

# A multimembership catalogue for 1876 open clusters using UCAC4 data

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## ABSTRACT

The main objective of this work is to determine the cluster members of 1876 open clusters, using positions and proper motions of the astrometric fourth United States Naval Observatory (USNO) CCD Astrograph Catalog (UCAC4). For this purpose, we apply three different methods, all based on a Bayesian approach, but with different formulations: a purely parametric method, another completely non-parametric algorithm and a third, recently developed by Sampedro & Alfaro, using both formulations at different steps of the whole process. The first and second statistical moments of the members' phase-space subspace, obtained after applying the three methods, are compared for every cluster. Although, on average, the three methods yield similar results, there are also specific differences between them, as well as for some particular clusters. The comparison with other published catalogues shows good agreement. We have also estimated, for the first time, the mean proper motion for a sample of 18 clusters. The results are organized in a single catalogue formed by two main files, one with the most relevant information for each cluster, partially including that in UCAC4, and the other showing the individual membership probabilities for each star in the cluster area. The final catalogue, with an interface design that enables an easy interaction with the user, is available in electronic format at the Stellar Systems Group (SSG-IAA) web site (<http://ssg.iaa.es/en/content/sampedro-cluster-catalog>).

**Key words:** methods: statistical – catalogues – open clusters and associations: general.

## 1 INTRODUCTION

Open clusters are excellent laboratories for exploring a large number of astrophysical problems: they are crucial to set constraints to the stellar evolutionary models (i.e. Landin et al. 2006; Charbonnel 2017, and references therein) and to study in detail the star formation process and the possible physical mechanisms driving and controlling all the different steps that lead from a molecular cloud to a gravitationally bound set of stars (i.e. Larson 1994; Elmegreen et al. 2000; Lada 2010). Since the pioneering work by Becker (1964), open clusters have also been used as suitable probes for drawing the Galactic structure in terms of shape, morphology, kinematics and chemical distribution (i.e. van den Bergh & McClure 1980; Janes, Tilley & Lynga 1988; Alfaro, Cabrera-Cano & Delgado 1991; Dias & Lépine 2005; Frinchaboy & Majewski 2008; Magrini et al. 2009; Lépine et al. 2011; Gilmore et al. 2012; Camargo, Bica & Bonatto 2013; Oliveira et al. 2013; Heiter et al. 2014; Junqueira et al. 2015; Costado et al. 2017; Magrini et al. 2017). Descending to conceptually smaller scales, open clusters are the observational basis of our knowledge of the initial mass function (Kroupa 2001;

Bastian, Covey & Meyer 2010), the existence or not of a primordial mass segregation, (Bonnell & Davies 1998; Allison et al. 2009; Parker, Goodwin & Allison 2011; Parker & Dale 2017), the time-scales for the destruction or dilution of these systems (Parker et al. 2009; Krumholz et al. 2014; Vicente, Sánchez & Alfaro 2016), the concurrence of different star formation bursts in a single cluster (Wünsch et al. 2017), etc. In short, open clusters represent unique observational targets for the development of Galactic and stellar physics (see Gilmore et al. 2012).

However, in most cases, making use of these stellar systems requires the preceding step of determining the physical members of the cluster. The classification between members and non-members of the stars located in the stellar field of the cluster depends on how we have defined what a stellar cluster is and which of its properties we use to carry out the classification, which is at the same time influenced by the observed variables available. Here we are using data from the fourth United States Naval Observatory (USNO) CCD Astrograph Catalog (UCAC4; Zacharias et al. 2013) – that is, positions and proper motions. Hence, we must use those phenomenological aspects of the actual definition of a cluster that leave their mark on the phase-space subspace formed by these variables. It is evident that the first clue for the existence of a stellar cluster is revealed by an increase in superficial stellar density in the plane of the sky. Since

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the 1930s, we have also known that the kinematic behaviour of a cluster is reflected in a greater density, with regard to the field stars, in the space of the kinematic variables. Moreover, if, as in this case, the aim is to carry out an analysis of the whole system of Galactic clusters, then we also need all-sky, accurate and homogeneous data.

Several efforts have been made over recent years to generate homogeneous and systematic astrometric catalogues, and to use them to analyse the membership to the system of open clusters of the catalogued stars. Examples of these works are Baumgardt, Detbarn & Wielen (2000), Loktin & Beshenov (2003), Beshenov & Loktin (2004), Dias et al. (2006), Dias et al. (2014), Kharchenko et al. (2005), Kharchenko et al. (2012) and Kharchenko et al. (2013, hereafter K13). Most of the techniques used to separate field and cluster populations address the problem from a statistical point of view, by computing the membership probabilities through the estimation of the probability density functions (PDFs) for cluster and stellar-field populations in the subspace formed by the astrometric variables. However, the model underlying the distribution of these variables is not unique, and nor is the algorithm used to determine the PDFs and to estimate the cluster membership probabilities.

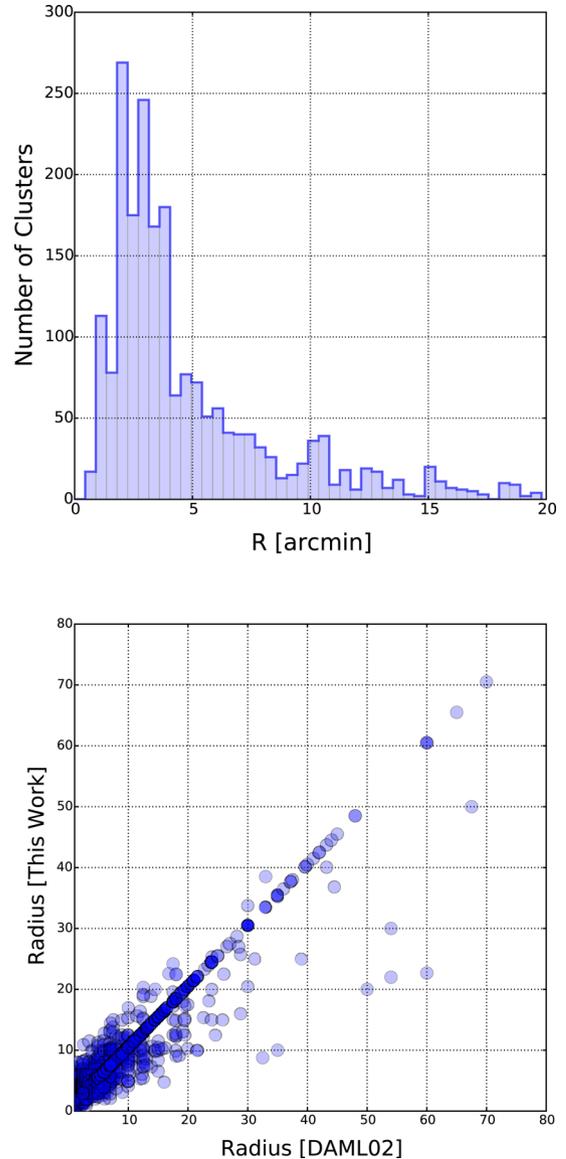
Since the pioneering work by Vasilevskis, Klemola & Preston (1958) using only proper motions, many other studies have proposed different approaches to the problem that respond to the observational fact of the distribution pattern diversity of the astrometric variables in the clusters studied (i.e. Cabrera-Caño & Alfaro 1985, 1990; Zhao & He 1990; Galadi-Enriquez, Jordi & Trullols 1998; Dias et al. 2014; Sampedro & Alfaro 2016; Gao 2016, among others). Cluster membership determination is therefore dependent on both the characteristics of the data and the distinct statistical approaches assumed by the different methods. In this work, we apply three different methods fully described in Cabrera-Caño & Alfaro (1985, 1990) and Sampedro & Alfaro (2016), to determine, in a homogeneous way, the members of the open clusters listed in Dias et al. (2002).

The paper is organized as follows. In Section 2, we summarize the data used in this work. In Section 3, we give a brief description of the three methods applied to determine the membership of the stars. In Section 4, we summarize the main results and we discuss the comparison with previous published catalogues. Finally, in Section 5, we highlight the main conclusions.

## 2 DATA

We use version 3.5 of the New Optically Visible Open Clusters and Candidates catalogue (hereafter DAML02; Dias et al. 2002), to select a sample of 2167 open clusters to be analysed. The stellar positions and the proper motions are taken from the UCAC4 (Zacharias et al. 2013). The catalogue contains data for over 113 million stars (105 million of them with proper-motion data), and is complete down to magnitude  $R = 16$ . The positional accuracy of the listed objects is about 15–100 mas per coordinate, depending on the magnitude. Formal errors in proper motions range from about 1 to 10  $\text{mas yr}^{-1}$ , depending on the magnitude and the observational history. Systematic errors in the proper motions are estimated to be about 1–4  $\text{mas yr}^{-1}$ .

As proposed by Sánchez, Vicente & Alfaro (2010), we should utilize the cluster angular radius as the best sampling radius to analyse the cluster membership when using the parametric model of the proper-motion distribution. However, given that we are dealing with three different methods, we proceed in a different way.



**Figure 1.** Top: the distribution of apparent angular cluster radii estimated in this work up to 20 arcmin. Bottom: a comparison between the radii estimated in this work and those from DAML02 (in arcmin).

We download the data from the UCAC4, using the VizieR service<sup>1</sup> (Ochsenbein, Bauer & Marcout 2000) for an initial cluster centre and within an area given by the DAML02 catalogued radius, plus an extra 15 arcmin to ensure we take all the stars in the cluster region. Then we again calculate apparent cluster angular radii through visual inspection of the radial density profiles (RDPs), where the apparent cluster angular radius is defined as the distance from the cluster centre where the RDP drops into the field. We should note that for this estimation we only make use of the stars catalogued in the UCAC4.

Fig. 1 shows the distribution of the apparent angular radii obtained in this work for the 2167 clusters listed in DAML02 up to an estimated radius of 20 arcmin. It also compares our final cluster radii and those listed in DAML02. We consider that the observed

<sup>1</sup> <http://vizier.u-strasbg.fr/viz-bin/VizieR>

**Table 1.** Redetermined central coordinates (J2000.0) for 10 open clusters are shown in the second and third columns. The coordinates from DAML02 are also included in the last two columns for comparison.

Name	$\alpha_{\text{this work}}$ (h:m:s)	$\delta_{\text{this work}}$ (g:m:s)	$\alpha_{\text{DAML02}}$ (h:m:s)	$\delta_{\text{DAML02}}$ (g:m:s)
Berkeley 28	06:52:07	02:54:47	06:52:12	02:56:00
Berkeley 39	07:46:48	−04:40:06	07:46:42	−04:36:00
Berkeley 43	19:15:32	11:16:20	19:15:36	11:13:00
Berkeley 45	19:19:05	15:42:47	19:19:12	15:43:00
Berkeley 50	20:10:01	34:57:58	20:10:24	34:58:00
BH 208	16:59:32	−37:07:20	16:59:36	−37:05:00
IC 1311	20:10:46	41:10:27	20:10:18	41:13:00
IC 1369	21:12:06	47:46:04	21:12:06	47:44:00
IC 361	04:18:54	58:15:00	04:19:00	58:18:00
Ruprecht 164	11:30:2:5	−60:45:10	11:30:51	−60:44:00

differences are mainly because of the fact that DAML02 is a compilation of different works based on different data sets. In addition, we compare our final estimated radii with those from K13, where, among other parameters, cluster radii were derived for 3006 stellar clusters using the PPMXL (Roesser, Demleitner & Schilbach 2010) and the Two-Micron All-Sky Survey (2MASS; Skrutskie et al. 2006) catalogues. In K13, the mode of the cluster radii distribution is about 7 arcmin while the same central value for our sample is about 2.5 arcmin. The difference is big but we should stress again that we are deriving a functional radius from and for a single catalogue. It would be wrong to consider this estimate as the true radius of the cluster, if that concept existed and was unique and absolute.

However, going further into the understanding of this difference we can envisage several causes. The observed disagreement might be mainly a result of the different techniques and different catalogues used to estimate the cluster radii. We have estimated apparent angular radii from all stars in the field, while K13 obtain the radii values from member stars, using a different procedure (Kharchenko et al. 2005, 2012). Therefore, the direct comparison must be taken with caution as it might be biased, showing strong disagreement. In addition, although Kharchenko et al. (2005) and K13 obtain similar cluster radii, several studies (Sharma et al. 2006; Bukowiecki et al. 2011) have claimed that using different data sources can lead to different cluster radius estimations.

We have also used the Digitized Sky Survey (DSS)<sup>2</sup> to redetermine the central positions, of 10 open clusters, according to our best analysis of the RDPs. The new central coordinates were estimated as the positions maximizing the central star density. Table 1 includes the new coordinates as well as those listed in DAML02 for the 10 clusters that showed different central positions.

For each cluster, once the UCAC4 data were downloaded, we discarded stars without proper-motion data or with errors in the proper motions larger than 12 mas yr<sup>−1</sup>. The latter stars should be considered as observational fails and they might not correspond to real stars, according to Zacharias et al. (2013). In addition, we exclude clusters with fewer than 20 stars and more than 10<sup>4</sup> stars. Cluster regions with few stars might suffer from subsampling effects as a result of patchy and heavy absorptions or incompleteness because of the photometric depth of the catalogue. However, open clusters with the largest angular diameters show a low surface brightness

with high contamination of foreground/background stars. This effect can prevent a proper determination of the cluster members using these methods. This pruning leads to a sample of 1876 open clusters analysed in this study.

### 3 METHODS USED IN THE MEMBERSHIP DETERMINATION

In this work, we apply three different methods to determine the membership for the previously selected 1876 open clusters. The first method (hereafter M1) is fully described in Sampedro & Alfaro (2016). Basically, it uses different sets of variables satisfying the simple condition of being more densely distributed for the cluster members than for the field stars. The membership probabilities are estimated in a one-dimensional space, defined by the Euclidean distances between every star and the cluster central overdensity, in the variable space of  $N$ -dimensions. Therefore, it reduces the estimation of membership probabilities from an  $N$ -dimensional space to a one-dimensional space. The method involves two iterative processes. Initially, the distances between every star and the cluster centre are estimated. Then, the distance distribution of the stars is modelled by a mixture of two one-dimensional Gaussians, one for the cluster members and other for the field stars. The parameters defining the total model and the membership probabilities are calculated through an iterative Wolfe estimation procedure (Wolfe 1970). The stars with a membership probability higher than 0.5 are selected as members according to the Bayes minimum error rate decision rule (Kulkarni & Harman 2011). In this work, this method is applied to the proper-motion data.

The second method (hereafter M2) is a Bayesian, non-parametric method developed by Cabrera-Caño & Alfaro (1990). It determines the members of the clusters using the positions (angular distances) and the proper-motion data without any a priori assumptions about the cluster and field star distributions, but assuming two hypotheses: (i) there are two populations in the field – cluster members and field stars – and (ii) the cluster members are more densely distributed in the phase space. Membership probabilities are calculated by using Gaussian kernel estimators in an iterative way through a discriminant analysis. In every iteration, three different probabilities for each star are estimated: one just using the positions of the stars, another using only the proper-motion data (kinematic probability), and the last using both positions and proper motions (joint probability). As a first option, cluster members in every iteration are selected as those stars with joint and kinematic probabilities higher or equal to 0.5.

The third method (hereafter M3) follows a parametric approach, fully described in Cabrera-Caño & Alfaro (1985). It uses only the stellar proper motions to determine the membership probabilities. This method fits the PDF of the whole sample by a mixture of two bivariate Gaussian distributions, one for the cluster members and other for the field stars. Through an iterative Wolfe estimation procedure, the parameters that define the total model are calculated, as well as the membership probabilities. The cluster members are those stars with a membership probability higher than 0.5.

Before starting with the membership analysis, we detect and remove sample outliers. In this task, we use a non-parametric technique (Cabrera-Caño & Alfaro 1985), which estimates the probabilities of the stars being outliers of the parent sample using the proper-motion data. In this way, any object with a probability of being an outlier greater than 0.5 was rejected for further analysis.

<sup>2</sup> <http://archive.stsci.edu/dss/>

**Table 2.** Comparison of the cluster mean proper motions obtained by the methods (see text for details).  $\Delta\mu_\alpha \cos \delta$ ,  $\Delta\mu_\delta$ ,  $\sigma_{\Delta\mu_\alpha \cos \delta}$  and  $\sigma_{\Delta\mu_\delta}$  are expressed in  $\text{mas yr}^{-1}$ .  $N$  is the number of clusters for which the compared methods converge to a solution.

Methods	M1–M2	M1–M3	M2–M3
$\Delta\mu_\alpha \cos \delta$	0.01	0.21	0.20
$\Delta\mu_\delta$	0.02	0.20	0.16
$1\sigma_{\Delta\mu_\alpha \cos \delta}$	0.81	1.11	1.39
$1\sigma_{\Delta\mu_\delta}$	0.90	1.16	1.28
$N$	1606	1713	1661

## 4 RESULTS

In this section, we compare the level of agreement achieved by the different methods described in Section 3. Furthermore, we compare our results with previous catalogues that also contain the main physical variables for each cluster, such as DAML02, K13 and D14 (Dias et al. 2014). The large number of objects listed in these catalogues will enable a comparison with a reliable statistical significance. We use the estimated probabilities, derived from the three different methods (the joint probability in the case of M2) as weight factors to calculate the cluster mean proper motions  $\mu_\alpha \cos \delta$  and  $\mu_\delta$ , as

$$\mu_\alpha \cos \delta = \frac{\sum_{i=1}^n P_i \mu_{\alpha \cos \delta, i}}{\sum_{i=1}^n P_i}$$

and

$$\mu_\delta = \frac{\sum_{i=1}^n P_i \mu_{\delta, i}}{\sum_{i=1}^n P_i},$$

where  $\mu_\alpha \cos \delta_i$  and  $\mu_{\delta, i}$  are the proper motion and  $P_i$  is the probability estimated by each method for the  $i$ th star. Similarly, the cluster proper-motion dispersion and the correlation coefficients have also been estimated. The data base containing the results is available in electronic format at the SSG-IAA<sup>3</sup> web site as two complementary files. The main parameters obtained by the three methods are included in a general catalogue. We also generate individual cluster-by-cluster files, giving the membership probabilities and additional information from the UCAC4 catalogue (more information is given in Appendix A).

### 4.1 Comparison of results from M1, M2 and M3

In this study, we investigate 1876 open clusters using the three different methods described in Section 3. Whereas M1 converges to a solution for a total of 1748 clusters (93 per cent), M2 and M3 do so for a total of 1693 (90 per cent) and 1819 (97 per cent), respectively. The three different approaches converge for 1584 (84 per cent) clusters in common.

Table 2 and Fig. 2 show the results of the comparison in  $\mu_\alpha \cos \delta$  and  $\mu_\delta$  between the methods. In Table 2,  $\Delta\mu_\alpha \cos \delta$  and  $\Delta\mu_\delta$  are the differences in the values of the cluster mean proper motions,  $\sigma_{\Delta\mu_\alpha \cos \delta}$  and  $\sigma_{\Delta\mu_\delta}$  represent their dispersions and  $N$  is the number of common clusters. The values of the mean differences are close to 0  $\text{mas yr}^{-1}$  with small dispersion (lower than 1.4  $\text{mas yr}^{-1}$ ), indicating good agreement between the methods. Furthermore, the mean differences among the methods, for the second moments of the cluster member distribution, are lower than the typical proper-motion errors in the UCAC4 catalogue, overall when comparing M1

and M3. There is also agreement in the number of cluster members estimated. However, for some cases, it is very difficult to distinguish between members and background stars, especially in regions associated with a patched absorption pattern or with clusters that have large apparent angular radii (i.e. with a low contrast between the cluster and the field). For these cases, we notice that M2 tends to determine lower fractions of members than M1 and M3. It is worth noting that M2 provides a different cluster view than that derived from the M1 and M3 methods; M2 makes use of four astrometric variables, while M1 and M3 do so for the two proper-motion components. Nevertheless, in this scenario, the results obtained from all methods should be considered with caution.

Finally, for six clusters (Alessi 53, Berkeley 76, Czernik 11, Loden 27, Pismis 5 and Ruprecht 3) none of the methods was capable of converging to a solution. This is not surprising as these clusters have few stars, with sparse distributions in the position and in the proper-motion spaces, showing the typical problems derived from small-number statistics. Moreover, Loden 27 is flagged as a dubious open cluster in DAML02 and Ruprecht 3 is not even considered as a physical system by Piatti, Dias & Sampedro (2017).

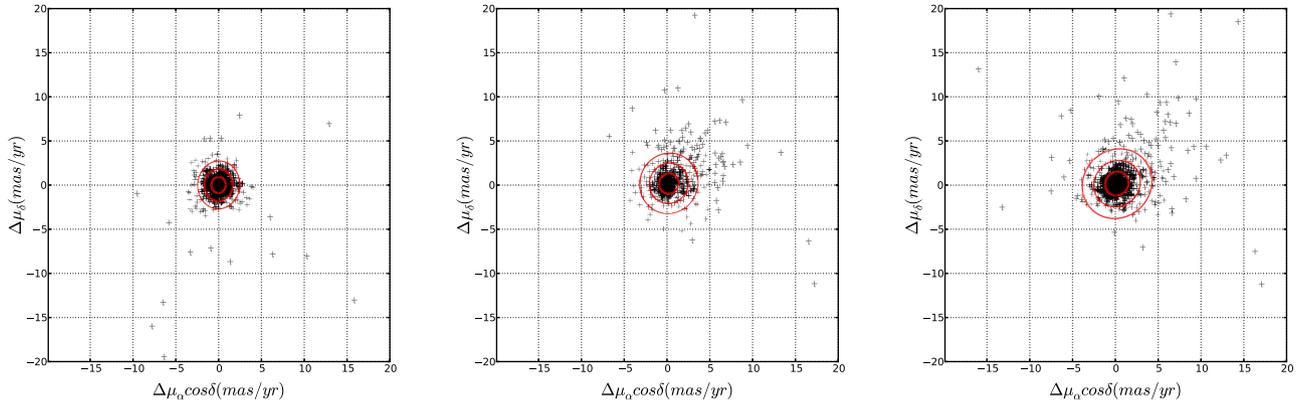
### 4.2 Comparison with other catalogues

A direct comparison of the results presented in this work with those in the literature is not straightforward, as different data, methods and criteria are used to select members, which can lead to different samples of stars used to determine the mean proper motions of the clusters. Therefore, we opt to compare our results with those published catalogues that present a large number of open clusters, such as DAML02, K13 and D14.

The DAML02 catalogue is a compendium of results published from different authors. Thus, it is based on different data and methods but it does present the mean proper motions in the Hipparcos system and it has been widely used in recent years. In K13, the cluster membership and mean proper motions are estimated for more than 3000 open clusters, using the PPMXL and 2MASS catalogues as the main observational basis. The authors use an interactive human control of a standardized set of multidimensional diagrams to determine kinematic and photometric membership probabilities for stars in the cluster region. D14 is a homogeneous study of 1805 open clusters using the UCAC4 catalogue. The authors determine mean proper motions and membership probabilities using a parametric model, fitting the PDF for the whole sample by two bivariate Gaussian functions, also taking into account the proper-motion errors of the stars.

Table 3 and Fig. 3 show the results of the comparison in  $\mu_\alpha \cos \delta$  and  $\mu_\delta$  for each cluster, reproducing very well, on average, the previous determinations. Dispersion differences lie below 2.8  $\text{mas yr}^{-1}$  in the comparison with D14 and DAML02 for the three methods. This indicates that there is no statistically significant separation between the compared distributions, taking into account the catalogued proper-motion errors. Basically, the higher dispersion in the comparison with DAML02 is likely to be because of the heterogeneity of the proper-motion sources used in DAML02. A few open clusters in Fig. 3 present differences greater than the  $3\sigma$  level. Generally, after a visual inspection, we notice that most of these open clusters do not show clear overdensity in the proper-motion space, using UCAC4 data. Likewise, some clusters such as Stock 1 and Alessi 13 present a high degree of field contamination and, in both cases, a proper determination of the cluster members is difficult. The greatest individual differences are found when comparing with K13, which is likely caused, as previously discussed, by the

<sup>3</sup> <http://ssg.iaa.es/en/content/sampedro-cluster-catalog>



**Figure 2.** Left: differences in the values of the cluster mean proper motions between M1 and M2. Centre: differences in the values of the cluster mean proper motions between M1 and M3. Right: differences in the values of the cluster mean proper motions between M2 and M3. The red ellipses represent the  $1\sigma$ ,  $2\sigma$  and  $3\sigma$  dispersion levels, respectively.

**Table 3.** Comparison of the mean proper motions obtained by the three methods used in this work with those published by DAML02, K13 and D14.  $\Delta\mu_\alpha \cos \delta$ ,  $\Delta\mu_\delta$ ,  $\sigma_{\Delta\mu_\alpha \cos \delta}$  and  $\sigma_{\Delta\mu_\delta}$  are expressed in  $\text{mas yr}^{-1}$ .  $N$  is the number of clusters in common between each method and the catalogue indicated.

	D14–M1	D14–M2	D14–M3	K13–M1	K13–M2	K13–M3	DAML02–M1	DAML02–M2	DAML02–M3
$\Delta\mu_\alpha \cos \delta$	0.03	0.04	−0.20	0.80	0.82	0.46	0.19	0.21	−0.10
$\Delta\mu_\delta$	0.01	0.00	−0.21	−0.50	−0.44	−0.80	−0.13	−0.08	−0.37
$\sigma_{\Delta\mu_\alpha \cos \delta}$	0.80	1.01	1.35	3.00	3.08	3.29	1.89	2.03	2.33
$\sigma_{\Delta\mu_\delta}$	0.74	0.91	1.36	3.15	3.31	3.43	1.79	1.92	2.78
$N$	1599	1558	1648	1563	1513	1625	1733	1675	1800

different data sources and methods used for the cluster membership analysis. Again, we notice that most of these differences are related to clusters that do not show a clear overdensity in the UCAC4 proper-motion space at the position indicated in K13.

We observe that D14 do not properly determine mean proper motions for the large and nearby clusters NGC 7092, Ruprecht 147, Mamajek 1, Blanco 1, NGC 752 and Stock 2. For these clusters, M1 and M3 obtain satisfactory results which agree with those published by Dias, Lépine & Alessi (2001) based on the Tycho-2 data (Høg et al. 2000).

Finally, we determine, for the first time, the membership and mean proper motions of 18 new cluster candidates. The candidates correspond to a cluster sample with apparent diameter smaller than 3 arcmin in DAML02. For most cases, there has been no confirmation of the real existence of the clusters through photometric analysis and the new cluster membership study could help to verify this issue.

## 5 CONCLUSIONS

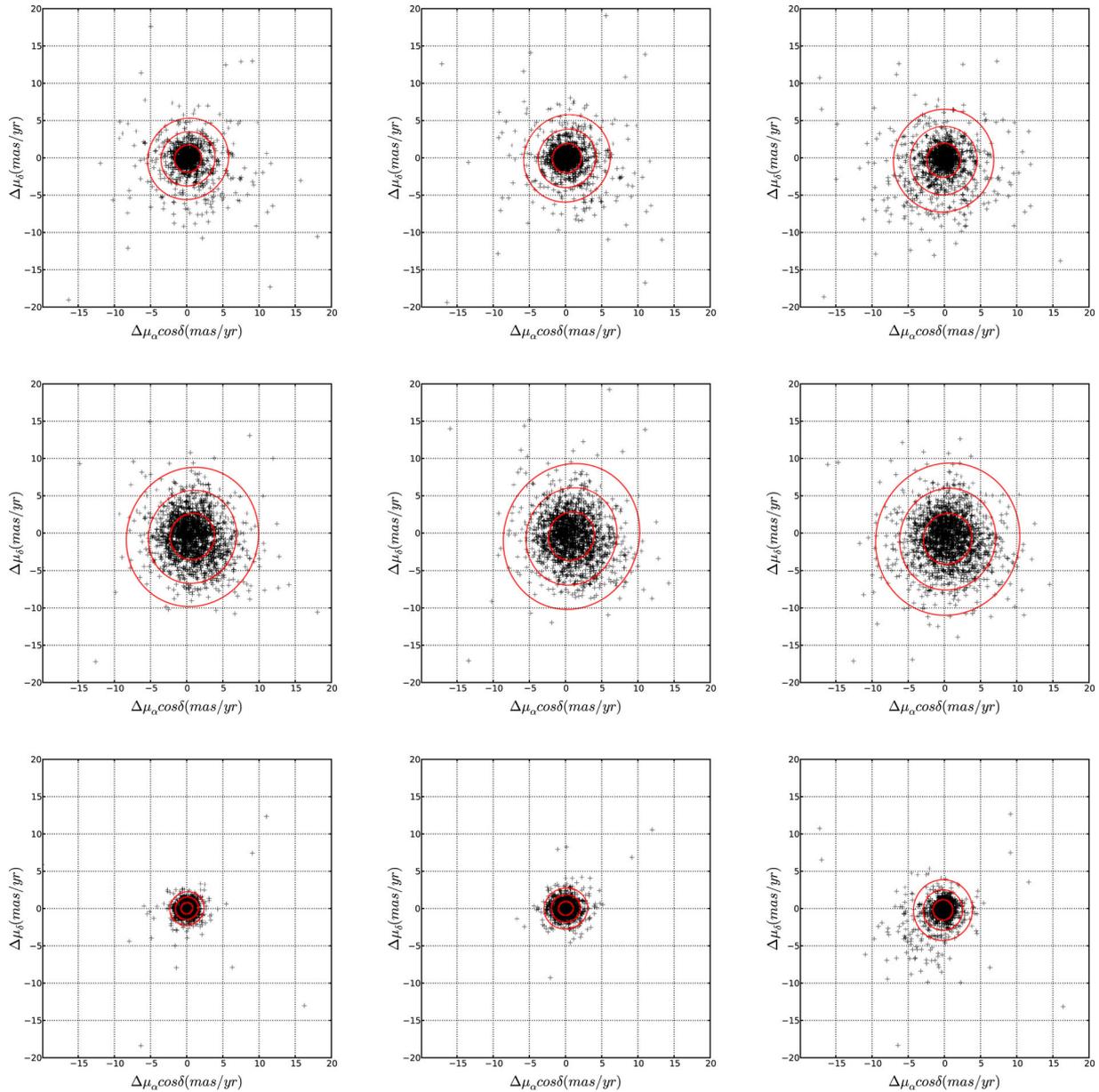
In this paper, we present a homogeneous multimembership and mean proper-motion catalogue of a sample of 1876 open clusters. The fundamental idea underlying this work is the comparison of different cluster membership analysis methods applied to the same set of data. In particular, we wished to check thoroughly, using real data, the method developed by Sampedro & Alfaro (2016) (M1), which enables the utilization of multiple physical variables for cluster membership analysis. The increase in the number of variables in a statistical analysis entails the problem that for the same sample the volumetric density in the  $N$ -variable space decreases drastically as we increase the number of variables, making, in many cases, any later analysis inviable. In Sampedro & Alfaro (2016) (M1), we

avoided this problem by transforming the  $N$ -dimensional space to a one-dimensional space defined by the distance to the distribution centre in that  $N$ -variable space. This method was specially designed for the analysis of *Gaia* data, in which we would have complete information of the phase space for many clusters.

The three different methods used in this study are based on the formulation and analysis of the distribution model of two populations in subspaces of phase space. The model chosen and the way in which the subsequent analysis is carried out can represent different aspects of what we could call a stellar cluster. In most cases, if the astrometric data distributions are well behaved, the results derived from the three analyses coincide. However, clear differences can appear if the spatial and kinematic (proper motions) structure of the cluster deviates from the Vasilevskis model, or if it has a fractal appearance (Sánchez & Alfaro 2009).

The M1 and M2 methods allow the incorporation of new variables for membership analysis, but it is precisely M1 that enables this incorporation in an easier and more elegant way and without preceding hypotheses concerning the variable distribution, except that the cluster members are more densely concentrated than the field stars in that space.

We have redetermined the apparent angular cluster radii directly from the means of the RDPs using the position data from the UCAC4 catalogue. An analysis of the RDPs has enabled us to correct the central coordinates of 10 open clusters. This work presents, for the first time, a catalogue of stellar clusters whose members have been selected through three different methods. It is a study that we consider fundamental for a subsequent comparison with *Gaia* data, but which by itself provides unique information, in quantity and variety. From the initial sample of 1876 clusters, the method M1 successfully converges to a solution for a total of 1748 clusters (93 per cent), M2 for a total of 1693 (90 per cent) and M3 for a



**Figure 3.** Comparison of our results for mean proper motions with those published in the literature. The first line presents the comparison with DAML02, the second line with K13 and the last line with D14. From left to right, the plots present the comparison with the M1, M2 and M3 methods, respectively. The ellipses represent the  $1\sigma$ ,  $2\sigma$  and  $3\sigma$  dispersion levels and their values are given in Table 3.

total of 1819 (97 per cent). The three methods yield membership analysis for a total of 1584 clusters in common (84 per cent).

The comparison of the first moments of the proper-motion distribution for the 1584 clusters that have a solution in the three methods corroborates what has been stated. Furthermore, the mean differences among the methods, for the second moments of the cluster member distribution, are lower than the typical proper-motion errors in the UCAC4 catalogue, overall when comparing M1 and M3. Most of them present differences below the errors catalogued and only a few show clear differences between the estimated values.

We improve the determination of the mean proper motions for the clusters NGC 7092, Ruprecht 147, Mamajek 1, Blanco 1, NGC 752 and Stock 2 for which the D14 analysis failed. For these clusters, M1

and M3 obtain satisfactory results, which agree with those published by Dias et al. (2001). We have determined mean proper motions and cluster membership probabilities for 18 open-cluster candidates for the first time.

The results are available in electronic format at the SSG-IAA (see footnote 3) web site as an unique data base with two main files (more information is given in Appendix A).

## ACKNOWLEDGEMENTS

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**Table 4.** Mean proper motions and dispersions, expressed in  $\text{mas yr}^{-1}$ , for 18 open clusters candidates with unpublished data in DAML02.

Cluster name	M1	M2	M3		$\mu_\delta, \sigma_{\mu_\delta}$	$\mu_\alpha \cos \delta, \sigma_{\mu_\alpha \cos \delta}$	$\mu_\delta, \sigma_{\mu_\delta}$
	$\mu_\alpha \cos \delta, \sigma_{\mu_\alpha \cos \delta}$	$\mu_\delta, \sigma_{\mu_\delta}$	$\mu_\alpha \cos \delta, \sigma_{\mu_\alpha \cos \delta}$	$\mu_\delta, \sigma_{\mu_\delta}$			
AH03 J1725 34.4	-3.40, 1.26	0.74, 1.05	-4.06, 5.40	-0.33, 4.53	-3.46, 1.24	0.79, 1.01	
Alessi 52	-5.78, 5.83	0.08, 4.54	-5.58, 5.07	0.06, 4.27	-5.86, 5.76	0.15, 4.43	
BH 208	–	–	-5.25, 8.11	-1.43, 7.60	-8.44, 1.85	-2.37, 6.59	
DC 1	0.76, 1.68	-0.34, 1.71	0.90, 4.54	-0.98, 4.29	0.36, 1.96	-0.54, 1.99	
Dutra Bica 12	1.27, 4.72	-3.82, 5.07	0.76, 7.78	-3.66, 8.46	1.23, 4.72	-3.81, 5.29	
FSR 0647	-2.64, 2.39	0.70, 3.95	-0.77, 13.53	-1.95, 7.70	-1.90, 6.61	0.36, 5.24	
FSR 0696	–	–	-7.27, 4.39	-3.14, 3.39	–	–	
FSR 0763	–	–	-12.68, 6.52	-4.59, 4.79	-9.33, 2.10	-2.67, 0.78	
FSR 0814	0.42, 2.26	-3.36, 2.04	0.68, 4.49	-3.40, 4.13	-0.90, 1.38	-5.39, 0.52	
FSR 0828	0.97, 3.53	-3.49, 3.29	1.19, 4.83	-3.39, 5.07	0.96, 3.36	-3.50, 3.52	
FSR 1308	–	–	–	–	-3.62, 2.01	1.04, 0.54	
FSR 1343	–	–	-3.30, 4.42	3.84, 3.96	-10.37, 0.45	-3.16, 2.08	
Juchert 10	-4.65, 4.59	-5.37, 4.18	-4.28, 5.00	-4.99, 5.13	-4.89, 2.05	-5.44, 3.51	
Kronberger 23	-0.92, 3.96	-0.44, 2.84	-0.78, 5.74	0.20, 5.33	-1.00, 4.11	-0.55, 2.79	
Majaess 30	0.31, 3.52	-3.27, 3.28	-1.38, 4.90	-2.60, 4.28	-0.71, 2.26	-3.72, 2.20	
Majaess 9	-2.32, 4.66	2.51, 3.53	-2.49, 3.73	1.63, 2.94	-2.32, 4.09	2.54, 3.55	
Teutsch 127	-1.44, 2.29	-2.86, 1.60	-0.26, 4.89	-3.70, 4.79	-3.39, 4.40	-1.99, 0.52	
Wit 3	0.18, 3.40	-3.89, 5.08	1.32, 7.17	-3.81, 5.41	0.16, 1.54	-2.61, 5.71	

us to finish this work in the proper way. LS and EJA acknowledge C. Husillos and T. Gallego for helping us to tailor the SSG-IAA data base. We acknowledge the IAA-CSIC, Universidade Federal de Itajubá (UNIFEI) and IAG/USP for hosting LS during the time this paper was worked on. We acknowledge financial support from the Spanish Ministry for Economy and Competitiveness and FEDER funds through grants AYA-2010-17631, BES-2011-049077, AYA-2013-40611-P and AYA2016-75931-C2-1-P. LS acknowledges financial support from the Brazilian funding agency FAPESP (post-doc fellowship process number 2016/21664-2). AM acknowledges financial support from the Brazilian funding agency FAPESP (post-doc fellowship process number 2014/11806-9). This research has made use of the VizieR catalogue access tool, CDS, Strasbourg, France. The original description of the VizieR service was published in Ochsenein et al. (2000).

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## APPENDIX A

Table A1 describes the parameters in the general catalogue, both for the cluster and the field populations. In addition, it includes distances, ages and colour excess from DAML02. Parameters related to methods not converging to a solution are set to  $-9999.99$ .

We also provide individual cluster-by-cluster files. These include membership probabilities (for M2, the joint and kinematic probabilities are provided), the outlier classification (outliers are set to 1), three additional columns encoding the membership classification of the methods (cluster members are set to 1 and field stars are set to 0) and additional information from UCAC4. Membership probabilities for those cases where a method does not converge to a solution are set to  $-1$ . An example of an individual cluster-by-cluster file is shown in Table A2 for the open cluster NGC 2225.

**Table A1.** Description of the parameters included in the general catalogue for the cluster and the field populations. We also provide the distances, ages and colour excess from the DAML02 catalogue. The subscript  $i$  refers to the three different methods applied in this work.

Position	Parameter	Description	Units	Type
1	Name	Cluster name	–	String
2	RA	Right ascension	deg	Float (5)
3	Dec.	Declination	deg	Float (5)
4	$N_{\text{stars, ini}}$	Number of stars in the cluster region	–	Integer
5	$N_{\text{stars}}$	Number of stars to be analysed	–	Integer
6	$N_{\text{outliers}}$	Number of outliers	–	Integer
7, 8, 9	$N_{M_i}$	Number of cluster members	–	Integer
10, 11, 12	$P_{M_i}$	Percentage of cluster members	–	Float (2)
13, 15, 17	$px_{c, M_i}$	Cluster mean proper motion ( $\mu_{\alpha} \cos(\delta)_c$ )	mas yr <sup>-1</sup>	Float (2)
14, 16, 18	$sigpx_{c, M_i}$	Cluster proper-motion dispersion ( $\sigma_{\mu_{\alpha} \cos(\delta)_c}$ )	mas yr <sup>-1</sup>	Float (2)
19, 21, 23	$py_{c, M_i}$	Cluster mean proper motion ( $\mu_{\delta_c}$ )	mas yr <sup>-1</sup>	Float (2)
20, 22, 24	$sigpy_{c, M_i}$	Cluster proper-motion dispersion ( $\sigma_{\delta_c}$ )	mas yr <sup>-1</sup>	Float (2)
25, 26, 27	$coefcorr_{c, M_i}$	Cluster correlation coefficient ( $\rho_c$ )	–	Float (2)
28	Distance	Distance from DAML02	pc	Integer
29	ColourExcess	Colour excess in BV from DAML02	–	Float (2)
30	Age	Age from DAML02 (in log $t$ )	–	Float (2)
31	Radius	Apparent angular cluster radius from this work	arcmin	Float (2)
32, 34, 36	$px_{f, M_i}$	Field mean proper motion ( $\mu_{\alpha} \cos(\delta)_f$ )	mas yr <sup>-1</sup>	Float (2)
33, 35, 37	$sigpx_{f, M_i}$	Field proper-motion dispersion ( $\sigma_{\mu_{\alpha} \cos(\delta)_f}$ )	mas yr <sup>-1</sup>	Float (2)
38, 40, 42	$py_{f, M_i}$	Field mean proper motion ( $\mu_{\delta_f}$ )	mas yr <sup>-1</sup>	Float (2)
39, 41, 43	$sigpy_{f, M_i}$	Field proper-motion dispersion ( $\sigma_{\delta_f}$ )	mas yr <sup>-1</sup>	Float (2)
44, 45, 46	$coefcorr_{f, M_i}$	Field correlation coefficient ( $\rho_f$ )	–	Float (2)

**Table A2.** Example of an individual cluster-by-cluster file for the open cluster NGC 2225.  $\alpha$  and  $\delta$  are expressed in degrees and  $\mu_\alpha$  and  $\mu_\delta$  in mas yr<sup>-1</sup>. *OutFlag* flags stars classified as outliers (1).  $P_{M_i}$  are the membership probabilities obtained by each method.  $Class_{M_i}$  flags stars classified as cluster member (1) or field stars (0). Membership probabilities, for those cases where a method does not converge to a solution, are set to  $-1$ .

UCAC4 ID	$\alpha$	$\delta$	$\mu_\alpha$	$\mu_\delta$	<i>OutFlag</i>	$P_{M_1}$	$P_{M_2,kinem}$	$P_{M_2,join}$	$P_{M_3}$	$Class_{M_1}$	$Class_{M_2}$	$Class_{M_3}$
402-013194	96.64324	-9.66397	-14.0	-3.7	0	0.01	0.86	0.87	-1	0	1	-1
402-013195	96.64333	-9.62977	-19.5	29.4	0	0.00	0.28	0.67	-1	0	0	-1
402-013199	96.64697	-9.64447	56.5	-5.8	1	0.00	0.00	0.00	-1	0	0	-1
402-013203	96.64961	-9.67991	-58.1	-28.5	1	0.00	0.00	0.00	-1	0	0	-1
402-013204	96.65088	-9.66370	-10.1	-28.2	0	0.00	0.16	0.54	-1	0	0	-1
402-013211	96.65479	-9.67519	-1.4	-8.0	0	0.95	0.84	0.88	-1	1	1	-1
402-013214	96.65607	-9.68540	-4.2	-2.5	0	0.93	0.87	0.99	-1	1	1	-1
402-013216	96.65660	-9.64626	-7.5	-10.8	0	0.37	0.86	0.82	-1	0	1	-1
402-013218	96.65835	-9.61792	-29.2	3.3	0	0.00	0.25	0.00	-1	0	0	-1

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