2	Paeonia subsect. Delavayanae
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19	Running head: Population genetics of Paeonia subsect. Delavayanae
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Population genetic dynamics of Himalayan-Hengduan tree peonies,

21 Abstract

22 According to the present taxonomical treatment, *Paeonia* subsect. *Delavayanae* consists of only two species (P. delavayi and P. ludlowii) endemic to the Himalayan-Hengduan Mountains. 23 Although P. ludlowii can be distinguished from P. delavavi on the basis of a series of 24 morphological characters, the species delimitation remains controversial because the more 25 widespread one, P. delavavi, exhibits considerable morphological diversity. Both chloroplast 26 DNA markers and nuclear microsatellites or simple sequence repeats (nSSR) are used herein to 27 reveal genetic diversity and relationships of the two taxa included in this subsection, and 28 29 ecological niche modeling (ENM) is employed to get insights into their paleodistribution. Our results show that genetic boundaries between the two currently recognized species are unclear, 30 probably due to recent divergence. Paeonia ludlowii is budding from P. delavavi, probably by 31 genetic isolation but also by shifting its niche to the harsher upland Tibetan conditions. Paeonia 32 delavavi itself would be, however, under active speciation, showing significant genetic 33 differentiation and morphological diversity. Whereas P. ludlowii would have endured the 34 35 Pleistocene glacial periods by in situ persistence in local, small refugia, a 'dual' model seems to apply for P. delavayi (in situ persistence and retreat to refugia). The rarity of P. ludlowii and 36 high evolutionary potential of *P. delavavi* imply high priority of *in situ* conservation of both 37 taxa. The Himalayan-Hengduan Mountains are an ideal arena for differentiation within subsect. 38 Delavayanae of Paeonia, by means of expansions/contractions/displacements, vertical 39 migration, and local survival/extinctions in response to the Neogene climate fluctuations and 40 41 geological changes.

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43 Keywords: biogeography; conservation; genetic diversity; *Paeonia*; differentiation.

45 **1. Introduction**

Mountain ranges are regarded as one of the most important drivers of plant evolutionary 46 divergence and speciation, both in the tropics and in temperate regions (Hugues and Atchison, 47 2015; Schwery et al., 2015). The formation and the complex orography of mountains, often 48 coupled with large climatic changes, shaped the historical population demography and genetic 49 diversity generally by strong selection, quick genetic drift and limited gene flow. As a 50 consequence, mountains are generally associated with high levels of plant diversity, both in 51 terms of species richness and endemism (e.g., Barthlott et al., 2005). The Qinghai-Tibetan 52 53 Plateau (QTP), the highest (with an average elevation exceeding 4500 m) and one of the largest plateaus of the world (with about 2.5 million km²), is one of the regions with most active plant 54 evolution (Liu et al., 2014; Wen et al., 2014; Favre et al., 2015; Xing and Ree, 2017), largely 55 due to its active tectonism (with major uplift events commencing 40 Ma and continuing at 56 present; Mulch and Camberlain, 2006), large altitudinal gradients (from 500 m at the foot of 57 the Himalayas to the 8848 m of the Mt. Everest), and the enormous diversity of habitats and 58 59 climatic conditions (e.g., the annual precipitation may vary from 100 mm to about 3000 mm; Favre et al., 2015). However, most speciation events happened in the southernmost part, the 60 Himalayas and the Hengduan Mountains, which are recognized as two of the 25 global 61 biodiversity hotspots. It is estimated that there are around 10,000 species (with 3160 endemisms) 62 in the Himalayas, and ca. 12,000 species (with 3500 endemisms) in the Hengduan Mountains 63 (Mittermeier et al., 2011). The division of Himalayas and Hengduan Moutains is somewhat 64 arbitrary, and indeed many species occur in both regions, which are increasingly referred to as 65 'Himalayan-Hengduan Mountains (HHM)' (e.g., Luo et al., 2016). 66

The HHM is an arena for active plant radiation and speciation in the world due to the 67 combination of continued mountain uplift, a very complex topography (the Hengduan 68 Mountains are probably the most rugged ones on Earth), and Quaternary climatic oscillations. 69 In response to these factors, plants have diversified through multiple mechanisms that include 70 allopatric speciation, hybridization, polyploidy, ecological adaptation, and even morphological 71 72 innovations. Mountain uplift is certainly the main trigger for some of the best-known HHM rapid radiations (e.g., Aconitum L./Delphinium L., Pedicularis L., Rhododendron L., Saussurea 73 DC., or the Ligularia Cass./Cremanthodium Benth./Parasenecio W.W. Sm. & J. Small complex; 74 Wen et al., 2014; Hugues and Atchison, 2015), and also for intraspecific differentiation in some 75 plant species [e.g., Hippophae tibetana Schltdl., Meconopsis integrifolia (Maxim.) Franch., and 76 Taxus wallichiana Zucc.; Wang et al., 2010; Yang et al., 2012; Liu et al., 2013], given that uplift 77 would have been active until Pliocene and Pleistocene (Li and Fang, 1999; Favre et al., 2015). 78 Nevertheless, the main triggers that mostly delineated the intraspecific genetic and 79 phylogeographic structure of HHM plants were the Quaternary climatic oscillations coupled 80 with the very complex regional topography. Although the QTP was not covered by a large ice 81 sheet (such as the Laurentide one), glaciers and ice caps occupied large parts (about 350,000 82 km²; Shi, 2002), especially in the Himalayas. In contrast, the Hengduan Mountains, particularly 83 84 their southern section, remained relatively ice-free (Li et al., 1991; Shi, 2002), likely having served as a refugium. The expectation of range shifts from the Himalayas (or other parts of the 85 QTP) to large glacial refugia located in the Hengduan Mountains during glacial periods 86 (followed by postglacial recolonizations) is however, not always met; many HHM plant species, 87 especially those cold-tolerant, instead of migrating into these (warm) refugia, would have 88

survived *in situ* in small refugia (Qiu et al., 2011; Liu et al., 2014; Luo et al., 2016).

Paeoniaceae is a small family of only one genus, Paeonia L., and ca. 33 species, 90 91 distributed in the northwestern corner of Africa, Europe, Asia and western North America (Hong, 2010). The genus has diverged into woody (sect. Moutan DC., commonly known as 92 'tree peony' and mudan in Chinese) as well as herbaceous forms (sect. Paeonia and sect. 93 Onaepia Lindl.). The tree peony is crowned 'the king of flowers' in China for its beauty. It first 94 95 appeared in royal gardens during the Tang dynasty (618–907 AD) or earlier, and today it is extensively cultivated in China for medicinal and oil uses. About 1000 ornamental cultivars 96 97 have been created in China alone (He and Xing, 2015). The section Moutan contains nine diploid (2n = 10) species endemic to China and is subdivided into two subsections, 98 Delavayanae Stern and Vaginatae Stern (Hong, 2010). The subsect. Delavayanae consists of 99 only two species (Hong, 2010). One is P. delavayi Franch., restricted to sparse thickets, woods 100 or forests in southeastern Tibet, northern Yunnan and western Sichuan provinces at altitudes of 101 1900-4000 m (Fig. 1A). The other is P. ludlowii (Stern & G. Taylor) D.Y. Hong, which is 102 103 endemic to a small area in three counties (Lhünzê, Mainling, Nyingchi) of southeastern Tibet (Fig. 1A). 104

The taxonomy of subsect. Delavayanae has been controversial because the type species, 105 P. delavavi, exhibits considerable morphological diversity in the width of leaf segments and the 106 number, size, and color of floral parts (Hong et al., 1998). A plethora of names at specific or 107 varietal ranks was given to morphoforms; for example, P. delavavi for the dark red flower form, 108 109 P. lutea Delavay ex Franch. for the yellow flower form, P. potaninii Komarov for the finely lobed form, P. trollioides Stapf ex Stern for the yellowish pink or pinkish yellow flower form, 110 and P. weisiensis Y. Wang & K. Li for the white or pale yellow flower form. All these names 111 have been synonymized under P. delavayi with the exception of P. lutea var. ludlowii Stern & 112 G. Taylor, which was upgraded to the species rank (P. ludlowii) on the basis of a series of 113 morphological characters (single carpels, larger follicles, yellow petals, and lack of stolons 114 (Hong, 1997, 2010; Hong et al., 1998). Such two-species taxonomic treatment for subsect. 115 Delavayanae was supported by recent molecular evidence (Zhang et al., 2009a; Zhou et al., 116 117 2014).

It is expected that, being dwellers in HHM region, P. delavayi and P. ludlowii would have 118 experienced range expansions/contractions, vertical migration, displacements, and local 119 survival/extinctions in given refugia as consequence of the Neogene climate fluctuations and 120 121 mountain building processes, leaving some imprints on their geographical patterns. In this study, 122 we will use maternally-inherited plastid DNA sequences and nuclear microsatellites or simple sequence repeats (nSSR), as well as ecological niche modeling (ENM), to unravel the 123 population dynamics of the HHM tree peonies (P. delavayi and P. ludlowii). ENM is often 124 employed to get insights into the paleodistribution of species, being an ideal complement to 125 genetic markers (Huang and Schaal, 2012). More specifically, we aimed to: (1) evaluate the 126 phylogenetic and relationships between P. delavavi and P. ludlowii; (2) reconstruct the 127 phylogeographic relationships among populations of both species; (3) estimate the levels of 128 intrapopulation genetic diversity, as well as the interpopulation genetic and phylogeographic 129 structure; (4) reveal the effects of complex landforms, the Quaternary climatic oscillations 130 (including the formation of ice-caps), and the last phases of mountain uplift on the genetic and 131 phylogeographic patterns; and (5) estimate the potential distribution both at present and in the 132

past and determine whether there is niche conservatism or niche divergence between the twospecies.

Data obtained in this study will shed light on the HHM tree peony diversification; such 135 information would be especially relevant for their conservation, as both P. delavavi and P. 136 ludlowii have been traditionally harvested for both horticultural (Zhang et al., 2011) and 137 medicinal (Hong, 2010; Yang et al., 2014) purposes. Although only P. ludlowii is considered 138 threatened (it is listed as 'endangered' in 2013 version of China's IUCN red list; MEP-CAS, 139 2013), only P. delavayi is, ironically, legally protected in China at national level (it is listed as 140 'third grade' in the National List of Rare and Endangered Plant Species of 1984 under the old 141 name P. delavayi var. lutea). In addition, large parts of P. delavayi's range (especially NW 142 Yunnan) are becoming major touristic hotspots of China. 143

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146 **2. Material and Methods**

147 2.1. Plant material and population sampling

A total of 17 wild populations of subsect. Delavayanae (13 populations of P. delavayi and 148 four populations of P. ludlowii) were sampled in the whole distribution area during the 149 flowering seasons from 1999 to 2006 (Tables 1 and S1). The number of sampled individuals 150 within each population ranged from two in small Tibetan populations to 15 individuals 151 depending on the number of patches and total size of studied populations. Since for P. delavayi 152 153 one patch usually consists of ramets of the same genet, only one sample per patch was collected. For the plastid DNA phylogenetic analysis (see below), we used P. jishanensis T. Hong & W. 154 Z. Zhao, P. ostii T. Hong & J. X. Zhang. and P. qiui Y. L. Pei & D. Y. Hong as outgroups 155 following Zhou et al. (2014). Voucher specimens were deposited in the PE herbarium. Young 156 leaves once collected were quickly dried in silica gel. 157

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159 2.2. DNA isolation, plastid DNA sequencing and microsatellite genotyping

Total genomic DNA was extracted from silica gel-dried leaves according to the mCTAB 160 method (Li et al., 2013). After preliminarily screening of seven chloroplast loci, the most 161 variable ones (*ndhC-trnV*^(UAC), *ndhF*, *psbA-trnH*, and *rps*16-*trnQ*) were finally selected as 162 markers for this study. The primers used and detailed PCR profiles are given in Table S2. 163 Polyethylene glycol (PEG8000) was used for purifying the PCR products. The fragments were 164 sequenced on an AB 3730xl DNA analyzer (Applied Biosystems Inc., Foster City, CA) using 165 both forward and reverse primers. Chloroplast sequences were further edited and assembled 166 using Sequencher v.4.6 (Gene Codes Corporation, Ann Arbor, MI) and adjusted manually. The 167 sequences were submitted to GenBank under accession numbers XXXXX-XXXXX (Table S1). 168 Each individual was genotyped at nine nSSR loci using the primer pairs designed by Wang 169 et al. (2009) and Zhang et al. (2011) (see Table S3). The 5' ends of the reverse primers were 170 labeled with one of the four fluorescent dyes (FAM, JOE, PET, or NED). The PCR products 171 labeled with different dyes were mixed together in equal ratio and 2.0 µL of the mixture was 172 combined with 7.7 µL of Hi-Di formamide and 0.3 µL of an internal size standard, GeneScan 173 600LIZ (Applied Biosystems Inc). The fragments were resolved on an AB 3730xl DNA 174 Analyzer (Applied Biosystems Inc). Fragment length in base pairs was calculated by 175 GeneMapper v.4.0 (Applied Biosystems Inc). 176

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178 **2.3. Genetic data analysis**

179 2.3.1. Estimation of genetic diversity

Genetic diversity was quantified based on both plastid DNA and nSSR data. Polymorphic 180 sites (S), number of unique haplotypes (h), haplotype diversity (H_d), nucleotide diversity (π) 181 and average number of nucleotide difference (k) were computed with DnaSP v.5.10.01 (Librado 182 and Rozas, 2009) using the combined plastid DNA sequence data. For the nSSR dataset, 183 percentage of polymorphic loci when the most common allele had a frequency of < 0.99 (P₉₉), 184 185 mean number of alleles per locus (A), observed heterozygosity (H_0), and unbiased expected heterozygosity or Nei's (1978) gene diversity (He) were calculated using GenAlEx v.6.5 186 (Peakall and Smouse, 2006). Allelic richness (AR; rarefacted to compensate for unequal sample 187 sizes; Hurlbert, 1971) was computed with FSTAT v.2.9.3 (Goudet, 1995). 188

Linkage disequilibrium between pairs of nSSR loci in each population was assessed using 189 FSTAT with significance determined using the Bonferroni correction (Rice, 1989). The 190 frequency of null alleles was estimated following the expectation maximization (EM) algorithm 191 of Dempster et al. (1977) using FreeNA (Chapuis and Estoup, 2007). To test whether 192 populations were under Hardy-Weinberg equilibrium, Wright's (1965) F_{IS} was estimated using 193 nSSR data following the method of Weir and Cockerham (1984) with Genetix v.4.05 (Belkhir 194 et al., 1996–2004). Statistical significance of F_{IS} values for each locus per population was tested 195 by permutation tests (10,000 randomizations), using the same program. Mean F_{IS} was 196 197 calculated by jackknifing over loci, with bootstrapping to obtain the 95% confidence interval.

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199 2.3.2. Phylogeographic and population structure analyses

Genetic relationships between haplotypes using concatenated sequences were depicted by a statistical parsimony network using TCS v.1.21 (Clement et al., 2000). The highly variable polyT, polyA, polyG, or polyC motifs in the sequences were excluded from the analysis, and indels longer than 1 bp were shortened to single base pair gaps and treated as a fifth character state.

205 Phylogenetic relationships among haplotypes were reconstructed using Bayesian inference (BI) with MrBayes v.3.1.2. (Ronquist and Huelsenbeck, 2003) and maximum parsimony (MP) 206 method with PAUP v.4.0a149 (Swofford, 2002) by treating gaps as missing values. The best-fit 207 model of substitution for each region was selected by means of jModelTest v.0.1.1 (Posada, 208 2008) for BI analyses. Bayesian analyses were initiated with random starting trees and four 209 Markov chains were run simultaneously for 10^6 generations and sampled every 100 generations. 210 The first 25% of the runs were discarded as burn-in. The 50% majority rule consensus tree and 211 posterior probabilities (PPs) of the nodes were calculated from the remaining trees. Nodes with 212 $PPs \ge 0.95$ were considered to be statistically supported. MP analyses were conducted by 213 heuristic searches with 10,000 replicates with random taxon addition. Bootstrap support (BS) 214 values were generated from 1000 replicates using simple heuristic search algorithm with 215 random sequence addition and TBR branch swapping. Parsimony-uninformative characters 216 were excluded from analysis. BS support \geq 70% was considered significant. 217

Pairwise genetic distance between populations based on nSSR dataset was calculated using two algorithms, Nei's (1972) standard genetic distance (D_S) and Rogers' (1972) distance (D_R). These distance matrices were converted into UPGMA (unweighted pair-group method using arithmetic averages) dendrograms (after 1000 bootstrap replicates) employing the programs
Populations v.1.2.30 (Langella, 1999) and TreeView v.1.6 (Page, 1996). A principal coordinate
analysis (PCoA), based on codominant genotypic distances, was also conducted using GenAlEx.

Population genetic structure was investigated with both markers using different approaches. 224 First, analysis of molecular variance (AMOVA) was used to examine the distribution of the 225 variance components of genetic diversity within and among populations, based on the complete 226 sample set and several nested analyses. These analyses were performed with Arlequin v.3.5.1.2 227 (Excoffier and Lischer, 2010) for plastid DNA data and with GenAlEx for nSSRs. Second, the 228 229 Bayesian algorithm implemented in Structure v.2.3.4 (Pritchard et al., 2000) was used with nSSR dataset. The admixture ancestry model with correlated allele frequencies was selected as 230 the most appropriate option for the analysis. The burn-in period and Markov chain Monte Carlo 231 (MCMC) were set to 50,000 and 500,000 iterations, respectively, and 20 replicates per K were 232 run. The most likely value of K was determined by the ΔK statistic of Evanno et al. (2005), with 233 the aid of Structure Harvester v.0.6.94 (Earl and vonHoldt, 2012). As the ΔK method tends to 234 235 identify K = 2 as the top level of hierarchical structure (Janes et al., 2017), it was combined with the method of choosing the smallest K after the log probability of data [ln Pr(X|K)] values 236 reached a plateau (Pritchard et al., 2010). Programs Clumpp v.1.1.2 (Jakobsson and Rosenberg, 237 2007) and Distruct v.1.1 (Rosenberg, 2004) were used to combine the results of the 20 replicates 238 of the best K and to to graphically display the results produced by Clumpp, respectively. And 239 third, Permut v.1.0 (Pons and Petit, 1996) was used with plastid DNA dataset for checking the 240 241 occurrence of significant phylogeographical structure by testing if G_{ST} (which only takes into account the allele frequencies) and N_{ST} (which uses the distance between different alleles) were 242 significantly different using 10,000 permutations. 243

The location of potential genetic barriers between populations was explored from both plastid DNA and nSSR datasets through the Monmonier's maximum-difference algorithm implemented in Barrier v.2.2 (Manni et al., 2004). The significance of barriers was tested for nSSR data by bootstraping 1000 Nei's genetic distances D_A (Nei et al., 1983) matrices that were previously obtained with Microsatellite Analyzer (MSA) v.4.05 (Dieringer and Schlötterer, 2003).

To test isolation-by-distance among populations, we estimated Wright's (1965) FST based 250 on nSSR loci following the method of Weir and Cockerham (1984) and Rousset (1997). 251 Statistical significance of all pairwise F_{ST} values was estimated with Genetix by permutation 252 tests (10,000 randomizations), whereas mean F_{ST} was calculated by jackknifing over loci, with 253 254 bootstrapping to obtain the 95% confidence interval. The correlation between the matrix of pairwise genetic differentiation $[F_{ST}/(1-F_{ST})]$ and the matrix of the log-transformed 255 geographical distances was computed by applying the Mantel test (Mantel, 1967) with 1000 256 permutations using IBDWS v.3.23 (Jensen et al., 2005). 257

To reveal recent (i.e., within the last several generations) gene flow between populations, we estimated migration rates based on our nSSR loci using the program BayesAss v.1.3 (Wilson and Rannala, 2003). We ran 3×10^{6} MCMC iterations, with a burn-in of 999,999 iterations and a sampling frequency of 2000 by setting delta at 0.15 (the default value).

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263 2.3.3. Population demographical dynamics

In order to detect deviations from the selective neutrality, Tajima's (1989) D, Fu and Li

265 (1993) F^* and D^* , Fu's (1997) F_s , and Ramos-Onsins and Rozas' (2002) R_2 were computed by 266 using DnaSP. Additionally, using the same software we calculated the mismatch distribution of 267 the pairwise differences between all individuals in a sample to test the population stability or 268 growth (Rogers and Harpending, 1992; Harpending, 1994) and the raggedness index HRI 269 (Harpending, 1994) between observed and expected mismatch distribution as an estimate of the 270 goodness-of-fit. The significance of all the indexes was tested by coalescent analysis using 271 10,000 replicates.

To test the recent dynamics of effective population size (bottlenecks), we ran the software Bottleneck v.1.2.02 (Piry et al., 1999) using our nSSR dataset for those populations with at least 10 individuals sampled (the minimum requirements of the program). Two different tests were used to check for bottlenecks: the sign test (Cornuet and Luikart, 1996) and the Wilcoxon signed-rank test (Luikart and Cornuet, 1998), both under the infinite allele model (IAM) and the stepwise mutation model (SMM).

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279 **2.4. Ecological niche modeling (ENM)**

ENM was performed to evaluate the potential distribution of Paeonia delavayi, P. ludlowii, 280 and P. delavayi + P. ludlowii. We employed the maximum entropy algorithm, as implemented 281 in MaxEnt v.3.3 (Phillips et al., 2006). The current distribution information for both species 282 was obtained from the sampling sites (Table 1), literature (e.g., Hong, 2010), and specimens 283 deposited in the main Chinese herbaria (www.cvh.ac.cn). After removing duplicate records 284 285 within each pixel (2.5 arc-min, ca. 5 km), we obtained 119 records of P. delavayi and 13 records of P. ludlowii. A set of 19 bioclimatic variables at 2.5 arc-min resolution covering the 286 distribution range (and neighboring areas) for both species under current conditions (1950-287 2000) were downloaded from the WorldClim website (www.worldclim.org; Hijmans et al., 288 2005). Of these, we selected a smaller set of eight relatively uncorrelated (r > |0.9|) variables 289 (see Supplementary Text 1). The distribution model under current conditions was projected to 290 the Last Glacial Maximum (LGM, ca. 21,000 yr BP) using palaeoclimatic layers simulated by 291 292 three widely used general circulation models (see Supplementary Text 1). For the cases with a 293 considerable number of occurrences (P. delavayi and P. delavayi + P. ludlowii), 20 replicates of MaxEnt (using the subsample method) were run, and model performance were assessed using 294 the area under the curve (AUC) of the receiver operating characteristic (ROC) plot with 25% 295 of the localities randomly selected to test the model. Given the low number of occurrences for 296 297 P. ludlowii (13), we used a methodology based on a jackknife (or 'leave-one-out') procedure to 298 test the model (Pearson et al., 2007), with the definitive model (i.e., using all occurrence points) running MaxEnt 20 times (using the bootstrap method) (see Supplementary Text 1). The 299 MaxEnt jackknife analysis was used to evaluate the relative importance of the eight bioclimatic 300 variables employed based on their gain values when used in isolation. All ENM predictions 301 were visualized in ArcGIS v.10.2 (ESRI, Redlands, CA, USA). 302

Niche similarity between *P. delavayi* and *P. ludlowii* was measured through two niche overlap indices, Hellinger-derived *I* and Schoener's *D* as implemented in ENMTools v.1.4.3 (Warren et al., 2010). Two quantitative tests of niche similarity—also implemented in ENMTools—were further used: the 'niche identity test' and the 'background test' (see Supplementary Text 1). For both tests, null distributions were generated from 100 pseudoreplicates. Finally, we estimated the niche breadth for each species by calculating the

- 309 inverse concentration statistic of Levins (1968), as implemented in ENMTools.
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311 3. Results

312 **3.1. Genetic diversity**

A total of 159 individuals from of Paeonia subsect. Delavayanae (137 individuals from 13 313 populations of P. delavavi and 22 individuals from four populations of P. ludlowii) were 314 sequenced for four loci of chloroplast genome, *ndh*C-*trn*V^(UAC), *ndh*F, *psb*A-*trn*H, and *rps*16-315 trnQ. We identified 14 chloroplast haplotypes in P. delavayi and one in P. ludlowii. Most 316 317 populations of *P. delavavi* had only one haplotype except DY-LIJ with three haplotypes, and DS-XIA and DY-WEI1 with two different haplotypes; the populations DT-BOM1, DT-BOM2, 318 DT-NYI and DY-DEQ shared the same haplotype (Fig. 1A). At species level, the values of 319 genetic diversity of *P. delavayi* were the following: $H_d = 0.90$, $\pi = 0.19 \times 10^{-2}$, and k = 4.98 (see 320 Table S4). No sequence variation was detected in *P. ludlowii*. 321

With nSSR, 155 individuals from the same 17 populations were genotyped, with all nine 322 surveyed microsatellites being polymorphic across populations. No significant linkage 323 disequilibrium was detected in any of the loci pairs. Only 14 of 90 of all the valid tests showed 324 a significant deviation from Hardy-Weinberg expectations (nine loci showed excess of 325 homozygotes whereas five loci showed excess of heterozygotes), although only one persisted 326 after Bonferroni correction (data not shown). The values of null allele frequency at all loci were 327 very low (all well below 0.100, with a mean of 0.039), indicating that null alleles are not 328 329 expected to cause significant problems in the analysis (cf. Dakin and Avise, 2004; Orsini et al., 2008). In fact, the differences between the 'raw' values of F_{ST} (one of the most sensitive 330 parameters when null alleles occur; Chapuis and Estoup, 2007; Chapuis et al., 2008) and those 331 after correcting for the presence of null alleles in our dataset were absolutely negligible (less 332 than 1%). 333

334 A total of 100 alleles were detected and the number of alleles per locus varied greatly among loci, from five (loci Jx05 and Pdel11) to 20 (Pdel05), with an average of 11.1 alleles. 335 The values of polymorphism were much higher for populations of *P. delavayi* ($P_{99} = 75.2$, A =336 2.402, AR = 1.731, $H_e = 0.369$) than those of P. ludlowii (P₉₉ = 5.6, A = 1.056, AR = 1.024, H_e 337 = 0.013; Table 2). The most variable populations within the study system were two from NW 338 Yunnan (DY-WEI2 and DY-XIG; Table 2). The four populations of P. ludlowii showed 339 extremely low levels of genetic diversity. In fact, three out of four populations were fixed for a 340 341 single genotype (Table 2). Private alleles were found for most P. delavayi populations (ranging 342 from one to six) but were absent in P. ludlowii ones (Table 2). Of the 11 alleles occurring in P. ludlowii, three were exclusive to this species, whereas the remaining eight were shared with 343 some of the P. delavayi populations. 344

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346 3.2. Phylogeographic and population structure analyses

The rooted TCS network illustrates the relationships among the 15 haplotypes (Fig. 1B). The network shows two large groups of haplotypes, with a weak geographic pattern. The first group included haplotypes H1, H9, H11–H13 and H15, corresponding to Tibetan populations (DT-BOM1, DT-BOM2 and DT-NYI) and Yunnan populations (DY-DEQ, DY-LIJ and DY-WEI1). The second group included haplotypes H2–H8, H10, and H14, corresponding to Yunnan (DY-DAL, DY-KUN, DY-WEI2 and DY-XIG) and Sichuan populations (DS-LIT, DS-MUL and DS-XIA) and also the four populations of *P. ludlowii*. In the phylogeny of the haplotypes (Fig. 2) we found a similar relationship to the network. The first group was weakly supported (PP = 0.77) and the second group was well supported (PP = 1.00; Fig. 2). The monophyly of *P. delavayi* was not supported by chloroplast data because *P. ludlowii* was nested to the second group. For details of phylogenetic analyses, see Table S5.

With nSSR markers, the UPGMA dendrograms based on Nei's (1972) genetic distance (D_s) and Rogers' (1972) distance (D_R) of nSSR had identical topologies, with the latter one showing somewhat higher bootstrap support at many branches (Figs. 3 and S1). Notably, the Tibetan populations of *P. delavayi* appeared as sister to the rest of populations, whereas *P. ludlowii* was sister to the populations of *P. delavayi* from Sichuan and Yunnan. A similar result was obtained in the PCoA (Fig. 4), with three clearly differentiated groups: *P. ludlowii* populations, the Tibetan populations of *P. delavayi*, and the remaining populations of *P. delavayi*.

The results of the nSSR's AMOVA also confirmed the highly significance and singularity of these three groups of populations (Table 3), as there was a considerable percentage of variance due to differences among these three clusters (38.5%) in nSSR data. However, the among-taxa component only accounted slightly higher, for up to 40.3% (and 42% in plastid DNA) of the total variance.

The results of Structure analyses for all individuals based on the nSSR markers are 370 represented in Figs. 5 and S2. According to Evanno's approach, K = 2 and 13 were the most 371 likely numbers of genetic clusters, whereas, when the $\ln Pr(X|K)$ was plotted, the 'plateau' was 372 373 approximately reached at K = 11-13 (Fig. 5). Despite the highest peak for ΔK was at K = 2, genetic clustering was highly different between the 20 runs (in only three runs clusters and 374 species coincided; Fig. S3); the two-cluster structure was, therefore, not considered as the most 375 biologically meaningful, although the fact that the four populations of P. ludlowii appear 376 together in all runs (Fig. S3) might be of relevance. At K = 13, in contrast, a highly stable 377 population structure was found, with 19 of 20 simulations showing the same pattern: each 378 population having its own cluster, with the exception of the four populations of *P. ludlowii* and 379 the three Tibetan populations of P. delavayi (each group having its own cluster), and populations 380 381 DY-LIJ and DY-XIG (these latter showing partial membership to multiple clusters). With the Permut analysis with our plastid DNA dataset, the $N_{\rm ST}$ (0.912 ± 0.054) was not significantly 382 higher than G_{ST} (0.921 ± 0.048), suggesting the absence of phylogeographical structure. 383

The analysis of genetic barriers based on Monmonier's algorithm with nSSR dataset 384 suggested that strong obstacles to gene flow would exist between Tibetan populations of P. 385 delavavi and P. ludlowii (Fig. 6A). The first barrier mainly separated Tibetan Nyingchi 386 population (DT-NYI) from the four P. ludlowii populations (green barrier, with 93% bootstrap 387 support), whereas the second one separated populations of P. ludlowii from the other Tibetan 388 populations of P. delavayi (DT-BOMI1 and DT-BOMI2, blue barrier, 90% BS). The third barrier 389 (in red) was weaker (with BS below 80%) and mainly contributed to the separation of most of 390 the populations in the eastern range of P. delavavi, as well as DY-DEQ from DT-BOM2 (i.e., 391 the eastern from the western range of *P. delavavi*). The significance of these separations 392 increased with the addition of new barriers, and no additional separations appeared until the 393 10th barrier was added (Fig. 6A). The genetic barriers revealed by plastid DNA data suggested 394 395 a relatively different scenario for the eastern populations but not for the western ones (see Fig. 396 6B).

The genetic differentiation among populations of *P. delavayi* based on F_{ST} was very large [mean F_{ST} (13 populations) = 0.510, 95% CI = 0.451–0.584; see also Table 4]. In contrast, the genetic differentiation between populations of *P. ludlowii* was very low [mean F_{ST} (four populations) = 0.014, 95% CI = -0.002–0.027]. The F_{ST} value between the two species was 0.458. The Mantel test performed for *P. delavayi* revealed a significant positive correlation between genetic and geographic distances ($r^2 = 0.190, P = 0.001$).

The BayesAss analysis revealed an almost total absence of recent gene flow between populations. Of the 272 pairwise estimates, only three indicated evidence of recent gene flow between populations: from DS-MUL to DS-XIA (m = 0.166), from DY-WEI1 to DS-LIT (m =0.223), and from DY-XIG to DY-LIJ (m = 0.119) (Table S6).

407

408 **3.3. Population demographical dynamics**

The mismatch pairwise distance analyses of total plastid DNA dataset (*P. delavayi* + *P. ludlowi*) and at specific level (*P. delavayi*) showed multimodal patterns that suggested dynamical equilibrium (Fig. S4). In addition, no significance was found in any of the studied indexes to detect deviation in the selective neutrality (Table S4). However, DY-LIJ and DY-WEI1 populations showed unimodal mismatch distributions that could be interpreted as these populations have experienced past demographic expansions (Fig. S4).

Bottleneck test results using nSSR data suggested that two populations (DS-MUL, DS-LIT, and DS-XIA) have suffered recent decreases in effective population size (Table 2). However, these results should be taken with extreme caution, as our sample sizes (N = 10-15) and number of polymorphic loci (2–9) are below the recommended threshold (20 individuals and 20 loci) to ensure enough statistical power (Cornuet and Luikart, 1996).

420

421 **3.4. Ecological niche modeling**

The AUC scores averaged across 20 runs were very high for both P. delavayi and P. 422 delavavi + P. ludlowii (mean \pm SD, 0.937 \pm 0.010 and 0.935 \pm 0.013, respectively), which 423 424 supported the predictive power of the model. The model for P. ludlowii also performed 425 reasonably well, as we found a high success rate (0.846) and statistical significance (P < 0.001) in the jackknife test. According to the MaxEnt jackknife tests of variable importance, the mean 426 temperature of the coldest quarter (bio11) was the most informative for the three models. 427 Although probability maps largely differed among models, there was a general loss of suitable 428 range for both P. delavayi and P. delavayi + P. ludlowii at the LGM when compared to the 429 430 present (Figs. 7 and S5; Table S7); such loss was mainly focused on southeastern areas (mainly Yunnan) for the MIROC model but on northwestern areas (Tibet and Arunachal Pradesh of 431 India) for the MPI model (Fig. S5). Variability among models was even more evident in P. 432 ludlowii, although they should be interpreted with extreme caution given the uncertainty of 433 projecting into the past with a small number of occurrences. The LGM models also indicated a 434 slight decrease in the mean altitude of the suitable habitats compared to the present (Table S7). 435

Niches of *P. delavayi* and *P. ludlowii* were not identical (Figs. S6A and S6B), although we
should take into account that niche identity tests can be seriously biased by the environmental
differences that exist between the regions in which two species do not overlap (Warren et al.,
2010). While background test overcomes this limitation, results were apparently contradictory:
when we compared *P. delavayi* occurrences to *P. ludlowii* background, the observed niche

441 overlap was significantly larger than the null distribution of niche overlap (P < 0.05) for D (but

442 not for *I*, which showed no differences; Figs. S6C and S6D). When the test was performed in

the inverse direction, observed *I* was significantly smaller than the expected by the null model (P = 0.00), while differences were not significant for *D* (Figs. S6E and S6F). Finally, the niche

breadth estimate based on ENMs was about five times larger for *P. delavavi* than *P. ludlowii*

446 (0.185 and 0.036, respectively).

447

448 **4. Discussion**

449 4.1. Relationships between *P. delavayi* and *P. ludlowii* and differentiation within *P. delavayi*

Our genetic data based on nSSR and plastid DNA do not support the current taxonomic 450 treatment of Hong et al. (1998) that P. delavavi and P. ludlowii are different species. First, the 451 populations of *P. ludlowii* are nested within *P. delavayi* in the UPGMA trees (Figs. 3 and S1); 452 second, the exclusion of K = 2 as the best clustering in the Structure analyses (Figs. 5 and S3); 453 and third, phylogenetic tree inferred from plastid DNA indicate that *P. ludlowii* is sister to two 454 Sichuan populations of P. delavayi, in a clade placed within a polytomy (Fig. 2). The recent 455 phylogeny of Zhou et al. (2014), which included all the species of section Moutan and 456 traditional cultivars, showed that P. delavayi and P. ludlowii are not distinguishable based on 457 14 chloroplast regions but, on the contrary, could be delimited based on 25 single-copy nuclear 458 459 markers. Leaving aside the discrepancy between nuclear and plastid markers (not uncommon in angiosperms; e.g., Zhang et al., 2015), the implications derived from the phylogeny of Zhou 460 461 et al. (2014) should be treated with caution as only a single individual from a small subset of the extant populations of both P. delavayi and P. ludlowii were included. 462

Despite the lack of clear boundaries between P. delavayi and P. ludlowii, there are some 463 signals of genetic distinctiveness, including the PCoA grouping patterns (P. ludlowii appears as 464 a relatively isolated entity; Fig. 4), the Structure clustering patterns (the four populations are 465 clustered together along all surveyed K values; Fig S2), and the AMOVA analyses (with the 466 among-taxa component showing a non-negligible value of 40.3%; Table 3). Such results could 467 likely reflect an ongoing differentiation process, apparently following the classical progenitor-468 derivative (P-D) model of speciation. Under such scenario, D taxon has budded off and 469 acquired new traits (or become fixed for features within the polymorphism of P species) while 470 P species remained largely unchanged (Gottlieb, 2003; Crawford, 2010). In addition to 471 molecular phylogenies (the expected pattern is D taxon to be nested within populations of the 472 473 P species; Crawford, 2010, see above), morphological traits, allelic/genetic diversity patterns, 474 and ecological modeling are in agreement with what is expected for P–D model of speciation. While some morphological traits of P. ludlowii (the D taxon) can be considered new compared 475 to P. delavayii (the P species), such as the number of carpels, others could be the result of the 476 fixation of a given trait among the variability of the P species (e.g., pure yellow corollas). The 477 lack of stolons in *P. ludlowii* can also be viewed as the fixation of a very rare trait within *P.* 478 *delavavi* (as the vast majority of populations shows a vigorous vegetative propagation; Hong, 479 2010). Regarding genetic diversity, P. ludlowii shows only a subset of the variability of the 480 putative progenitor (values of P₉₉, A, AR and H_e are much higher for P. delavayi compared to P. 481 ludlowii; Table 2) with few unique alleles (three exclusive alleles for P. ludlowii compared to 482 89 for P. delavayi). 483

484 Niche similarity tests are suggestive that there is at least some degree of niche divergence

of P. ludlowii with respect to P. delavayi. According to the background tests, the niche model 485 of P. delavayi can predict the niche of P. ludlowii (i.e., niche conservatism), but the model of P. 486 ludlowii is not able to predict that of P. delavayi (i.e., niche differentiation, Fig. S6), which can 487 also be visualized in the habitat suitability maps (Fig. 7). Such asymmetrical niche 488 differentiation may be due to limited availability of the preferred environmental conditions 489 within the range of the species showing greater differentiation than expected (e.g., Culumber et 490 al., 2012) or to some degree of habitat selection or specialization within a range of 491 environmental conditions (e.g., Sackett et al., 2014). Compared to P. delavayi, P. ludlowii has 492 493 a much narrower niche (its niche breadth is below one-fifth that of P. delavavi), whereas its current populations tend to occur at higher elevations (altitude = 3588 ± 429 m vs. 3010 ± 637 494 m), at colder (mean annual temperature = 6.9 ± 2.4 °C vs. 9.3 ± 3.8 ; mean temperature of coldest 495 guarter = -5.8 ± 2.5 °C vs. 2.5 ± 4.2 °C) and at clearly much drier (annual precipitation = 603 496 \pm 100 mm vs. 882 \pm 154 mm) sites. This putative specialization of *P. ludlowii* toward the harsher 497 upland Tibetan conditions is fully compatible with the scenario of a P–D model of speciation. 498

499 Instead of remaining unchanged, as stated by the classical definition of a P–D species pair, *P. delavayi* itself seems to be in process of differentiation, which is mainly (but not exclusively) 500 allopatric. As noted in the Introduction, none of the multiple variants of P. delavayi (often 501 described as subspecies or varieties, or even as 'species') has merited taxonomic recognition 502 (Hong et al., 1998, Hong, 2010). We believe that the variations of the morphological characters 503 that were used to define these entities might represent, however, the fixation of a given character 504 505 (within the polymorphism of the species) at the local level [e.g., yellow petals in Eryuan of NW Yunnan (P. lutea), or withish petals around Weixi of NW Yunnan (P. weisiensis)]. Genetic data, 506 rather than morphology, unambiguously indicate the onset of an ongoing speciation process 507 within P. delavayi: indeed, the Tibetan populations of P. delavayi as a whole are in an advanced 508 process of divergence, as shown in all our genetic analyses (Figs. 3, 4, 5 and S1), and may merit 509 recognition as a distinct evolutionary lineage. In addition, the Sichuan and Yunnan populations 510 are also showing clear signals of genetic divergence among them; notably, genetic structure 511 Bayesian approaches indicate that, as a general norm, each population has its own cluster. F_{ST} 512 513 values among populations are very high for nuclear (around 50%; Tables 3 and 4) markers and extremely so for chloroplast markers (98%; Table 3). Although divergence based on plastid 514 DNA was only slightly lower, F_{ST} values based on nSSR were much lower (0.302 vs. 0. 510) 515 for another tree peony, Paeonia rockii (Yuan et al., 2011, 2012); this is a very pertinent 516 comparison given that the same set of nSSR were used, although it should be taken into account 517 518 that P. rockii reproduces exclusively by sexual means (Chen et al., 1997). A series of life-history traits, in addition to topographical isolation (the Hengduan Mountains have an extremely 519 complex topography, with elevation gradients of up to 5000 m), might act as stimuli for the 520 incipient allopatric differentiation: (1) a general poor performance of sexual reproduction 521 observed in the field (despite that P. delavavi may be outcrossing; Li et al., 2014), (2) lack of 522 adaptation of seeds to long-distance dispersal, and (3) the (likely) preponderance of clonal 523 propagation (Hong et al., 1998; Hong, 2010; Li et al., 2012). The isolation-by-distance found 524 among P. delavayi populations is also supporting this pattern. 525

- 527 **4.2.** Genetic diversity and phylogeography: the role of Hengduan Mountains as a refugium
- 528 for subsect. *Delavayanae*

The high richness of plant species in China, especially the overrepresentation of relict 529 lineages, has been ascribed to the existence of large refugial areas (partly due to the lack of 530 Pleistocene extensive glaciations; López-Pujol et al., 2011; Huang et al., 2015). Although large 531 glaciers existed in the QTP, these never formed a unified ice sheet such as Fennoscandia and 532 Laurentide ones (Li et al., 1991; Shi, 2002; Owen et al., 2008; Kirchner et al., 2011). The low-533 altitude parts of the Hengduan Mountains mostly remained ice-free, especially at its 534 southernmost section (Li et al., 1991; Shi, 2002). Indeed, the Hengduan Mountains are regarded 535 as a Pleistocene glacial refugium (Zhang et al., 2009b; López-Pujol et al., 2011; Qiu et al., 2011; 536 537 Liu et al., 2012; Tang, 2015); even very relic elements, such as *Cunninghamia lanceolata* Lamb, Davidia involucrata Baill. or Taiwania cryptomerioides Hayata, found a refugium in the 538 southern Hengduan Mountains, where temperature and precipitation kept relatively high even 539 during the LGM (e.g., Jiang et al., 2011; Lu et al., 2013; Liu and Jiang, 2016; Tian and Jiang, 540 2016). Despite that levels of genetic diversity of P. delavavi as a whole are lower than those of 541 Paeonia rockii (Yuan et al., 2012), it should be taken into account that the latter occurs in an 542 543 area (Qinling and central China ranges) that was never glaciated and was home of some of the most important glacial refugia of China (e.g., López-Pujol et al., 2011; Huang et al., 2015). 544 Indeed, if we consider only those populations located in areas that were ice-free or almost ice-545 free (those at latitudes below 28°N), then the levels of genetic diversity of P. delavavi 546 populations ($H_e = 0.448$) are closer to those of *Paeonia rockii* ($H_e = 0.498$; Yuan et al., 2012) 547 and even higher to those expected for endemic species ($H_e = 0.420$; Nybom, 2004). 548

549 Judging from the values of polymorphism included in Table 2, it seems that the extent of glaciers and ice caps at the LGM within the Hengduan Mountains would have played a major 550 role in the genetic variability harbored by the populations of *P. delavavi*, as found for other 551 regional endemisms [e.g., Sinopodophyllum hexandrum (Royle) T.S. Ying; Li et al., 2011]. We 552 believe that this is a reliable assumption given that microsatellites, mainly due to their high 553 mutation rates and high incidence of homoplasy, are generally only suitable to reveal events on 554 timescales of just several thousands of years (Jarne and Lagoda, 1996). Almost all populations 555 located in the south section of Hengduan Mountains (DY-XIG, DY-DAL, DY-WEI2, and DY-556 557 LIJ)—that was relatively ice-free at the LGM (Fig. 8), would have been refugial for P. delavayi, because they show a large number of alleles (including private ones) and show the highest H_e 558 values (Table 2); in fact, mean H_e values of these four populations plus DY-WEI1 ($H_e = 0.460$, 559 SD = 0.067) are much larger as a whole than those of the four populations (DY-DEQ, DS-XIA, 560 DS-LIT, and DS-MUL) located in the northern section of the Hengduan mountains, much more 561 affected by glaciations ($H_e = 0.314$, SD = 0.071). The Lijiang-Xianggelila region probably 562 harbored the most important refuges for the taxon, as these are the only two populations (DY-563 LIJ and DY-XIG) that show partial membership to multiple clusters at K = 13 (Fig. 5). 564 Xianggelila, in addition, is the population harboring the highest number of alleles, while almost 565 all described petal colours have been observed there (which could be the result of secondary 566 contact of yellow and deep-pink flowered morphotypes; Hong et al., 1998; Hong, 2010; Zhang 567 et al., 2011). Pollen records indicate the presence of broad-leaved forests [Quercus L., Betula 568 L. and Castanopsis (D. Don) Spach; Yao et al., 2015] in the area, suggesting certain level of 569 climatic stability. Under such scenario, populations would have maintained relatively large 570 sizes while keeping certain levels of gene flow. 571

572 The northern populations (i.e., DY-DEQ, DS-XIA, DS-LIT, and DS-MUL) of *P. delavayi*

probably escaped the coldest periods of the Pleistocene by means of downward migrations to 573 574 the adjacent valleys. All these populations, although showing low levels of heterozygosity, harbor exclusive allele variants (Table 2), suggesting in situ persistence. Glaciers here were 575 more extensive (Fu et al., 2013) and populations, perhaps with the single exception of DS-XIA 576 (which has the highest heterozygosity among this group; Table 2), would have directly been 577 affected by LGM ice caps (as populations were located on the surroundings of these; Fig. 8). 578 Moreover, the northern part of the Hengduan Mountains was colder and drier than the southern 579 one, even showing more differences than at present (Jiang et al., 2011; Tian and Jiang, 2016), 580 581 which would have favoured the persistence of the taxon in microrefugia instead of macrorefugia (Rull, 2009). On the contrary, the nSSR data suggest that the Tibetan populations of P. delavayi 582 might be the result of recent recolonizations (perhaps after the LGM): both their levels of 583 heterozygosity and number of alleles are the lowest within the taxon, and have no exclusive 584 alleles (with the exception of DT-BOM1; Table 2). The mountain range that dominated this area, 585 Nyaingentanglha, was extensively glaciated at the LGM, and the three populations (DT-BOM1, 586 587 DT-BOM2, and DT-NYI) are located in an area that was supposedly covered by ice caps at the LGM (Li et al., 1991; Fig. 8). 588

ENM results are equivocal, given that the three models (CCSM, MIROC, and MPI) 589 indicate a very different scenario for P. delavavi at the LGM (Figs. 7 and S5). The MIROC 590 model seems to be unrealistic, as most of the reconstructed suitable areas occur in regions that 591 were heavily glaciated (Li et al., 1991; Shi, 2002; Fig. 8) and the 'lost' area compared to the 592 present corresponds to the more polymorphic populations (Fig. S5). Both the CCSM and MPI 593 594 models (especially the latter) basically agree with the continuous presence of the taxon in southern macrorefugia (Figs. 7 and S5), a scenario compatible with the mismatch analysis (Fig. 595 S4). The ENM also indicates a slight decrease in the mean altitude of the suitable habitats for 596 P. delavayi at the LGM compared to the present (of about 300 m; Table S7). It is agreed that 597 mountain species would have tracked the Pleistocene climatic oscillations by means of 598 altitudinal changes; admixture as consequence of downward migrations in the colder periods 599 600 would have blurred genetic footprints of allopatric divergence, as it is often reported in plants inhabiting the mountains of the Mediterranean Basin (Nieto-Feliner, 2014; Jiménez-Mejías et 601 al., 2015) and the Korean Peninsula (Chung et al., 2017). For P. delavavi, the large altitudinal 602 gradients generally prevented populations from secondary contacts among closely-located 603 populations (with some exceptions, e.g., Xianggelila) and, therefore, did not compensate the 604 strong geographical isolation that is found at present (there is almost a total absence of ongoing 605 606 recent gene flow; Table S6). Low seed production and seedling establishment, aggravated by the lack of adaptation to long-distance dispersal of P. delavayi seeds (Hong, 2010) might have 607 608 contributed to avoid wider altitudinal displacements.

The haplotype distribution within and among P. delavavi populations also indicates that 609 this taxon survived in multiple refugia in the Hengduan Mountains, in agreement with 610 microsatellites. Most populations are fixed for a single haplotype and, with the exception of H1, 611 no haplotypes are present in more than one population (Fig. 1A). Judging from the high number 612 of missing haplotypes, local extinction would have also been common, as expected for a region 613 that was partially covered by ice caps. However, plastid DNA may also reflect older events; 614 extinction of haplotypes and generation of new ones would have also been the result of 615 mountain building episodes, some of which took place until very recent phases of the 616

Pleistocene. Although there is still much controversy regarding the tempo and pace of the major 617 uplift events of the OTP, there is relative consensus on the fact that Hengduan Mountains are 618 of very recent origin, probably with abrupt upliftings during the Pliocene (Favre et al., 2015) 619 and probably extending into the Pleistocene (e.g., Li and Fang, 1999). However, as no mutation 620 rates are available for plastid DNA of Paeonia, we are not able to distinguish whether the 621 phylogeographic patterns are mainly attributable to geologic or to climatic events. The oldest 622 divergence event detected in our cpDNA network (that between haplotypes H1 and H7; Fig. 623 1B) may even pre-date the Hengduan Mountains uplift. 624

625 In agreement with microsatellites, the haplotype architecture of P. delavavi is also suggesting that the three Tibetan populations (DT-BOM1, DT-BOM2, and DT-NYI) could have 626 been the result of migration from warmer places (most probably at lower elevations in the 627 Hengduan Mountains), after the retreat of the glaciers that almost completely covered the 628 eastern section of Nyaingêntanglha range (where the three populations are located; see Fig. 8). 629 Regarding *P. ludlowii*, all the studied populations are fixed for a single haplotype (H2) that, 630 according to the haplotype network, it is a derived one. However, the fact that the H2 haplotype 631 appears as basal to haplotypes H4 and H5 in the Bayesian tree (Fig. 2) and the many mutational 632 steps (9–11) that separates H2 from its closest haplotypes both indicate a long isolation of these 633 populations, probably mainly driven by genetic drift. Such isolation is also indicated by the 634 Barrier analyses (Fig. 6), and would have been accompanied by a process of ecological 635 specialization that it is almost complete. The fixation of some morphological characters (e.g., 636 637 low number of carpels and lack of stolons) should be viewed in the context of this speciation process (Liu et al., 2014). The extremely low levels of hetezygosity and allelic richness as 638 revealed by nSSR are expected for a scenario of long-term persistence [in or around 639 (micro)refugia] within a heavily glaciated area (Fig. 8) under extreme cold and dry conditions 640 (e.g., Tian and Jiang, 2016), in which populations were probably small and isolated. Therefore, 641 we can propose a dual model for subsect. Delavayanae, in which the two main hypotheses of 642 Ouaternary history of OTP plant species are not mutually exclusive: (i) in situ persistence in 643 local refugia in the QTP and (ii) tabula rasa (retreat to SE refugia in glacial periods followed 644 by recolonization in postglacial ones) (Qiu et al., 2011; Liu et al., 2014). Such a double scenario 645 has also been reported in other regional endemics including Sinopodophyllum hexandrum 646 (Royle) T.S. Ying (Li et al., 2011), Lepisorus clathratus Ching (Wang et al., 2011), or Anisodus 647 tanguticus Pascher (Wan et al., 2016). 648

649

650 **4.3. Conclusions: evolutionary and conservation remarks**

The HHM region, but especially the Hengduan Mountains, is probably the largest 651 'evolutionary front' of the world's North Temperate Zone (López-Pujol et al., 2011). Some of 652 the most amazing plant radiations are taking place there, with lineages in which dozens to 653 hundreds of new species have arisen in the last million years (Wen et al., 2014; Hugues and 654 Atchison, 2015). For some extreme cases of 'rapid' radiations, each species can be limited to a 655 single mountain (Zhang et al., 2009b). In contrast to some of these spectacular examples, such 656 as Pedicularis or Saussurea, our case study (subsect. Delavayanae) could be regarded as a sort 657 of a 'slow', gradual radiation, and this 'slowness' in its diversification could be partly due to 658 the demographic dynamics of the species; the population viability of both P. delavayi and P. 659 ludlowii seems to mainly rely on longevity (individuals of both taxonomic entities usually reach 660

15-20 years; Yang et al., 2007; Li et al., 2012) which makes recruitment (which is rare in the 661 662 field: He, 2008; Hong, 2010) relatively unnecessary, especially for the case of *P. delavavi*. Our results indicate that at present at least three entities are clearly recognizable on the genetic 663 grounds (what is known as P. ludlowii, the Sichuan/Yunnan populations of P. delavavii, and the 664 Tibetan populations of *P. delavayi*), with the two latter not morphologically recognizable yet. 665 Carrying out additional studies may help to understand the ongoing speciation process, and 666 these may include (1) examining the stability of the diagnostic morphological traits (through 667 common-garden experiments); and (2) seeing, through comparative pollination studies, whether 668 pollinators have effectively influenced the fixation of floral and reproductive traits and thus, the 669 extent (if any) of pollinator-driven ecological speciation. Indeed, the preliminary results of 670 Shuai and Zang (2016) suggest that *P. ludlowii* has a wider spectrum of pollinators—with higher 671 frequencies of visits-compared to P. delavayi, and that insects show some degree of 672 phenotypic selection regarding several floral traits (e.g., petal length and petal width) which is 673 also variable between the two taxa. Although further studies are needed, such results might 674 675 indicate that pollinators play some role in the process of ecological differentiation of *P. ludlowii*. The rarity of *P. ludlowii* and high evolutionary potential of *P. delavavi* imply high priority 676 of in situ conservation of both taxa. Considering the high genetic differentiation among 677 populations of *P. delavayi* and variable morphology, and given that this taxon is under active 678 speciation, as many populations as possible should be conserved. Its extensive harvest due to 679 the medicinal properties (Hong, 2010; Yang et al., 2014), as well as its habitat fragmentation 680 681 (mainly as consequence of tourism growth, road construction, and economic development in general; e.g., Gu et al., 2013; Ye et al., 2015), if not banned or stopped, might contribute to 682 further increases in genetic differentiation, as it has been reported in the literature (e.g., Cruse-683 Sanders and Hamrick, 2004; Chung et al., 2014). As for further conservation efforts for P. 684 *delavavi*, these should be directed towards the most polymorphic populations (and putatively 685 contact zones), which, based on our results, are located in NW Yunnan: the axis Lijiang-686 Xianggelila, Weixi, and Dali. 687

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698 Appendix A. Suplementary material

699 Supplementary data associated with this article can be found, in the online version at

- 700 http://xxxxxxxxxxxxxxxxxxxx
- 701702 References
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- 989

- 990 LEGEND OF FIGURES
 991 Fig. 1. (A) Map showing the location of the studied populations of *Paeonia delavayi* and *P. ludlowii* (population codes are given in Table 1) and the distribution of plastid DNA
 993 haplotypes. Pie sizes are proportional to the haplotype frequency. (B) Haplotype network
 994 constructed by TCS v.1.2.1. The small open circles represent missing haplotypes. The size
 995 of coloured circles is approximately proportional to the observed frequency of haplotypes.
 996 * = outgroup position.





Fig. 2. Majority-rule consensus tree from a Bayesian analysis of the concatenate sequences of
 plastid DNA of *Paeonia* subct *Delavayanae* with Bayesian posterior probabilities indicated
 below branches and bootstrap values above branches. Supported branches are indicated in
 bold.





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0.01

- Fig. 3. Unweighted pair-group method using arithmetic averages (UPGMA) dendrogram using
 Rogers' (1972) distance (D_R) of 17 populations of *Paeonia* subsect. *Delavayanae* based on
 nSSR data. Numbers above branches represent bootstrap support for 1000 replicates; only
 values equal to or greater than 50% are given. *Paeonia delavayi* populations are in bold,
 those of *P. ludlowii* in italics.



Fig. 4. Principal coordinate analysis (PCoA) of 17 populations of *Paeonia* subsect.
 Delavayanae based on nSSR, using codominant genetic distances. *Paeonia delavayi* populations are in bold, those of *P. ludlowii* in italics.



1018 **Fig. 5.** Results of Structure analysis for all individuals of *Paeonia* subsect. *Delavayanae* 1019 studied, based on nSSR data. (A) The most likely *K* was estimated by choosing the smallest 1020 *K* after the log probability of data [ln Pr(X|K)] values reached a plateau (Pritchard et al., 1021 2010), and (B) the ΔK statistic of Evanno et al. (2005). (C) Assignation of individuals to 1022 genetic clusters at K = 13.

1023





Fig. 6. Schematic representation of the first nine barriers detected in *Paeonia* subsect. *Delavayanae* using the program Barrier based on nSSR (A) and plastid DNA (B). For the nSSR, results are based on 1000 bootstrap matrices of Nei et al. (1983) genetic distance (DA). Numbers indicate bootstrap support (after nine barriers). The location of populations is indicated by dots (full dots, *P. delavayi*; empty dots, *P. ludlovii*) whereas the polygons result from the Voronoi tessellation.



Fig. 7. Comparison of potential distributions as probability of occurrence for *Paeonia delavayi*, *P. ludlowii*, and *P. delavayi* + *P. ludlowii* using MaxEnt (Phillips et al., 2006), at the present time and at three climatic scenarios of the Last Glacial Maximum (LGM, ca. 21,000 years BP). The maximum training sensitivity plus specificity logistic threshold has been used to discriminate between suitable (gray-shaded) and unsuitable habitat. The darker color indicates a higher probability of occurrence. Black circles and black squares represent extant occurrence points of *P. delavayi* and *P. ludlowii*, respectively.



1042Fig. 8. Reconstruction of the extension of the Last Glacial Maximum (LGM, ca. 21,000 years1043BP) glaciers and ice caps in the study area, following the original map of Li et al. (1991).1044The value of expected heterozygosity (H_e ; red bar) and the number of alleles (blue bar) are1045given for each population. In light blue, total number of alleles (TA), in dark blue, number1046of private aleles (PA).



Population	Locality (village, country, province)	Latitude (N), longitude (E)	Altitude (m)	Plastid DNA (N)	nSSR (V)
abbreviation*					
P. delavayi					
DS-LIT	Maiwa, Litang, Sichuan	29°01'N, 100°38'E	3153	15	15
DS-MUL	Shawan, Muli, Sichuan	28°54'N, 100°55'E	3486	15	15
DS-XIA	Baiyi, Xiangcheng, Sichuan	28°54'N, 99°34'E	3175	15	10
DT-BOM1	Guxiang, Bomi, Tibet	29°55'N, 95°45'E	2600	5	5
DT-BOM2	Sumzom, Bomi, Tibet	29°44'N, 96°01'E	3100	2	2
DT-NYI	Zanba, Nyingchi, Tibet	29°35'N, 94°15'E	3000	4	4
DY-DAL	Cangshan, Dali, Yunnan	25°55'N, 100°02'E	2564	15	15
DY-DEQ	Mingyong, Dêqên, Yunnan	28°29'N, 98°56'E	2850	15	15
DY-KUN	Xishan, Kunming, Yunnan	24°57′N, 102°38′E	2000	15	15
DY-LIJ	Maoniuping, Lijiang, Yunnan	27º11'N, 100º16'E	2690	6	6
DY-WEI1	Duoduo, Weixi, Yunnan	27°42'N, 99°02'E	3483	8	8
DY-WEI2	Laboluo, Weixi, Yunnan	27°40'N, 98°50'E	3720	9	5
DY-XIG	Hala, Xianggelila, Yunnan	27°57'N, 99°35'E	3200–3300	13	15
P. ludlowii					
LT-MAI1	Zhare, Mainling, Tibet	29°12'N, 94°18'E	2980	10	10
LT-MAI2	Jinxuega, Mainling, Tibet	29°30'N, 94°48'E	2900	9	9
LT-MAI3	Gangga, Mainling, Tibet	29°18'N, 94°24'E	2900	2	2
LT-MAI4	Between Gangga and Mainling, Mainling, Tibet	29°12'N, 94°12'E	3000	4	4
Total				159	155

Table 1. Details of sampled populations of *Paeonia delavayi* and *P. ludlowii*. *N* = number of individuals investigated. * The population abbreviation consists of

TABLES

Table 2. Genetic divers loci when the most con individuals); $TA = total$ observed heterozygosit	ity parameter mon allele h number of i y; H_e = aunbi	s and fixatio ad a frequen alleles; <i>PA</i> = ased expecte	n index in] icy of < 0.9 number of ed heterozy	17 populations of 9; <i>A</i> = mean nu f private alleles gosity or Nei's	of Paeonia umber of a RA = nv (1978) ge	subsect. <i>Delava</i> lleles per locus; umber of rare all ne diversity; F_{IS}	<i>yanae</i> based on <i>AR</i> = allelic rich eles (those occu = fixation index	nSSR. P ₉₉ ⁻ mess (adjuu urring at fre . SE = stan	 percentisticate for a sted for a squencies dard erro 	age of poly sample si below 0. r.	ymorphic ze of two 05 ; $H_0 =$
Population symbol	P_{99} (%)	V	AR	TA/PA/RA	H_0	$H_{\rm e}({ m SE})$	$F_{ m IS}{}^1$	[A]	M ²	SM	M^2
P. delavayi											
DS-LIT	88.9	2.222	1.675	20/5/1	0.283	0.347 (0.071)	0.193*	0.111	0.020	0.414	0.191
DS-MUL	55.6	1.667	1.439	15/5/0	0.346	0.240(0.085)	-0.469***	0.146	0.031	0.214	0.047
DS-XIA	88.9	2.222	1.770	20/3/0	0.279	0.396 (0.079)	0.310*	0.026	0.020	0.040	0.098
DT-BOM1	66.7	1.667	1.577	15/1/0	0.422	0.328 (0.085)	-0.333^{ns}	I	I	I	I
DT-BOM2	44.4	1.444	1.444	13/0/0	0.333	0.296 (0.117)	-0.200^{ns}	I	I	I	I
DT-NYI	44.4	1.444	1.405	13/0/0	0.176	0.233 (0.093)	0.296^{ns}	I	I	I	I
DY-DAL	77.8	3.222	1.926	29/6/6	0.360	0.429~(0.094)	0.165*	0.579	0.289	0.587	0.594
DY-DEQ	88.9	2.222	1.505	20/1/3	0.210	0.272 (0.083)	0.320^{ns}	0.529	0.578	0.340	0.844
DY-KUN	77.8	3.111	1.803	28/3/7	0.444	0.387 (0.087)	-0.153*	0.602	0.531	0.325	0.656
DY-LIJ	88.9	3.111	1.932	28/3/0	0.332	$0.435\ (0.080)$	0.252**	I	I	I	I
DY-WEI1	88.9	2.222	1.760	20/0/0	0.389	0.390 (0.078)	$0.004^{ m ns}$	I	I	I	I
DY-WEI2	88.9	2.778	2.186	25/6/0	0.367	0.564~(0.085)	0.392**	I	I	I	I
DY-XIG	77.8	3.889	2.077	35/4/10	0.399	0.481 (0.107)	0.171^{ns}	0.319	0.344	0.607	0.766
Mean (95% CI)	75.2	2.402	1.731	21.6/2.8/2.1	0.334	0.369	0.061				
							(-0.065, 0.190)				
P. ludlowii											
LT-MAI1	22.2	1.222	1.097	11/0/0	0.056	0.051 (0.034)	$-0.098^{ m ns}$	0.341	1.000	0.217	1.000
LT-MAI2	0.0	1.000	1.000	0/0/6	0.000	0.000 (0.000)	I	I	I	I	I
LT-MAI3	0.0	1.000	1.000	0/0/6	0.000	0.000 (0.000)	I	Ι	Ι	I	I
LT-MAI4	0.0	1.000	1.000	0/0/6	0.000	0.000 (0.000)	I	I	I	I	I
Mean (95% CI)	5.6	1.056	1.024	9.5/0/0	0.014	0.013	-0.050				

57, -0.021)	0.066	56, 0.195)	
)0.0	-	(-0.0)	
	0.285 (0.023)		
	0.259		correction.
	18.8/2.2/1.6		after the Bonferroni
	1.564		ificant values
	2.085		ificant; in bold, sign
	58.8)01; ^{ns} not sign
	subsec.	% CI)	01; **P < 0.0
	P.	anae (95	; **P < 0.
	Mean	Delavay	$^{1}*P < 0.05$

² Numbers reported are *P* values of sign tests (left) and Wilcoxon signed-rank tests (right) under IAM (infinite allele model), and SMM (stepwise mutation model) conducted using the program Bottleneck. Significant P values (at the 0.05 level) are boldfaced.

		u	ISSR			pla	istid DNA	
l axon/source	df	SS	Vc	%	df	SS	Vc	%
P. delavayi + P. ludlowii								
Among taxa	1	175.805	2.017	40.3***	1	158.699	3.264	42.02***
Among populations within taxa	15	401.649	1.395	27.9***	15	618.048	4.433	57.08***
Within populations	293	465.285	1.588	31.8***	142	9.989	0.070	0.91^{***}
P. delavayi (no regional categories)								
Among populations	12	401.295	1.575	46.4***	12	618.048	4.966	98.40**
Within populations	253	460.935	1.822	53.6 ^{ns}	124	9.989	0.081	1.60^{**}
P. ludlowii (no regional categories)								
Among populations	3	0.355	0.001	$^{\rm su}$ 6.0	С	0.000	0.000	I
Within populations	40	4.350	0.109	99.1^{ns}	18	0.000	0.000	Ι
Paeonia subsect. Delavayanae								
Among 3 genetic clusters ¹	2	230.119	1.773	38.5***	I	Ι	I	I
Among populations within clusters	14	347.335	1.245	27.0***	Ι	Ι	I	I
Within populations	293	465.285	1.588	34.5***	I		I	I

Table 3. Analysis of molecular variance (AMOVA) of *Paeonia* subsect. *Delavayanae* based on nSSR and plastid DNA variation. *df* = degrees of freedom; SS = 0.01 · ** D > 0.001 · n_{suct} circuition V U ** . V V 4 * • د

variation.	*P < 0.0.	5; ** $P <$	0.01; **F	o < 0.001	; ^{ns} not si	ignificar	ıt; NA, nc	ot applical	ole; in bo	ld, signifi	cant valu	es after t	he Bonfer	rroni corr	ection.		
	DS-	DS-	-SQ	DT-	DT-	DT-	DY-	DY-	DY-	DY-	DY-	DY-	DY-	LT-	LT-	LT-	LT-
	LIT	MUL	XIA	BOM1	BOM2	IYN	DAL	DEQ	KUN	LIJ	WEII	WE12	XIG	MAII	MAI2	MAI3	MAI4
DS-LIT	I																
DS-MUL	0.602***	I															
DS-XIA	0.576***	0.520***	Ι														
DT-BOM1	0.622**	0.672**	0.587**	I													
DT-BOM2	0.599 ^{ns}	0.683 ^{ns}	0.580 ^{ns}	0.310^{ns}	I												
DT-NYI	0.616**	0.687**	0.620*	0.432^{ns}	$0.044^{\rm ns}$	Ι											
DY-DAL	0.530***	0.534***	0.382***	0.490**	0.439^{ns}	0.496**	Ι										
DY-DEQ	0.629***	0.680***	0.494***	0.671^{**}	$0.691^{\rm ns}$	0.728**	0.464***	Ι									
DY-KUN	0.540***	0.546***	0.495***	0.574^{**}	0.576^{ns}	0.604^{**}	0.507***	0.573***	I								
DY-LIJ	0.487***	0.381***	0.321***	0.527*	$0.481^{\rm ns}$	0.557*	0.409***	0.557***	0.422***	Ι							
DY-WEI1	0.562***	0.573***	0.377**	0.517*	$0.552^{\rm ns}$	0.588*	0.424***	0.528***	0.481^{***}	0.422**	Ι						
DY-WEI2	0.501**	0.582**	0.330^{**}	0.464^{ns}	0.439^{ns}	0.521*	0.270**	0.373**	0.441^{**}	0.380^{**}	0.170^{*}	I					
DY-XIG	0.437***	0.497***	0.361***	0.505**	0.477^{ns}	0.536**	0.421^{***}	0.472***	0.368***	0.213***	0.372***	0.332**	I				
LT-MAI1	0.777***	0.824***	0.718**	0.856**	0.911^{ns}	0.902*	0.664***	0.831^{***}	0.736***	0.717^{**}	0.738**	0.708**	0.688^{***}	I			
LT-MAI2	0.763**	0.825**	0.702**	0.855*	0.936^{ns}	0.914*	0.646^{**}	0.826^{**}	0.720**	0.699**	0.720^{**}	0.676*	0.668^{**}	0.049^{ns}	I		
LT-MAI3	0.705^{ns}	0.786 ^{ns}	0.595 ^{ns}	0.778^{ns}	$0.861^{\rm ns}$	0.836^{ns}	0.568 ^{ns}	$0.754^{\rm ns}$	0.661^{ns}	0.587^{ns}	0.620^{ns}	0.485^{ns}	0.583^{ns}	-0.062^{ns}	NA	Ι	
LT-MAI4	0.739*	0.808*	0.661*	0.825^{ns}	0.913^{ns}	0.888^{ns}	0.615*	0.802^{*}	0.695*	0.655*	0.680*	0.608^{ns}	0.635*	0.016^{ns}	NA	NA	

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Supplementary Material

Population genetic dynamics of Himalayan-Hengduan tree peonies, *Paeonia* subsect. *Delavayanae*

Jin-Mei Zhang, Jordi López-Pujol, Xun Gong, Hua-Feng Wang, Roser Vilatersana and Shi-Liang Zhou

Supplementary Text 1. Extended material of Ecological niche modeling methodology.

Ecological niche modeling (ENM) was performed to evaluate the potential distribution of Paeonia delavayi, P. ludlowii, and P. delavayi + P. ludlowii. We employed the maximum entropy algorithm, as implemented in MaxEnt v.3.3 (Phillips et al., 2006). The current distribution information for both species was obtained from the sampling sites (Table 1), literature (e.g., Hong, 2010), and specimens deposited in the main Chinese herbaria (through the Chinese Virtual Herbarium platform; www.cvh.ac.cn). After removing duplicate records within each pixel (2.5 arc-min, ca. 5 km), we obtained a total of 119 presence records of P. delavayi and 13 records of P. ludlowii. A set of 19 bioclimatic variables at 2.5 arc-min resolution covering the distribution range (and neighboring areas) for both species under current conditions (1950-2000) were downloaded from the WorldClim website (www.worldclim.org; Hijmans et al., 2005). Although finer resolutions are available (30 arcsec), these may not be appropriate given uncertainties associated with geo-referencing approximate localities (a common situation for old herbarium records in China, especially for remote areas) or with geo-reference errors. After a correlation analysis in a random sample of 1000 points within the study area, we selected a smaller set of eight (relatively) uncorrelated variables: mean diurnal range (bio2), isothermality (bio3), temperature seasonality (bio4), mean temperature of the coldest quarter (bio11), annual precipitation (bio12), precipitation of the driest month (bio14), precipitation seasonality (bio15), and precipitation of the coldest quarter (bio19). The selection of variables from pairs or groups of highly correlated $(r \ge |0.9|)$ ones was done on the basis of their relative contribution to the models (percent contribution, permutation importance, jackknife of regularized gaining train), making sure that the top most influential variables for the two species were selected.

The distribution model under current conditions was projected to the Last Glacial Maximum (LGM, *ca.* 21,000 yr BP) using palaeoclimatic layers simulated by the Community Climate System Model version 4 (CCSM4; Gent et al., 2011), the Model for Interdisciplinary Research on Climate Earth System Model (MIROC-ESM; Watanabe et al., 2011), and the New Earth System Model of Max Planck Institute for Meteorology (MPI-ESM; http://www.mpimet.mpg.de/en/science/models/mpi-esm/). For the models with a considerable number of occurrences (*P. delavayi* and *P. delavayi* + *P. ludlowii*) 20 replicates of MaxEnt (using the subsample method) were run, and model performance was assessed using the area

under the curve (AUC) of the receiver operating characteristic (ROC) plot with 25% of the localities randomly selected to test the model. AUC scores range between 0.5 (randomness) and 1 (exact match), and a value above 0.9 is considered a good performance of the model (Swets, 1988). Given the low number of occurrences for *P. ludlowii* (13), we used a methodology based on a jackknife (or 'leave-one-out') procedure to test the model (Pearson et al., 2007). With this procedure the model is built (or 'trained') using n-1 occurrences, and tested using the discarded locality. Following the recommendation of Pearson et al. (2007), the Lowest Presence Threshold (LPT, also commonly referred as 'minimum training presence' in the MaxEnt terminology) was used as the cut-off value to decide whether the discarded locality is 'suitable' or 'unsuitable'. Performance of models for *P. ludlowii* was evaluated through success rate (percentage of right predictions) and statistical significance (a *P*-value computed across the set of jackknife predictions, which was done using the software provided by Pearson et al. (2007). To get the definitive model (that is, using all occurrence points) MaxEnt was run 20 times using the bootstrap method.

The MaxEnt jackknife analysis was used to evaluate the relative importance of the eight bioclimatic variables employed based on their gain values when used in isolation. To convert the continuous value projection to a binary presence/absence distribution, we applied the maximum sensitivity plus specificity logistic threshold, which is very robust with all types of data (Liu et al., 2016). All ENM predictions were visualized in ArcGIS v. 10.2 (ESRI, Redlands, CA, USA). The suitable area (in km²) for all models at each time slice was also calculated in ArcGIS. To estimate suitable area gains or losses (or unchanged areas) for the LGM scenarios with respect to the present, binary output maps were overlapped with the Intersect Tool of ArcGIS.

Niche similarity between P. delavayi and P. ludlowii was measured through two niche overlap indices, Hellinger-derived I and Schoener's D. These metrics are implemented in the software ENMTools v.1.4.3 (Warren et al., 2010). The I and D values range from 0 (when the two species show completely discordant ENMs) to 1 (complete niche overlap). Two quantitative tests of niche similarity that are also implemented in ENMTools were further used. The 'niche identity test' determines whether two ENMs are identical by comparing the empirical I and D values to those generated from a number of pseudoreplicated datasets that are obtained by pooling all the occurrences of the two species and randomly splitting them into two new groups. The 'background test' determines whether ENMs are more similar (or less similar) than would be expected given the underlying environmental differences between the regions in which the entities to compare occur. A null distribution of I and D values was generated by comparing the actual occurrence records of P. delavavi with a set of randomly simulated occurrences within the range of P. ludlowii, and vice versa. Niche conservatism (niches more similar than expected) or divergence (niches more different than expected) can be interpreted when the empirical niche overlap values are significantly larger and smaller than those of the null hypothesis, respectively. Backgrounds were delimited by creating a buffer zone of 20 km around the occurrence points of each species, with the aid of the specific tools included in ArcGIS. For both tests, null distributions were generated from 100 pseudoreplicates. Finally, we estimated the niche breadth for each species by calculating the inverse concentration statistic of Levins (1968), as implemented in ENMTools; values range from 0 (only one pixel shows suitability greater than zero) to 1 (all pixels equally suitable).

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Fig. S1. Unweighted pair-group method using arithmetic averages (UPGMA) dendrogram using Nei's (1972) standard genetic distance (D_s), based on nSSR variation. Numbers above branches represent bootstrap support for 1000 replicates; only values equal to or greater than 50% are given. *Paeonia delavayi* populations are in bold, those of *P. ludlowii* in italics.



Reference

Nei, M., 1972. Genetic distance between populations. Am. Nat. 106, 289–291.

Fig. S2. Results of Structure analysis for all individuals of *Paeonia* subsect. *Delavayanae* studied, based on nSSR data. Assignation of individuals to genetic clusters from K = 2 to K = 13.





Fig. S3. Results of Structure analysis for all individuals of *Paeonia* subsect. *Delavayanae* studied, based on nSSR data. Representation of the 20 runs at K = 2.



Fig. S4. Observed and expected mismatch distribution under constant population size of *Paeonia* subsect. *Delavayanae*.

Fig. S5. Comparison of potential distributions for *Paeonia delavayi*, *P. ludlowii*, and *P. delavayi* + *P. ludlowii* between the present time and each of the three climatic scenarios assayed for the Last Glacial Maximum (LGM, ca. 21,000 years BP). In gray, no change between present and past climatic scenarios; in red, LGM contraction areas (that is, areas gained in the present compared to the LGM); in green, LGM expansion areas.



Fig. S6. Results of the two tests of niche similarity between *Paeonia delavayi* and *P. ludlowii* as obtained using ENMTools. (A) and (B), niche identity tests between both species based on Hellinger-derived *I* and Schoener's *D*, respectively; (C) and (D), background tests between *P. delavayi* occurrences and *P. ludlowii* background based on *I* and *D*, respectively; (E) and (F), background tests between *P. ludlowii* occurrences and *P. delavayi* background based on *I* and *D*, respectively. The arrows represent the observed niche overlap between ENMs (I = 0.569 and D = 0.307), whereas the histograms are those expected under the null hypotheses.



Population abbreviation*	Locality (village, county, province)	Voucher	psbA-trnH	rps16-trnQ	ndhF	ndhC-trnV(UAC)
P. delavayi						
DS-LIT	Maiwa, Litang, Sichuan	WY06078-LTP (PE)	MH025548	MH025569	MH025590	MH025611
DS-MUL	Shawan, Muli, Sichuan	91035 (KUN)	MH025549	MH025570	MH025591	MH025612
DS-XIA	Baiyi, Xiangcheng, Sichuan	WY06074-XCP (PE)	MH025553-554	MH025574-575	MH025595-596	MH025616-617
DT-BOM1	Guxiang, Bomi, Tibet	H060015 (PE)	MH025538	MH025559	MH025580	MH025601
DT-BOM2	Sumzom, Bomi, Tibet	H060016 (PE)	MH025539	MH025560	MH025581	MH025602
DT-NYI	Zanba, Nyingchi, Tibet	H060012 (PE)	MH025541	MH025562	MH025583	MH025604
DY-DAL	Cangshan, Dali, Yunnan	91027 (KUN)	MH025542	MH025563	MH025584	MH025605
DY-DEQ	Yunnan, Dêqên, Mingyong	701 (PE)	MH025540	MH025561	MH025582	MH025603
DY-KUN	Xishan, Kunming, Yunnan	WH05 (PE)	MH025544	MH025565	MH025586	MH025607
DY-LIJ	Maoniuping, Lijiang,Yunnan	82101 (KUN)	MH025545-547	MH025566-568	MH025587-589	MH025608-610
DY-WEI1	Duoduo, Weixi, Yunnan	82410 (KUN)	MH025550-551	MH025571-572	MH025592-593	MH025613-614
DY-WEI2	Laboluo, Weixi, Yunnan	82507 (KUN)	MH025552	MH025573	MH025594	MH025615
DY-XIG	Hala, Xianggelila, Yunnan	R05 (PE)	MH025543	MH025564	MH025585	MH025606
P. ludlowii						
LT-MAI1	Zhare, Mainling, Tibet	H03072 (PE)	MH025556	MH025577	MH025598	MH025619
LT-MAI2	Jinxuega, Mainling, Tibet	H03082 (PE)	MH025558	MH025579	MH025600	MH025621
LT-MAI3	Gangga, Mainling, Tibet	H06013 (PE)	MH025557	MH025578	MH025599	MH025620

Table S1. Localities and voucher information for the population studied, and GenBank accession numbers for the four plastid DNA regions studied of *Paeonia* subsect. *Delavayanae*.

LT-MAI4	Between Gangga and Mainling, Mainling, Tibet	H06014 (PE)	MH025555	MH025576	MH025597	MH025618
<u>Outgroups</u>						
P. jishanensis		BOP001735	MH051894	MH051896	MH051899	MH051902
P. ostii		BOP001481	KJ946192	MH051897	MH051900	KJ945956
P. qiui		BOP001030	MH051895	MH051898	MH051901	MH051903

* The population abbreviation consists of the first letter of species epithet, the first letter of province, and the first tree letters of county.

Note: KUN = Herbarium of Kunming Institute of Botany, Chinese Academy of Sciences; PE = Herbarium of the Institute of Botany, Chinese Academy of Sciences.

Table S2. Plastid genes, primers designed and annealing temperatures (Ta) for each plastid DNA region used in this study. PCRs were performed using the following program: 94 °C for 4 min; 35 cycles of 94 °C for 30 s, Ta °C for 30 s, and 72 °C for 2 min; and with a final extension at 72 °C for 10 min.

Gene	Primer name	Sequence (5'-3')	Ta (°C)
psbA-trnH	26f	CGCGCATGGTGGATTCACAATCC	52
	475r	GTTATGCATGAACGTAATGCTC	
rps16-trnQ	6022f	CGTTGCTTTCTACCACATCG	58
	7383r	CTATTCGGAGGTTCGAATCC	
ndhC- trnV ^(UAC)	51189f	CGGATTCGAAATTGTAACCAAGC	56
	52001r	TGAAAAACAAGGGTCCTTGGC	
ndhF	110492f	CAATTATTCGCCTATCAA	52
	111588r	GTCTCAATTGGGTTATATGATG	

Table S3. Repeat motifs, annealing temperatures (Ta), primer sequences and the range of alleles detected per locus for the nine nSSR studied loci of *Paeonia* subsect. *Delavayanae*. The PCRs were performed using the following program 3 min at 94 °C, followed by 25 cycles of 30 s at 94 °C, 30 s at Ta °C, and 45 s at 72 °C, with a final extension of 10 min at 72 °C.

Locus	Repeat motif	Primer sequences (5'–3')	Size (bp)	Ta (°C)
Jx02*	(TC)9	F: TTGGTTGGTGAAGGTGTT	289-331	54
		R: CTTCGATAACCGCAGGAGGAT		
Jx05*	(CT)17	F: GCCACAAGAAAACAAAAACC	214-246	54
		R: CCTTCACCACTACTTCCCCAT		
Jx17*	(TC)20	F: CAAACTACCTGAATGTTCGGCTC	187–225	54
		R: CATCAAATTACCAAAGAAATCCT		
Jx27*	(TC)5	F: GTTATAGAACCACTGACAT	303-321	48
		R: TGAGAGACAAATAATCGTG		
Pdel05**	(AG)15	F: CCAATGTGGAAAATGAGTT	180-226	50
		R: CAAGCACAAGATGTAAGAA		
Pdel11**	(TGG)6	F: CTGCCATTTCTTGCCTTCTTTGT	230-242	54
		R: TCTACCCTGCCAACAGCACATAC		
Pdel20	(TC)6	F: TATAAATGGGAAGCAGACTCAA	253-329	54
		R: TATACTCAGCCTCGAAAAGAAG		
Pdel22	(AG)9	F: TCGCCCAACCTGTCGTGGAGAT	300-328	54
		R: TTGAATAGAGCGGAATGGAAAA		
Pdel35	(GA)10	F: ATGTCACCGAAAGTTGTGC	293-313	54
		R: AAAGCCTGGTGCAGTTATT		

*Wang et al. 2009; **Zhang et al. 2011

References

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Zhang, J., Liu, J., Sun, H., Yu, J., Wang, J., Zhou, S., 2011. Nuclear and chloroplast SSR markers in *Paeonia delavayi* (Paeoniaceae) and cross-species amplification in *P. ludlowii*. Am. J. Bot. 98, e346–e348.

Table S4. Summary of genetic diversity statistics, neutrality tests and mismatch distribution analyses in *Paeonia* subsect. *Delavayanae*. Number of sequences (*N*), polymorphic sites (*S*), number of unique haplotypes (*h*), haplotype diversity (H_d), nucleotide diversity (π), average number of nucleotide difference (*k*). SD = standard desviation. Significance of neutrality parameters were tested by coalescent analyses with their significance (ns, no significant). (R_2), Ramos-Onsins and Rozas's index, (HRI), Harpending's raggedness index.

Group	N	S	h	$H_{\rm d}({ m SD})$	π (SD) × 10 ⁻³	k	Tajima's D	Fu & Li's <i>D</i> *	Fu & Li's <i>F</i> *	Fu's Fs	R_2	HRI
P. delavayi + P. ludlowii	159	22	14	$0.90 (\pm 0.5 \times 10^{-4})$	$2.08 (\pm 0.06)$	5.57	1.19 ns	1,81 ns	1.88 ns	2.95 ns	0.13 ns	0,07 ns
P. delavayi	137	18	13	$0.90 \ (\pm 0.7 \times 10^{-4})$	$1.86 (\pm 0.06)$	4.98	1.43 ns	1.69 ns	1.90 ns	2.42 ns	0.14 ns	0.05 ns
P. delavayi (Sichuan + Yunnan)	126	18	13	$0.91 (\pm 0.007)$	$1.83 (\pm 0.07)$	4.91	1.32 ns	1.69 ns	1.85 ns	2.13 ns	0.14 ns	0.05 ns
DY-WEI1	8	1	2	$0.43 (\pm 0.028)$	$0.16 (\pm 0.06)$	0.43	0.33 ns	0.89 ns	0.83 ns	0.54 ns	0.21 ns	0.20 ns
DY-LIJ	9	2	3	0.67 (± 0.110)	$0.29 (\pm 0.07)$	0.78	0.20 ns	-0.22 ns	-0.14 ns	-0.11 ns	0.21 ns	0.26 ns

Dataset	psbA-trnH	rps16-trnQ	ndhF	ndhC- trnV(UAC)	Concatenated sequences
N° sequences	18	18	18	18	18
Total characters	372	1131	616	626	2745
Variable characters (%)	5 (1.34)	19 (1.68)	16 (2.60)	16 (2.56)	56 (2.04)
Parsimony informative characters	4	8	12	16	30
No. of steps	16	13	21	6	42
No. of trees	6	6	33	2	38
Consistence index (CI)	0.33	0.62	0.75	1.00	0.78
Retention index (RI)	0.00	0.67	0.88	1.00	0.84
Homoplasy index (HI)	0.67	0.38	0.25	0.00	0.22
Model of evolution	F81+G	GTR+G	GTR+G	F81	

Table S5. Numerical results from phylogenetic analyses of each independent plastid DNA region and concatenated sequences in *Paeonia* subsect. *Delavayanae*.

Table S6. Mean recent migration rates (m) among the studied populations of *Paeonia* subsect. *Delavayanae* estimated from nine nSSR data using the BayesAss program. Values on the diagonal (underlined) are the proportions of individuals in each generation that are not migrants. Simulations in BayesAss show that in instances where there is no information in the data, the mean m and 95% confidence interval for datasets of 17 populations are 0.0105 and 0.0000–0.0933, respectively; values in bold are the m rates that are informative.

Populations	DS- LIT	DS- MUL	DS- XIA	DT- BOM1	DT- BOM2	DT- NYI	DY- DAL	DY- DEQ	DY- KUN	DY- LIJ	DY- WEI1	DY- WEI2	DY- XIG	L- MAI1	L- MAI2	L- MAI3	L- MAI4
DS-LIT	0.6867	0.0062	0.0055	0.0064	0.0058	0.0066	0.0060	0.0058	0.0054	0.0066	0.2233	0.0061	0.0057	0.0059	0.0062	0.0060	0.0060
DS-MUL	0.0013	<u>0.9806</u>	0.0012	0.0014	0.0012	0.0011	0.0011	0.0012	0.0013	0.0015	0.0015	0.0011	0.0009	0.0012	0.0014	0.0009	0.0012
DS-XIA	0.0096	0.1659	<u>0.6935</u>	0.0102	0.0091	0.0090	0.0101	0.0096	0.0091	0.0093	0.0093	0.0089	0.0095	0.0089	0.0088	0.0092	0.0100
DT-BOM1	0.0035	0.0030	0.0027	<u>0.9548</u>	0.0026	0.0028	0.0025	0.0027	0.0026	0.0028	0.0030	0.0023	0.0033	0.0036	0.0025	0.0026	0.0028
DT-BOM2	0.0141	0.0145	0.0146	0.0391	<u>0.7489</u>	0.0141	0.0135	0.0136	0.0141	0.0133	0.0134	0.0148	0.0135	0.0148	0.0141	0.0148	0.0148
DT-NYI	0.0141	0.0143	0.0152	0.0665	0.0132	0.7219	0.0136	0.0134	0.0137	0.0145	0.0136	0.0142	0.0151	0.0148	0.0138	0.0134	0.0148
DY-DAL	0.0011	0.0013	0.0011	0.0014	0.0015	0.0011	<u>0.9798</u>	0.0012	0.0012	0.0013	0.0014	0.0012	0.0012	0.0011	0.0015	0.0013	0.0013
DY-DEQ	0.0011	0.0013	0.0013	0.0012	0.0014	0.0013	0.0012	0.9804	0.0012	0.0012	0.0012	0.0011	0.0013	0.0012	0.0012	0.0012	0.0014
DY-KUN	0.0013	0.0012	0.0013	0.0013	0.0013	0.0014	0.0014	0.0010	<u>0.9799</u>	0.0011	0.0011	0.0012	0.0012	0.0014	0.0013	0.0011	0.0016
DY-LIJ	0.0099	0.0479	0.0100	0.0096	0.0091	0.0094	0.0094	0.0102	0.0105	<u>0.6983</u>	0.0095	0.0096	0.1190	0.0095	0.0088	0.0097	0.0096
DY-WEI1	0.0019	0.0021	0.0019	0.0020	0.0020	0.0021	0.0022	0.0020	0.0019	0.0017	<u>0.9678</u>	0.0020	0.0020	0.0023	0.0021	0.0019	0.0021
DY-WEI2	0.0033	0.0026	0.0027	0.0029	0.0031	0.0030	0.0037	0.0038	0.0035	0.0040	0.0035	<u>0.9489</u>	0.0033	0.0030	0.0034	0.0029	0.0023
DY-XIG	0.0011	0.0010	0.0013	0.0012	0.0010	0.0014	0.0010	0.0015	0.0014	0.0011	0.0012	0.0011	<u>0.9804</u>	0.0014	0.0014	0.0011	0.0012
L-MAI1	0.0018	0.0015	0.0016	0.0019	0.0020	0.0017	0.0016	0.0017	0.0019	0.0015	0.0015	0.0018	0.0017	<u>0.9727</u>	0.0020	0.0016	0.0016
L-MAI2	0.0136	0.0140	0.0139	0.0129	0.0128	0.0132	0.0139	0.0129	0.0115	0.0122	0.0127	0.0124	0.0126	0.0990	<u>0.7076</u>	0.0120	0.0129
L-MAI3	0.0137	0.0150	0.0146	0.0140	0.0134	0.0144	0.0144	0.0152	0.0148	0.0142	0.0133	0.0145	0.0141	0.0382	0.0133	<u>0.7484</u>	0.0144
L-MAI4	0.0135	0.0131	0.0149	0.0148	0.0146	0.0131	0.0147	0.0130	0.0132	0.0145	0.0152	0.0148	0.0138	0.0646	0.0150	0.0148	0.7224

Model	Predicted area (km ²)	Difference respect to present (km ² and %)	Overlap with present (km ² and %)	Mean elevation (m)				
P. delavayi								
Present	273,749	_		3160				
LGM-CCSM	152,514	-121,235 (-44.29)	129,386 (47.26)	2824				
LGM-MIROC	72,759	-200,990 (-73.42)	63,164 (23.07)	3052				
LGM-MPI	193,679	-80,070 (-29.25)	151,635 (55.39)	2754				
Average LGM	139,651	-134,098 (-48.99)	114,728 (41.91)	2877				
P. ludlowii								
Present	32,791			3116				
LGM-CCSM	117,116	+84,325 (+257.16)	31,515 (96.11)	3143				
LGM-MIROC	31,529	-1262 (-3.85)	17,943 (54.72)	2828				
LGM-MPI	20,677	-12,114 (-36.94)	273 (0.83)	1281				
Average LGM	56,441	+23,649 (+72.12)	16,577 (50.55)	2417				
P. delavayi + P. ludlowii								
Present	248,243	_		2950				
LGM-CCSM	116,221	-132,022 (-53.18)	95,293 (38.39)	2875				
LGM-MIROC	59,071	-189,172 (-76.20)	49,473 (19.93)	2996				
LGM-MPI	159,494	-88,749 (-35.75)	113,026 (45.53)	2727				
Average LGM	111,595	-136,647 (-55.05)	85,931 (34.62)	2866				

Table S7. Predicted potential distribution of *Paeonia* subsect. *Delavayanae* for all general circulation models used in this study.