 275 (Table 6). The contribution to percentage of dissimilarity of Pisione parapari Moreira, Quintas and Troncoso, 2000, P. remota (Southern, 1914), Polygordius appendiculatus 276 277 Fraipont, 1887, Spio decoratus Bobretzky, 1870, Nematodes, Diogenes pugilator (Roux, 1829) and Gastrosaccus spinifer (Goës, 1864) was consistent among all the pair-278 279 wise comparisons between Miño and the remaining localities (Table 6). Moreover, 280 Eurydice sp. and Fabulina fabula (Gmelin, 1791) contributed only to dissimilarity of

281 Miño with Gelfa and Lima, respectively. Noticeably, the abundance of these species

was larger in Miño compared to the remaining localities except for *S. decoratus* and *F.* 

283 *fabula*, which reported lower abundances in Miño (Table 6).

284 *3.3 Relationship between environmental variables and macrobenthic assemblages* 

285 Results of DISTLM showed that 11 environmental variables (i.e. all the studied

variables except fine gravel and very coarse sand) together explained 71.51% of the

variance in N (Table 7). The variable that explained its greatest amount of variation was

the fine sand content that alone explained the 56.30% followed by medium sand content

289 (3.65%) and depth (6.02%) (Table 7). However, after fitting these three variables, the p-

values associated with the conditional test to add silt/clay and the subsequently fitted

terms in the model were not statistically significant (p > 0.05) (Table 7). In this way,

based on forward selection results, the best model to explain N would be include only

the first three variables: fine sand, medium sand and depth that together explained the

294 65.89% of variation in N (Table 7).

Results of DISTLM also showed that all the studied variables, except medium sand and 295 296 very fine sand, together explained 60.64% of the variance in S (Table 7). The variable that explained the greatest amount of variation (24.65%) was salinity, followed by 297 silt/clay content (6.65%) and depth (4.48%) (Table 7). Moreover, NO<sub>3</sub>, NH<sub>3</sub> and organic 298 matter content added significantly to explain variation of S, as evidenced by the p-299 300 values < 0.05 (Table 7). However, after fitting these six variables, the p-values 301 associated with the conditional test to add very coarse sand and the subsequently fitted terms in the model were not statistically significant (p > 0.05) (Table 7). In this way, 302 based on forward selection results, the best model to explain S would be include the 303 304 first six variables: salinity, silt/clay, depth, NO<sub>3</sub>, NH<sub>3</sub> and organic matter content that together explained the 58.45% of variation in S (Table 7). 305

306 Results of DISTLM showed that all studied variables, except oxygen concentration, salinity and coarse gravel, together explained 59% of the variability in H<sup>2</sup> (Table 7). The 307 variable that explained the greatest amount of variation was silt/clay content that alone 308 309 explained the 46.46% (Table 7). However, after silt and clay content, the p-values associated with the conditional test to add organic matter content and the subsequently 310 fitted terms in the model were not statistically significant (p > 0.05) (Table 7). In this 311 way, based on forward selection results, the best model to explain H' would be to 312 include only silt/clay content. 313 Results of DISTLM showed that all studied variables, except fine gravel and organic 314 315 matter content, explained 64.41% of the variability in the multivariate structure of macrobenthic assemblages (Table 7). The variable that alone explained the greatest 316 amount of variation (25.61%) was very coarse sand content followed by very fine sand 317 content (8.27%) and salinity (4.99%) (Table 7). Moreover, all the remaining variables 318 added significantly to explain variation in the multivariate structure of macrobenthic 319 assemblages, as evidenced by the p-values < 0.05 (Table 7). In this way, based on 320 forward selection results, the best model to explain the multivariate structure of 321 macrobenthic assemblages would include the 11 variables (Table 7). 322 323 The first two dbRDA axes explained 66% of the fitted variation, which is about 42.2% of the total variation in the structure of the macrobenthic assemblages (Figure 6). All 324 dbRDA axes together explain 100% of the fitted variation and 64% of the total 325 variation. 326 The results of the multivariate analyses based on two sets of variables (sediment and 327 328 water column) are shown in Table 8. The set of sedimentary variables showed the greatest descriptive power, which explained 70% of the variation in the macrobenthic 329

abundance and more than 54% in the diversity (S and H') and the multivariate structure

of macrobenthic assemblages (Table 8). Once the sedimentary variables were fitted, the set of water column variables only added another 7% to the explained variation in the multivariate structure of macrobenthic assemblages although this was statistically significant. However, in terms of abundance and diversity, the set of water column variables appeared to be redundant in the model (p > 0.05) (Table 8).

336

## 337 Discussion

338 Coastal ecosystems face strong anthropogenic pressures but are also naturally variable

339 (Crossland et al., 2005; Halpern et al., 2015). This variability makes difficult to

340 determine whether structural changes in benthic assemblages are due to natural or

anthropogenic disturbances, unless the latter are severe (Hardman-Mountford et al.,

342 2005; Rubal et al., 2014). In this way, studies that provide a consistent description of

their assemblages and associated habitats are useful in establishing a baseline for the

detection of ecological changes and anthropogenic impacts (Desroy et al., 2002;

Hardman-Mountford et al., 2005; Claudet and Fraschetti, 2010; Dutertre et al., 2013;

346 Schückel et al., 2015; Veiga et al., 2016). Particularly in the shallow subtidal, previous

347 studies have focused in analysing patterns of macrobenthic assemblages along salinity

or depth gradients (e.g. Bris and Glémarec, 1996; Ysebaert et al., 2003; Barros et al.,

349 2008). Moreover, many studies have focused in strongly anthropogenic areas (e.g. Lu,

350 2005; Sánchez-Moyano et al., 2010; Nishijima et al., 2015).

351 The present study explored spatial patterns of macrobenthic assemblages and main

352 features of habitat (i.e. sediment and water column) in an area characterized by

relatively low levels of anthropogenic pressure (Reis, 2012; Reis et al., 2014; Rubal et

al., 2014). Our results indicated that environmental variables varied significantly among

355 localities with the only exceptions of nutrients, salinity, temperature and sediment gravel content. Moreover, the abundance and multivariate structure of macrobenthic 356 assemblages changed significantly among localities, despite these did not exhibit clear 357 differences in their depth or salinity. However, diversity measures (S and H') did not 358 differ significantly among localities. Therefore, our results partially supported the 359 hypothesis that predicted that variables of the water column and sediment beside 360 macrobenthic response variables (N and the multivariate structure of assemblages) 361 362 would differ among localities.

Identifying main environmental variables that shape spatial patterns of benthic fauna is 363 364 not an easy task because they differ among areas (Lu, 2005). None mechanism alone has been universally capable of elucidating faunal patterns identified in different 365 environments, and at any given site, different interacting factors may be involved 366 (Snelgrove and Butman, 1994). Our results provide a good estimation of the percentage 367 of spatial variation of macrobenthos which can be explained by natural environmental 368 conditions providing models for abundance, diversity (S and H') and multivariate 369 370 structure of assemblages in which significant and most relevant environmental variables were selected. Results of the multiple multivariate regression analyses highlighted that 371 the environmental variables here considered explained a substantial proportion of the 372 variability, between 59% for the H' and 72% for the abundance of macrobenthos. 373 Similarly, Dutertre et al. (2013) found that spatial distribution of macrobenthos along 374 375 the subtidal coastal fringe of South Brittany showed a relatively high correlation with environmental factors, and that the combination of 16 natural abiotic variables, 376 377 including sediment characteristics, bathymetry and hydrodynamic conditions, explained 378 the 51% of macrobenthic distribution. This contrasts with other previous studies that 379 yielded low values of the rank correlation coefficient between macrobenthic

assemblages and environmental variables (Lu, 2005) or that found a lower degree of
variation explained by environmental variables (Olsson et al., 2013; Schückel et al.,
2015).

383 Our results also indicated that the identity and number of environmental variables selected for each model were dependent on macrobenthic response variable. Content in 384 385 different sedimentary grain sizes was usually the most relevant factor, explaining the 386 highest percentage of spatial variability in N (fine sand, 56%), H' (silt/clay, 45%) and the multivariate structure of macrobenthic assemblages (very coarse sand, 26%). 387 Sediment features (i.e. mud content, mean grain diameter and sorting index) were also 388 389 the responsible for 20% of the variation in the macrofaunal distribution along the coastal fringe of South Brittany (Dutertre et al., 2013). The relationship between 390 sediment and macrobenthos has been often described in terms of the range of 391 392 granulometric variations tolerated by each species (Ellingsen, 2002; Van Hoey et al., 2004; Hily et al., 2008). According to their lifestyle, macrobenthic species require 393 particular sediment features for instance for tube building, burrowing, or feeding (Self 394 395 and Jumars, 1988; Pinedo et al., 2000). Moreover, the relationship between animals and sediment is quite changeable because sedimentary type can influence other variables 396 like microbial content and food supply, and trophic interactions (Snelgrove and Butman, 397 398 1994; Barros et al., 2008). For example, organic content tends to increase in fine sediments in nearshore ecosystems, improving the food supply for many benthic species 399 400 (Gray and Elliot, 2009). In our case, the locality of Minho had significantly higher content in very coarse, coarse and medium sand and showed as expected the lowest 401 402 organic matter content. Moreover, the multivariate structure of the macrobenthic 403 assemblage in Miño was significantly different to that of other localities. However, in 404 contrast to expected, Miño displayed higher abundance values. This can be explained by

405	the relatively high abundance of some typically interstitial taxa as <i>Pisione</i> , <i>Polygordius</i>
406	or Nematodes that usually live in coarse and medium sand sediments as those found in
407	Miño, which were absent or less abundant in the remaining localities. Mesh-size (0.5
408	mm) could be partially responsible for such differences, because other published works
409	rely on 1-mm mesh which usually do not retain such numbers of small-sized species as
410	those found in this paper at Miño (Lu, 2005; Glockzin and Zetter, 2008; Dutertre et al.,
411	2013). Similarly, Dutertre et al. (2013) found that homogeneous fine sediments showed
412	a relatively low value of species richness because of the absence of interstitial fauna.
413	Salinity is often regarded among the major factors affecting the species richness and
414	composition of macrobenthic assemblages (e.g. Laine, 2003; Bonsdorff, 2006; Gorgina
415	et al., 2010). Lu (2005) found that species number and abundance was negatively
416	correlated with salinity because a lower salinity values mean higher freshwater input
417	from rivers, which is, in turn, a source of nutrients that can be used as food supply.
418	Thus, freshwater input might have affected positively to macrobenthos here. Our results
419	showed that salinity was the most important factor in explaining spatial pattern of
420	macrobenthic richness and the third most relevant for the multivariate structure of
421	macrobenthic assemblages. However, it was not selected in the models of N and H'
422	because its contribution was not significant. However, in our study area, salinity was
423	not significantly different among localities contrasting with previous studies focused on
424	macrobenthos along estuarine gradients (Bris and Glémarec, 1996; Ysebaert et al.,
425	2003; Barros et al., 2008); this may explain its minor relevance in shaping macrobenthic
426	assemblages in the north Portuguese coast. Similarly, Glockzin and Zettler (2008) found
427	that salinity did not vary strongly in their study area and thus its potential value as the
428	main predictor on the benthic assemblage was insignificant.

429 Depth, through its impact on other factors such as organic content, sorting or permeability of sediment, has been usually considered mainly responsible for spatial 430 patterns of macrobenthic assemblages (Bonsdorff, 2006; Zettler et al., 2006; Glockzin 431 432 and Zettler, 2008; Gorgina et al., 2010; Schückel et al., 2015). In our study, however, depth was only the third most relevant factor for N and S, explained a low percentage of 433 variability in the assemblage multivariate structure and it was not selected for H'. This 434 lower influence of depth could be explained because in our study area depth range was 435 small (between 9.5 and 15 m). 436

When variables were individually analysed by sets (i.e. sediment and water column), 437 438 results showed a significant contribution of both sets although sedimentary variables explained a higher percentage of variation in all the considered macrobenthic response 439 variables. However, when sets were sequentially fitted, water variables did not add a 440 significant contribution once that sedimentary set was already included in the model. 441 This points out that, at the scale of the present study, sedimentary variables seem to be 442 more relevant that those of the water column in shaping spatial patterns of 443 macrobenthos, although there is a degree of interdependence among both. Therefore, 444 these results supported our second hypothesis (i.e. sedimentary variables would explain 445 a greater percentage of variability in spatial patterns of macrobenthos than that 446 explained by variables of the water column environment). This may be explained by the 447 greater sediment heterogeneity across localities whereas the water column environment 448 449 was more homogenous. Dutertre et al. (2013) indicated that physical-chemical properties of the water column were important to understand the broad-scale species 450 451 distribution in coastal ecosystems. Schückel et al. (2015) studied macrobenthic assemblages in three tidal channels of the Wadden Sea, and found that differences in 452 453 spatial distribution of macrofaunal assemblages and species composition were best

454	explained by the variability of tidal current velocity and depth, followed by sed	iment
455	characteristics. Therefore, within the tidal channels (i.e. smaller spatial scale), the	he
456	significance of sediment characteristics became higher, in concordance with ou	r results
457	(Schückel et al., 2015).	
		6

Our models highlighted that most variation in spatial patterns of macrobenthic assemblages was explained by the studied environmental variables. However, some 459 460 variation still remains unexplained (between 21 and 48%). This suggests that other abiotic or biological variables did not considered in the present study, such as food 461 supply, source of larvae or interspecies competition might also play a significant role 462 463 (Wildish, 1977).

458

To develop realistic conservation and management strategies to identify and ameliorate 464 465 anthropogenic impacts, managers need baseline ecological measurements from appropriate spatial scales. The lack of these baseline data often results in poor decision-466 making and environmental policy (Yaffee, 1997). Therefore, analysis and measurement 467 468 of the relationship between species and their environment is essential to build predictive models, that provide a global visualisation harmonious with ecosystem management 469 (Gorgina and Zettler, 2010; Méléder et al., 2010). In this way, our study distinguished 470 471 and categorised different natural environmental factors responsible for shaping macrobenthic assemblages from soft-bottom shallow sublittoral in the North Portuguese 472 473 coast and contributes to incorporate a local ecological study in a wide-ranging 474 ecosystem background. The good correlation obtained between natural environmental factors and different macrobenthic response variables advocates that our models may be 475 476 useful to support conservation and management strategies. Moreover, results of this study showed some deviations from generality (i.e. lack of relevance for salinity or 477 478 depth) proving that our understanding of benthic assemblages is locally specific and

thus, conclusions based on studies considering different habitats, spatial scales or 479 480 different components of the benthos cannot be extrapolated as generalities. On the one hand, we establish appropriate baseline data that will be indispensable to future 481 monitoring studies that aim detecting ecological changes as those derived from 482 anthropogenic impacts or preserving the integrity of shallow subtidal areas including 483 their associated biodiversity (Desroy et al., 2002). On the other hand, our results allow 484 focusing on certain environmental factors that have proved shaping the distribution of 485 macrobenthic abundance, diversity and multivariate structure of the assemblage 486 (Gorgina and Zettler, 2010; Méléder et al., 2010). The identification of key 487 environmental factors, mainly related with sediment, is crucial for the conservation of 488 the studied habitat. Any anthropogenic activity that potentially modifies these 489 environmental factors (i.e. dredging, modification of the hydrological regimen due to 490 491 coastal structures or changes in fluvial discharge) could, in turn, modify intensely the structure of soft bottom assemblages. The disturbance of sedimentary habitat for any of 492 493 the previous activities could be especially dramatic in this particular area because of the 494 habitat fragmentation (dominated by rocks) and the limited sources of sediment to buffer these changes. 495

496

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- 508 **References**
- 509 Anderson, M.J., 2001a. A new method for non-parametric multivariate analysis of
- 510 variance. Austral Ecol. 26, 32–46. <u>http://dx.doi.org/10.1111/j.1442-</u>
- 511 <u>9993.2001.01070.pp.x</u>
- 512 Anderson, M.J., 2001b. Permutation tests for univariate or multivariate analysis of
- 513 variance and regression. Can. J. Fish. Aquat. Sci. 58, 626–639.
- 514 <u>http://dx.doi.org/10.1139/f01-004</u>
- 515 Anderson, M.J., 2002. DISTLM v.2: a FORTRAN computer program to calculate a
- 516 distance-based multivariate analysis for a linear model. Department of Statistics,
- 517 University of Auckland.
- 518 Anderson, M.J., 2006. Distance-based tests for homogeneity of multivariate dispersions.
- 519 Biometrics 62, 245–253. <u>http://dx.doi.org/10.1111/j.1541-0420.2005.00440.x</u>
- Araújo, R., Sousa-Pinto, I., Bárbara, I., Quintino, V., 2006. Macroalgal communities of
  intertidal rock pools in the northwest coast of Portugal. Acta Oecol. 30, 192–
- 522 202. http://dx.doi.org/10.1016/j.actao.2006.04.002
- 523 Bale, A.J., Kenny, A.J., 2005. Sediment analysis and seabed characterization, in:
- 524 Eleftheriou, A., McIntyre, A. (Eds.), Methods for the Study of Marine Benthos.
- 525 Blackwell Publishing, Oxford, pp. 43-86.

526	Barros, F., Hatj, V., Figueiredo, M.B., Magalhães, W.F., Dórea, H.S., Emídio, E.S.,
527	2008. The structure of the benthic macrofaunal assemblages and sediments
528	characteristics of the Paraguaçu estuarine system, NE, Brazil. Estuar. Coast.
529	Shelf S. 78, 753–762. <u>http://dx.doi.org/10.1016/j.ecss.2008.02.016</u>
530	Blanchet, H., de Montaudouin, X., Chardy, P., Bachelet, G., 2005. Structuring factors
531	and recent changes in subtidal macrozoobenthic communities of a coastal
532	lagoon, Arcachon Bay (France). Estuar. Coast. Shelf S. 64, 561-576.
533	http://dx.doi.org/10.1016/j.ecss.2005.03.016
534	Bonsdorff, E., 2006. Zoobenthic diversity-gradients in the Baltic Sea: Continuous post-
535	glacial succession in a stressed ecosystem. J. Exp. Mar. Biol. Ecol. 330, 383–
536	391. http://dx.doi.org/10.1016/j.jembe.2005.12.041
537	Bris, L.H., Glémarec, M., 1996. Marine and Brackish Ecosystems of South Brittany
538	(Lorient and Vilaine Bays) with Particular Reference to the Effect of the
539	Turbidity Maxima. Estuar. Coast. Shelf S. 42, 737–753.
540	http://dx.doi.org/10.1006/ecss.1996.0047
541	Clarke, K.R., 1993. Nonparametric multivariate analyses of changes in community
542	structure. Austral J. Ecol. 18, 117–143. <u>http://dx.doi.org/10.1111/j.1442-</u>
543	<u>9993.1993.tb00438.x</u>
544	Claudet, J., Fraschetti, S., 2010. Human-driven impacts on marine habitats: A regional
545	meta-analysis in the Mediterranean Sea. Biol. Cons. 143, 2195–2206.
546	http://dx.doi.org/10.1016/j.biocon.2010.06.004
547	Crossland, C.J., Baird, D., Ducrotoy, J-P., Lindeboom, H., Buddemeier, R.W.,
548	Dennison, W.C., Maxwell, B.A., Smith, S.W., Swaney, D.P., 2005. The Coastal
549	Zone — a Domain of Global Interactions, in: Crossland, C.J., Kremer, H.H.,

	ACCEPTED MANUSCRIPT
550	Lindeboom, H.J., Marshall Crossland, J.I., Le Tissier, M.D.A. (Eds.), Coastal
551	Fluxes in the Anthropocene. Springer-Verlag, Berlin Heidelberg, pp. 1–37.
552	http://dx.doi.org/10.1007/3-540-27851-6_1
553	Dauvin, J.C., 2007. Paradox of estuarine quality: benthic indicators and indices,
554	consensus or debate for the future. Mar. Pollut. Bull. 55, 271–281.
555	http://dx.doi.org/10.1016/j.marpolbul.2006.08.017
556	Dauvin, J.C., Thiébaut, E., Gomez Gesteira, J.L., Ghertsos, K., Gentil, F., Ropert, M.,
557	Sylvan, B., 2004. Spatial structure of a subtidal macrobenthic community in the
558	Bay of Veys (western Bay of Seine, English Channel). J. Exp. Mar. Biol. Ecol.
559	307, 217-235. http://dx.doi.org/10.1016/j.jembe.2004.02.005
560	Desroy, N., Warembourg, C., Dewarumez, J.M, Dauvin, J. C., 2002. Macrobenthic
561	resources of the shallow soft-bottom sediments in the eastern English Channel
562	and southern North Sea. Ices J. Mar. Sci. 60, 120–131.
563	http://dx.doi.org/10.1006/jmsc.2002.1333
564	Dutertre, M., Hamon, D., Chevalier, C., Ehrhold, A., 2013. The use of the relationships
565	between environmental factors and benthic macrofaunal distribution in the
566	establishment of a baseline for coastal management. Ices J. Mar. Sci 70, 294–
567	308. <u>http://dx.doi.org/10.1093/icesjms/fss170</u>
568	Ellingsen, K., 2002. Soft-sediment benthic biodiversity on the continental shelf in
569	relation to environmental variability. Mar. Ecol. Prog. Ser. 232, 15–27.
570	http://dx.doi.org/10.3354/meps232015
571	Esquete, P., Rubal, M., Veiga, P., Troncoso, J.S., 2015. A new species of heterochelous
572	tanaidacean Tanaissus (Paratanaoidea: Tanaissuidae) from the north-west

	ACCEPTED MANUSCRIPT
573	Iberian Peninsula. Zootaxa 3995, 189–202.
574	http://dx.doi.org/10.11646/zootaxa.3995.1.17
575	Glockzin, M., Zettler, M.L., 2008. Spatial macrozoobenthic distribution patterns in
576	relation to major environmental factors- A case study from the Pomeranian Bay
577	(Southern Baltic Sea). J. Sea Res. 59, 144–161.
578	http://dx.doi.org/10.1016/j.seares.2008.01.002
579	Gorgina, M., Glockzin, M., Zettler, M.L., 2010. Distribution of benthic macrofaunal
580	communities in the western Baltic Sea with regard to near-bottom environmental
581	parameters. 1. Causal analysis. J. Marine Syst. 79, 112–123.
582	http://dx.doi.org/10.1016/j.jmarsys.2009.07.006
583	Gorgina, M., Zettler, M.L., 2010. Diversity and distribution of benthic macrofauna in
584	the Baltic Sea. Data inventory and its use for species distribution modelling and
585	prediction. J. Sea Res. 64, 313–321.
586	http://dx.doi.org/10.1016/j.seares.2010.04.005
587	Gray, J.S., Elliott, M., 2009. Ecology of Marine Sediments. From Science to
588	Management, second ed. Oxford University Press, Oxford.
589	Halpern, B.S., Frazier, M., Potapenko, J., Casey, K.S., Koenig, K., Longo, C., Lowndes,
590	J.S., Rockwood, R.C., Selig, E.R., Selkoe, K.A., Walbridge, S., 2015. Spatial
591	and temporal changes in cumulative human impacts on the world's ocean. Nat.
592	Commun. 6, 7615. http://dx.doi.org/10.1038/ncomms8615
593	Halpern, B.S., Walbridge, S., Selkoe, K.A., Kappel, C.V., Micheli, F., D'Agrosa, C.,
594	Bruno, J.F., Casey, K.S., Ebert, C., Fox, H.E., Fujita, R., Heinemann, D.,
595	Lenihan, H.S., Madin, E.M.P., Perry, M.T., Selig, E.R., Spalding, M., Steneck,

	ACCEPTED MANUSCRIPT
596	R., Watson, R., 2008. A Global Map of Human Impact on Marine Ecosystems.
597	Science 319, 948–952. http://dx.doi.org/10.1126/science.1149345
598	Hardman-Mountford, N.J., Allen, J.I., Frost, M.T., Hawkins, S.J., Kendall, M.A.,
599	Mieszkowska, N., Richardson, K.A., Somerfield, P.J., 2005. Diagnostic
600	monitoring of a changing environment: An alternative UK perspective. Mar.
601	Pollut. Bull. 50, 1463–1471. <u>http://dx.doi.org/10.1016/j.marpolbul.2005.06.022</u>
602	Hily, C., Le Loc'h, F., Grall, J., Glemarec, M., 2008. Soft bottom macrobenthic
603	communities of North Biscay revisited: long-term evolution under fisheries-
604	climate forcing. Estuar. Coast. Shelf S. 78, 413–425.
605	http://dx.doi.org/10.1016/j.ecss.2008.01.004
606	Laine, A.O., 2003. Distribution of soft-bottom macrofauna in the deep open Baltic Sea
607	in relation to environmental variability. Estuar. Coast. Shelf S. 57, 87–97.
608	http://dx.doi.org/10.1016/S0272-7714(02)00333-5
609	Legendre, P., Anderson, M.J., 1999. Distance-based redundancy analysis: testing
610	multispecies responses in multifactorial ecological experiments. Ecol. Monogr.
611	69, 1–24. http://dx.doi.org/10.1890/0012-
612	9615(1999)069[0001:DBRATM]2.0.CO;2
613	Lemos, R.T., Pires, H.O., 2004. The upwelling regime off the west Portuguese coast,
614	1941–2000. Int. J. Climatol. 24, 511–524. <u>http://dx.doi.org/10.1002/joc.1009</u>
615	Lu, L., 2005. The relationship between soft-bottom macrobenthic communities and
616	environmental variables in Singaporean waters. Mar. Pollut. Bull. 51, 1034-
617	1040. http://dx.doi.org/10.1016/j.marpolbul.2005.02.013

618	Mann, K.H.	, Lazier, J.R	.N., 2006.	Dynamics	of marine	ecosystems.	Biological-	-Physical
-----	------------	---------------	------------	----------	-----------	-------------	-------------	-----------

619 Interactions in the Oceans, third ed. Blackwell Publishing.

620 <u>http://dx.doi.org/10.1002/9781118687901</u>

- 621 McArdle, B.H., Anderson, M.J., 2001. Fitting multivariate models to community data: a
- 622 comment on distance-based redundancy analysis. Ecology 82, 290–297.

623 <u>http://dx.doi.org/10.1890/0012-9658(2001)082[0290:FMMTCD]2.0.CO;2</u>

- 624 Méléder, V., Populus, J., Guillaumont, B., Perrot, T., Mouquet, P., 2010. Predictive
- 625 modelling of seabed habitats: case study of subtidal kelp forests on the coast of
- 626 Brittany, France. Mar. Biol. 157, 1525–1541. <u>http://dx.doi.org/10.1007/s00227-</u>
- 627 <u>010-1426-4</u>
- 628 Nishijima, W., Umehara, A., Okuda, T., Nakai, S., 2015. Variations in macrobenthic
- 629 community structures in relation to environmental variables in the Seto Inland
  630 Sea, Japan. Mar. Pollut. Bull. 92, 90–98.
- 631 <u>http://dx.doi.org/10.1016/j.marpolbul.2014.12.051</u>
- 632 Olsson, J., Bergström, L., Gårdmark, A., 2013. Top-down regulation, climate and multi-

633 decadal changes in coastal zoobenthos communities in two Baltic Sea areas.

- 634 PLoS ONE 8(5): e64767. <u>http://dx.doi.org/10.1371/journal.pone.0064767</u>
- 635 Pinedo, S., Sardá, R., Rey, C., Bhaud, C., 2000. Effect of sediment particle size on
- 636 recruitment of *Owenia fusiformis* in the Bay of Blanes (NW Mediterranean Sea):
- 637 an experimental approach to explain field distribution. Mar. Ecol. Prog. Ser. 203,
- 638 205–213. <u>http://dx.doi.org/10.3354/meps203205</u>
- 639 Pratt, D.R., Lohrer, A.M., Pilditch, C.A., Thrush, S.F., 2014. Changes in ecosystem
- function across sedimentary gradients in estuaries. Ecosystems 17, 182–194.
- 641 <u>http://dx.doi.org/10.1007/s10021-013-9716-6</u>

642	Ramey, P.A., Bodnar, E., 2008. Active post-settlement selection by a deposit-feeding
643	polychaete, Polygordius jouinae, for sands with relatively high organic content.
644	Limnol. Oceanogr. 53, 1512–1520. http://dx.doi.org/10.4319/lo.2008.53.4.1512
645	Reis, P.A., 2012. Cirripides Chthamalus montagui e Pollicipes pollicipes como
646	biomonitores de contaminação por metais e hidrocarbonetos aromáticos
647	policíclicos ao longo da costa noroeste de Portugal. PhD dissertation. University
648	of Porto, Portugal.
649	Reis, P.A., Cassiano, J., Veiga, P., Rubal, M., Sousa-Pinto, I., 2014. Fucus spiralis as
650	monitoring tool of metal contamination in the northwest coast of Portugal under
651	the European Water Framework Directives. Environ. Monit. Assess. 186, 5447–
652	5460. http://dx.doi.org/10.1007/s10661-014-3794-6
653	Rodríguez, M., Costa, J.Z., Monterroso, Ó., Riera, R., Pérez, O., Sacramento, A.T.,
654	Ramos, E., 2011. Estudo de caracterização dos habitats da faixa costeira do
655	Litoral Norte de Portugal, CIMA S.L Informe Técnico 74.
656	Rubal, M., Veiga, P., Vieira, R., Sousa-Pinto, I., 2011. Seasonal patterns of tidepool
657	macroalgal assemblages in the North of Portugal. Consistence between species
658	and functional group approaches. J. Sea Res. 66:187–194.
659	http://dx.doi.org/10.1016/j.seares.2011.07.003
660	Rubal, M., Veiga, P., Fontoura, P., Sousa-Pinto, I., In press. A new Batillipes
661	(Tardigrada, Heterotardigrada, Batillipedidae) from North Portugal (Atlantic
662	Ocean). Mar. Biodiv. http://dx.doi.org/10.1007/s12526-016-0526-x
663	Rubal, M., Veiga, P., Reis, P.A., Bertocci, I., Sousa-Pinto, I., 2014. Effects of subtle
664	pollution at different levels of biological organisation on species-rich

	ACCEPTED MANUSCRIPT
665	assemblages. Environ. Pollut. 191, 101-110.
666	http://dx.doi.org/10.1016/j.envpol.2014.04.019
667	Sánchez-Moyano, J.E., García-Asencio, I., García-Gómez, J.C., 2010. Spatial and
668	temporal variation of the benthic macrofauna in a grossly polluted estuary from
669	southwestern Spain. Helgol. Mar. Res. 64, 155–168.
670	http://dx.doi.org/10.1007/s10152-009-0175-6
671	Schückel, U., Beck, M., Kröncke, I., 2015. Macrofauna communities of tidal channels
672	in Jade Bay (German Wadden Sea): spatial patterns, relationships with
673	environmental characteristics, and comparative aspects. Mar. Biodiv. 45, 841-
674	855. <u>http://dx.doi.org/10.1007/s12526-014-0308-2</u>
675	Self, R.F.L., Jumars, P.A., 1988. Cross-phyletic patterns of particle selection by deposit
676	feeders. J. Mar. Res. 46, 119–143.
677	Skalar, 2004. The San Plus Continuous Flow Analyzer - User Manual: Section 8 -
678	Operating the SFA Analyzer. Skalar Analytical B.V., DE Breda, Netherlands.
679	Snelgrove, P.V.R., 1998. The biodiversity of macrofaunal organisms in marine
680	sediments. Biodivers. Conserv. 7, 1123-1132.
681	http://dx.doi.org/10.1023/A:1008867313340
682	Snelgrove, P.V.R., Butman, C.A., 1994. Animal-sediment relationships revisited: Cause
683	versus effect. Oceanogr. Mar. Biol. Ann. Rev. 32, 111–177.
684	Terlizzi, A., Benedetti-Cecchi, L., Bevilacqua, S., Fraschetti, S., Guidetti, P., Anderson,
685	M.J., 2005. Multivariate and univariate asymmetrical analyses in environmental
686	impact assessment: a case study of Mediterranean subtidal sessile assemblages.
687	Mar. Ecol. Prog. Ser. 289, 27e42. http://dx.doi.org/10.3354/meps289027

688	Underwood, A.J., 1997. Experiments in Ecology: Their logical design and interpretation
689	using analysis of variances, Cambridge University Press, Cambridge.
690	Van Hoey, G., Degraer, S., Vinx, M., 2004. Macrobenthic community structure of soft-
691	bottom sediments at the Belgian Continental Shelf. Estuar. Coast. Shelf S., 59,
692	599–613. <u>http://dx.doi.org/10.1016/j.ecss.2003.11.005</u>
693	Veiga, P., Rubal, M., Besteiro, C., 2009. Shallow sublittoral meiofauna communities
694	and sediment polycyclic aromatic hydrocarbons (PAHs) content on the Galician
695	coast (NW Spain), six months after the Prestige oil spill. Mar. Pollut. Bull. 58,
696	581–588. <u>http://dx.doi.org/10.1016/j.marpolbul.2008.11.002</u>
697	Veiga, P., Rubal, M., Cacabelos, E., Maldonado, C., Sousa-Pinto, I., 2014. Spatial
698	variability of macrobenthic zonation on exposed sandy beaches. J. Sea Res. 90,
699	1–9. <u>http://dx.doi.org/10.1016/j.seares.2014.02.009</u>
700	Veiga, P., Rubal, M., Vieira, R., Arenas, F., Sousa-Pinto, I., 2013. Spatial variability in
701	intertidal macroalgal assemblages on the North Portuguese coast. Consistence
702	between species and functional groups approaches. Helgoland Mar. Res. 67,
703	191-201. http://dx.doi.org/10.1007/s10152-012-0315-2
704	Veiga, P., Torres, A.C., Aneiros, F., Sousa-Pinto, I., Troncoso, J.S., Rubal, M., 2016.
705	Consistent patterns of variation in macrobenthic assemblages and environmental
706	variables over multiple spatial scales using taxonomic and functional
707	approaches. Mar. Environ. Res. 120, 191–201.
708	http://dx.doi.org/10.1016/j.marenvres.2016.08.011
709	Wildish, D.J., 1977. Factors controlling marine and estuarine sublittoral macrofauna.
710	Helgoland. Wiss. Meer. 30, 445–454.

711	Worm, B., Barbier, E.B., Beaumont, N., Duffy, J.E., Folke, C., Halpern, B.S., Jackson,
712	J.B.C., Lotze, H.K., Micheli, F., Palumbi, S.R., Sala, E., Selkoe, K.A.,
713	Stachowicz, J.J., Watson, R., 2006. Impacts of biodiversity loss on ocean
714	ecosystem services. Science 314, 787-790.
715	http://dx.doi.org/10.1126/science.1132294
716	Yaffee, S.L., 1997. Why environmental policy nightmares recur. Conserv. Biol. 11,
717	328–337. <u>http://dx.doi.org/10.1046/j.1523-1739.1997.95204.x</u>
718	Yamanaka, T., White, P.C.L., Spencer, M., Raffaelli, D., 2012. Patterns and processes
719	in abundance-body size relationships for marine benthic invertebrates. J. Anim.
720	Ecol. 81, 463–471. <u>http://dx.doi.org/10.1111/j.1365-2656.2011.01921.x</u>
721	Ysebaert, T., Herman, P.M.J., Meire, P., Craeymeersch, J., Verbeek, H., Heip, C.H.R.,
722	2003. Large-scale spatial patterns in estuaries: estuarine macrobenthic
723	communities in the Schelde estuary, NW Europe. Estuar. Coast. Shelf S. 57,
724	335-355. http://dx.doi.org/10.1016/S0272-7714(02)00359-1
725	Zettler, M.L., Röhner, M., Frankowski, J., 2006. Long term changes of
726	macrozoobenthos in the Arkona Basin (Baltic Sea). Boreal Env. Res. 11, 247-
727	260.
728	

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730	Fig. 1 Mean values (+SE) of oxygen concentration (A), salinity (B) and temperature (C)
731	in the water column at each site of the four studied localities. Letters indicate significant
732	differences among localities ( $p < 0.05$ ) as detected by SNK test. ns: not significant
733	differences among localities.
734	Fig. 2 Mean values (+SE) of NO <sub>3</sub> : nitrate (A), PO <sub>4</sub> : phosphorus (B) and NH <sub>3</sub> :
735	ammonium (C) in the water column at each site of the four studied localities. ns: not
736	significant differences among localities.
737	Fig. 3 Mean values (+SE) of sediment grain size (A-H) and organic matter content (I) in
738	percentage, at each site of the four studied localities. Letters indicate significant
739	differences among localities ( $p < 0.05$ ) as detected by SNK test. ns: not significant
740	differences among localities.
741	Fig. 4 Mean values (+SE) of N: total number of individuals (A), S: taxon richness (B)
742	and H': Shannon diversity index of macrobenthos at each site of the four studied
743	localities. Letters indicate significant differences among localities ( $p < 0.05$ ) as detected
744	by SNK test. ns: not significant differences among localities.
745	Fig. 5 nMDS ordination of macrobenthic assemblages based on square-root transformed
746	abundances and Bray-Curtis similarities in the four studied localities. Stress: 0.18.
747	Fig. 6 Distance-based redundancy (dbRDA) plot illustrating the DISTLM model based
748	on the macrobenthic assemblages and the fitted environmental variables as vectors
749	based on DistLM analysis in Table 7 for the assemblage.
750	

## 729 Figure caption

Locality	Site	Latitude	Longitude	Depth (m)	Md (mm)	Sediment type	QDΦ	Sorting
Miño	1	41°51′10.01″N	8°52′54.00″W	13	$0.548 \pm 0.068$	coarse sand	$0.485\pm0.057$	well sorted
	2	41°51′2.82″N	8°52′40.80″W	10	$0.435\pm0.020$	medium sand	$0.510 \pm 0.022$	moderately well sorted
	3	41°50′44.87″N	8°52′44.76″W	13.5	$0.332\pm0.018$	medium sand	$0.410\pm0.058$	well sorted
Âncora	1	41°48′54.72″N	8°52′54.48″W	15	0.162 ±0.001	fine sand	$0.378\pm0.009$	well sorted
	2	41°48′37.80″N	8°52′48.78″W	12.5	$0.164 \pm 0.001$	fine sand	$0.371\pm0.004$	well sorted
	3	41°48′17.47″N	8°53′4.86″W	13	$0.178 \pm 0.001$	fine sand	$0.325\pm0.002$	very well sorted
Gelfa	1	41°47′8.88″N	8°53′14.28″W	15	$0.239 \pm 0.079$	fine sand	$0.45\pm0.010$	well sorted
	2	41°46′52.38″N	8°53′6.84″W	12.5	$0.232 \pm 0.070$	fine sand	$0.426\pm0.051$	well sorted
	3	41°46′34.26″N	8°53′4.20″W	11	$0.169 \pm 0.002$	fine sand	$0.375\pm0.008$	well sorted
Lima	1	41°40′13.86″N	8°50′22.74″W	9.5	$0.152 \pm 0.002$	fine sand	$0.602\pm0.015$	moderately well sorted
	2	41°39′57.73″N	8°50′22.50″W	10.5	$0.140\pm0.008$	fine sand	$0.599 \pm 0.011$	moderately well sorted
	3	41°39'39.72''N	8°50′24.43″W	11	$0.137 \pm 5.561 \ 10^{-6}$	fine sand	$0.578 \pm 0.002$	moderately well sorted
				ER TE				

**Table 1**. Environmental data of the studied localities. Md: median particle size of sediment; QD $\Phi$ : sorting coefficient of sediment.

**Table 2.** Results of ANOVAs testing differences in coastal water variables amonglocalities; \*\*: p < 0.01; \*\*\*: p < 0.001; \*: significant; ns: not significant. Relevantsignificant differences indicated in bold.

Source of variation	df		kygen ntration	Sali	nity	Temperature		
Source of variation	ui	MS	F	MS	F	MS	F	
Locality	3	0.0009	<b>9.34</b> **	0.0012	4.11	0.0645	<b>8.82</b> **	
Residual	8	0.0001		0.0003	0.0003			
Total	11							
Cochran's test		0.8	8525 <sup>s</sup>	0.96	597 <sup>s</sup>	$0.8702^{s}$		
Transformation		n	one	no	ne	none		
	đf	ľ	NO <sub>3</sub>	N	H <sub>3</sub>	PO	D <sub>4</sub>	
	ui	MS	F	MS	F	MS	F	
Locality	3	2.1746	1.34	83.500	2.83	0.0247	1.79	
Site (Locality)	8	1.6239	$45.05^{***}$	29.531	0.55	0.0138	$8.04^{***}$	
Residual 24		0.0360				0.0017		
Total	35							
Cochran's test		0.2	2818 <sup>ns</sup>	0.16	99 <sup>ns</sup>	0.3314 <sup>ns</sup>		
Transformation		Sqr	t(X+1)	no	ne	Sqrt(	X+1)	

**Table 3.** Results of ANOVAs testing differences in sedimentary grain sizes and organic matter content (in percentage) among localities. \*: p < 0.05; \*\*: p < 0.01; \*\*\*: p < 0.001; \*: p < 0.001; \*:

Course of mariation	46	Coarse g	gravel	Fine g	ravel	Very coa	rse sand	Coarse	sand	Medium	sand
Source of variation	ai	MS	F	MS	F	MS	F	MS	F	MS	F
Locality	3	0.0002	1.00	0.0208	3.43	0.6724	11.79**	1591.7965	<b>6.34</b> <sup>*</sup>	2716.0566	12.96**
Site (Locality)	8	0.0002	1.00	0.0061	2.55	0.057	12.88***	251.2306	$9.48^{***}$	209.5186	0.74
Residual	12	0.0002		0.0024		0.0044	S S	26.4955		284.7926	
Total	23										
Cochran's test		1.00	) <sup>s</sup>	0.40	$24^{ns}$	0.40	$22^{ns}$	0.742	25 <sup>s</sup>	0.530	$2^{ns}$
Transformation		non	e	noi	ne	Sqrt(	X+1)	non	e	non	e
							1			_	
	đf	Fine s	and	Very fir	ne sand	Silt/	clay	Organic	matter		
	ui	MS	F	MS	F	MS	F	MS	F		
Locality	3	3927.4412	$22.25^{***}$	868.6802	<b>21.78</b> ***	115.6667	<b>41.87</b> <sup>***</sup>	0.5483	<b>16.07</b> **	_	
Site (Locality)	8	176.5464	1.07	39.8859	1.53	2.7623	1.64	0.0341	0.57		
Residual	12	164.8805		26.1224		1.6820		0.0596			
Total	23										
Cochran's test		0.528	9 <sup>ns</sup>	0.5119 <sup>ns</sup>		$0.5007^{ns}$		0.5182 <sup>ns</sup>			
Transformation		non	e	noi	ne	no	none		e		

**Table 4.** Results of ANOVAs testing for differences in the total number of individuals(N), taxon richness (S) and Shannon's diversity index (H') of macrobenthos amonglocalities; \*: p < 0.05; \*\*\*: p < 0.001; <sup>ns</sup>: not significant; <sup>s</sup>: significant. Relevantsignificant differences indicated in bold.

Course of variation	đf		N	N S		H'	
Source of variation	ai	MS	F	MS	F	MS F	
Locality	3	11.57	<b>6.5</b> 6 <sup>*</sup>	27.66	0.58	$2.21  5.46^*$	
Site (Locality)	8 1.76 7.62***		$7.62^{***}$	47.40 7.558***		$0.41 \ 2.84^{*}$	
Residual	48	0.23		6.25		0.14	
Total	59						
Cochran's test		0.	19 <sup>ns</sup>	$0.26^{ns}$		$0.32^{s}$	
Transformation		Ln (X+1)		none		none	

**Table 5**. Results of PERMANOVAs testing differences in the structure of macrobenthic assemblage among localities. Analyses based on Bray–Curtis dissimilarity matrix from square-root transformed data. All tests used 999 random permutations. Monte Carlo p-values were considered for pair-wise test because the number of unique permutations was lower than 30. \*: p < 0.05; \*\*: p < 0.01. Relevant significant differences indicated in bold.

Source of variation	df	MS	Pseudo-F	Unique perms
Lo	3	15402	2.68*	984
Si (Lo)	8	5752.8	$5.42^{**}$	996
Residual	48	1061.6		
Total	59			
Pair-wise test			ť	
Miño vs Âncora			1.74*	
Miño vs Gelfa			1.92*	
Miño vs Lima			2.27**	
Âncora vs Gelfa			0.89	
Âncora vs Lima			1.32	
Gelfa vs Lima			0.10	

**Table 6.** Contribution ( $\delta i$ ) of individual taxa to the average Bray-Curtis dissimilarity among localities that showed significant differences in the

# structure of their assemblages.

Species	Average Abundance				N	Miño-Ancora			Miño-G	elfa	Miño-Lima				
Species	Miño	Âncora	Gelfa	Lima	δi	δi%	$\delta i/SD(\delta i)$	δi	δi%	$\delta i/SD(\delta i)$	δί	δi%	$\delta i/SD(\delta i)$		
Pisione parapari	8.9	0	0	0	10.23	13.83	1.25	11.59	14.34	1.23	10.72	12.76	1.24		
Nematoda spp.	6.41	0.13	0.23	0.23	7.65	10.34	2.08	8.55	10.58	2.08	7.93	9.44	2.02		
Gastrosaccus spinifer	5.75	4.03	1.01	0.78	7.48	10.12	1.15	6.66	8.24	1.14	6.52	7.76	1.18		
Polygordius appendiculatus	6.05	0	0	0	6.78	9.17	1.1	7.67	9.48	1.1	7.1	8.45	1.1		
Diogenes pugilator	4.24	3.09	1.54	1.18	5.38	7.27	1.05	6.89	8.53	0.9	6.52	7.76	0.92		
Spio decoratus	1.24	4.62	3.41	3.46	4.5	6.08	1.04	3.85	4.77	1.54	3.16	3.77	1.69		
Pisione remota	3.01	0	0	0	3.3	4.46	1.01	3.71	4.59	1.01	3.45	4.1	1.01		
Bodotria arenosa	0.48	1.86	1.27	1.05	2.14	2.89	1.36	1.72	2.12	1.26	1.58	1.89	1.09		
Bathyporeia tenuipes	0.07	1.58	0.74	0.57	2.11	2.85	1.26	1.1	1.37	0.88	0.93	1.1	0.69		
Eurydice sp.	1.75	1.38	0.34	0.33	1.92	2.6	1.14	2.81	3.47	1.01	2.47	2.94	1.07		
Scolelepis squamata	0.36	1.42	0.09	0.16	1.66	2.24	1.4	0.73	0.9	0.68	0.68	0.81	0.73		
Nototropis falcatus	1.05	1.39	0.99	0.87	1.59	2.14	0.88	1.74	2.15	0.79	1.4	1.66	0.83		
Pontocrates altamarinus	0.28	1.16	1.4	1.89	1.5	2.02	1.14	2.01	2.49	1.06	2.42	2.88	1.66		
Calanoida spp.	1.27	0.2	0.42	0	1.47	1.99	0.81	2.07	2.56	0.87	1.59	1.89	0.8		
Nephtys cirrosa	0.95	1.58	1.03	0.63	1.32	1.78	1.28	1.41	1.75	1.21	1.27	1.51	1.24		
Micronephtys stammeri	0.56	1.01	0.88	1.57	1.28	1.73	1.18	1.23	1.52	1.1	1.62	1.93	1.47		
Magelona johnstoni	0.07	0.97	0.99	1.26	1.25	1.69	1.02	1.45	1.79	0.99	1.73	2.05	1.19		
Bathyporeia nana	0	0.77			1.06	1.43	0.81								
Urothoe brevicornis	0.19	0.57	0.33		0.84	1.14	0.82	0.77	0.96	0.73					
Liocarcinus marmoreus	0.64	0.33	0.07	0	0.84	1.14	1.01	1.06	1.31	0.97	0.97	1.15	0.98		
Saccocirrus papillocercus	0.65	0	0	0	0.69	0.94	0.63	0.78	0.96	0.64	0.72	0.86	0.63		
Diastylis rathkei	0	0.48	0.56	1.06	0.65	0.88	0.72	0.77	0.95	0.54	1.46	1.73	1.37		
Spiophanes bombyx	0	0.51	0.74	0.67	0.65	0.88	0.71	1.04	1.29	0.86	0.9	1.07	0.84		
Fabulina fabula	0			1.85							2.55	3.04	1.27		
Mediomastus fragilis	0			1.33			`)´				1.76	2.1	1.1		
Glycera tridactyla	0.4		0.36	0.77		<u> </u>		0.84	1.04	0.83	1.17	1.39	1.05		
Nephtys assimilis	0			0.69		· ) `					0.97	1.16	0.95		
Iphinoe trispinosa	0			0.7		-					0.92	1.1	0.63		
Hippomedon denticulatus	0			0.56	(	\					0.77	0.91	0.69		
Owenia fusiformis	0			0.5		/					0.68	0.81	0.75		
Spisula solida	0.45		0	0				0.75	0.93	0.77	0.68	0.8	0.78		
Magelona mirabilis	0			0.49	X'						0.67	0.8	0.66		
Eocuma dollfusi	0.07		0.65		Y			1.06	1.32	0.83					
Tanaidacea spp.	0.07		0.4					0.59	0.73	0.59					

**Table 7**. Results of DistLM carried out to ascertain the role of different environmental variables, considering forward-selection of variables, where amount explained by each variable added to model is conditional on variables already in the model (i.e. those variables listed above it)on the total number of individuals (N), taxon richness (S), Shannon's diversity index (H') and the structure of macrobenthic assemblages. %Var: percentage of variance in species data explained by that variable; Cum. %: cumulative percentage of variance explained. Variables significantly related to macrobenthic assemblages indicated in bold.

	Ν				S		
Variable	%Var	Pseudo-F	Cum (%)	Variable	%Var	Pseudo-F	Cum (%)
Fine sand	56.30	74.73**	56.30	Salinity	24.65	18.97**	24.65
Medium sand	3.56	5.06*	59.87	Silt/clay	6.65	$5.52^{*}$	31.30
Depth	6.02	<b>9.89</b> **	65.89	Depth	4.48	<b>3.91</b> *	35.77
Silt/Clay	1.23	2.10	67.12	NO <sub>3</sub>	10.68	10.95**	46.44
NO <sub>3</sub>	1.89	3.29	69.01	NH <sub>3</sub>	6.46	<b>7.4</b> 1 <sup>**</sup>	52.90
Very fine sand	1.52	2.74	70.53	Organic matter	5.59	<b>7.14</b> *	58.49
NH <sub>3</sub>	0.27	0.47	70.79	Very coarse sand	0.46	0.58	58.95
Organic matter	0.44	0.77	71.28	Fine sand	0.85	1.08	59.80
Coarse gravel	0.04	6.29 10 <sup>-2</sup>	71.26	Oxygen	0.12	0.15	59.92
Oxygen	0.03	5.47 10 <sup>-2</sup>	71.30	Fine gravel	0.34	0.41	60.25
Salinity	0.21	0.36	71.51	Coarse gravel	0.04	0.47	60.64
			Y				
		/	Assemblage				
Variable	%Var	Pseudo-F	Cum (%)	Variable	%Var	Pseudo-F	Cum (%)
Silt/clay	46.46	50.33**	46.46	Very coarse sand	25.61	20.00**	25.61
Organic matter	2.96	3.33	49.42	Very fine sand	8.27	7.13**	33.88
NH <sub>3</sub>	1.92	2.21	51.33	Salinity	4.99	<b>4.58</b> <sup>**</sup>	38.87
Depth	1.70	1.99	53.03	Fine sand	4.23	<b>4.09</b> <sup>**</sup>	43.11
NO <sub>3</sub>	2.87	3.52	55.90	Medium sand	4.22	<b>4.33</b> **	47.33
Fine sand	0.96	1.17	56.86	Silt/clay	3.17	3.39**	50.50
Very coarse sand	0.46	0.56	57.31	Oxygen	4.12	<b>4.7</b> 2 <sup>**</sup>	54.62
Medium sand	0.11	0.13	57.42	Coarse gravel	2.85	3.42**	57.47
Very fine sand	1.56	1.91	58.99	Depth	3.15	<b>4.01</b> <sup>**</sup>	60.62
Fine gravel	0.05	5.92 10 <sup>-2</sup>	59.04	NH <sub>3</sub>	2.27	3.00**	62.89
-				NO <sub>3</sub>	1.52	$2.05^{*}$	64.41

**Table 8**. Results of DistLM carried out to ascertain the role of on sets of environmental variables (coastal water and sediment), considering each set of variables taken individually (ignoring other sets) and forward-selection of sets of variables, where amount explained by each set added to model is conditional onset already in the model (i.e. that listed above it) on the total number of individuals (N), taxon richness (S), Shannon diversity index (H') and the structure of macrobenthic assemblages. %Var: percentage of variance in species data explained by that set of variables; Cum. %: cumulative percentage of variance explained. Sets of variables significantly related to macrobenthic assemblages indicated in bold.

	Ν				S			
Set	%Var	Pseudo-F	Cum (%)	Set	%Var	Pseudo-F	Cum (%)	
Sets individually				Sets individually	Y			
Sediment	69.67	14.65**		Sediment	54.78	$7.72^{**}$		
Water	63.50	12.93**		Water	48.81	7.08**		
Sets fitted sequentially				Sets fitted sequentially				
Sediment	69.67	14.65**	69.67	Sediment	54.78	$7.72^{**}$	54.78	
Water	1.84	1.03	71.51	Water	5.85	2.38	60.64	
H'				Assemblage				
Set	%Var	Pseudo-F	Cum (%)	Set	%Var	Pseudo-F	Cum (%)	
Sets individually			(	Sets individually				
Sediment	56.93	<b>8.43</b> <sup>**</sup>	~	Sediment	57.17	<b>8.5</b> 1 <sup>**</sup>		
Water	37.62	<b>4.48</b> <sup>**</sup>		Water	51.26	<b>7.81</b> <sup>**</sup>		
Sets fitted sequentially				Sets fitted sequentially				
Sediment	56.93	<b>8.43</b> ***	56.93	Sediment	57.17	8.51**	57.18	
Water	2.11	0.83	59.04	Water	7.25	3.26**	64.41	







Fig. 2





Fig. 4









Spatial patterns of macrobenthos in shallow sublittoral sediments. Abundance, diversity and multivariate structure of macrobenthos. Models selecting main environmental variables shaping macrobenthos. Most of macrobenthic variability was explained by environmental variables. Sedimentary variables were more relevant that those of the water column.

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