Demographic consequences for a threatened plant after the loss of its only disperser. Habitat suitability buffers limited seed dispersal JAVIER RODRÍGUEZ-PÉREZ^{1,2}* AND ANNA TRAVESET¹ ¹ Institut Mediterrani d'Estudis Avançats – IMEDEA (CSIC-UIB) Miquel Marquès, 21 E07190 Esporles, Mallorca, Balearic Islands, Spain ² Present address: UFZ Helmholtz Centre for Environmental Research — Umweltforschungszentrum, Department of Ecological Modeling, PF 500136, D04301 Leipzig, Germany * author to whom correspondence should be addressed. E-mail address: jvr.rodriguez@gmail.com

Running title: Demographic consequences in Daphne rodriguezii

- 22 Abstract
- 23

24 Seed dispersal links the end of a plant's reproductive cycle with the establishment of new 25 recruits. Dispersal over short distances may lead to the local aggregation of individuals, 26 slower population growth and, ultimately, to lower population densities. In this study, we 27 analyse the demographic consequences for the shrub Daphne rodriguezii after the loss of its 28 only seed disperser in an island ecosystem (Menorca Island, W Mediterranean). During a 29 period of 8-10 years, we collected demographic data from five populations, four where the 30 disperser is extinct (disrupted) and the only one in which it still persists (undisrupted). We 31 calculated basic deterministic variables, analysed life table response experiments (LTRE) 32 and their covariation among demographic traits, and simulated future population 33 vulnerability. Population growth rate (λ) was either stable or negative and independent of 34 whether the population was disrupted or not. Current and past population dynamics were 35 similar in the two largest populations (one being the undisrupted), which suggests that the 36 environmental conditions allow them to be stable regardless of seed disperser presence. 37 Variation in λ was dependent on rainfall variability and was highly influenced by stasis and 38 growth. There also existed trade-offs between the former life traits and fecundity, which 39 indicate strong competition when resources are limiting (e.g. high plant aggregation due to 40 limited seed dispersal or low rainfall), and that could ultimately affect high-elasticity 41 demographic traits. Our study suggests that the population dynamic of D. rodriguezii is 42 stable under the current conditions, and that where dispersal is limiting, important 43 environmental changes (e.g. in habitat suitability and/or rainfall regime) might lead to local 44 extinctions.

46 KEYWORDS: Covariation between vital rates, plant-disperser interactions, LTREs,

- 47 Matrix population models, Mutualism disruption
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48 Introduction

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50 The phase of seed dispersal links the end of the reproductive cycle of adult plants with 51 the establishment of recruits, and can thus significantly influence the population 52 dynamics of plant species (Wang and Smith 2002; Levin et al., 2003; Howe and Miriti 53 2004; Cousens et al., 2008). All other factors being equal, seed dispersal over short 54 distances may lead to local aggregation of plants, slower population growth and, 55 ultimately, to lower population densities (Levin et al., 2003; Cousens et al., 2008). This 56 pattern may emerge in environments which vary greatly in space since seedlings have 57 very different chances of survival depending on the site they arrive at (e.g. in ecological 58 succession or habitat degradation). Under such circumstances, the biased survival of 59 seedlings in sites to which seeds have been dispersed results in pervasive and positive 60 demographic impacts (Miriti et al., 2001; Godínez-Álvarez et al., 2002; Brodie et al., 61 2009a; Brodie et al., 2009b). Alternatively, mother plants might be concentrated in 62 patches highly favourable for plant survival, suggesting that the disadvantage of local 63 dispersal (i.e. seedlings emerging below adults may have high density-dependent 64 competition) must be weighed against the disadvantage of wider dispersal into less 65 suitable habitats (Cousens et al., 2008).

Vital rates (e.g. survival, fecundity, growth) determine the demographic
parameters that describe the development of individuals throughout life-cycle
transitions (de Kroon et al., 2000). In addition to disentangling which of these
demographic transitions influence plant population dynamics, it is crucial to estimate
those environmental variables linked to any of them (Boyce et al., 2006; Buckley et al.,
2010) and if they buffer or delay the effects of environmental disturbances (Honnay et
al., 2005). In long-lived species, the greater the variation in transitions (e.g. seed-to-

73	seedling transitions), the lower is the elasticity value (defined as the relative change in
74	population growth rate resulting from small relative changes in the matrix elements; de
75	Kroon et al., 2000; Zuidema and Franco 2001) and the less likely that such transitions
76	will influence the population growth rate (Pfister 1998; Doak et al., 2005). Despite
77	stasis (i.e. the probability of surviving and remaining in a given stage from one
78	recording date to the next; Zuidema and Franco 2001) having the highest elasticity of
79	the life-cycle transitions (Silvertown et al., 1996), it is important to estimate the
80	correlation between stasis and the rest of transitions in determining elasticity values
81	(Lambers et al., 1989; Roff 2002). Plant-animal interactions (pollination, seed dispersal,
82	seed predation, etc) influence low-elasticity transitions (fecundity in long-lived species),
83	but it is also possible that their demographic effect on plant population dynamics is
84	stronger than previously thought (Ehrlén 2002; Godínez-Álvarez and Jordano 2007), as
85	fecundity can covariate with high-elasticity transitions (Boyce et al., 2006).
86	Our study is focused on Daphne rodriguezii, an endemic and endangered shrub
87	of Menorca Island (Balearic Islands, W Mediterranean). In this species, seedling
88	survival depends highly on other nurse-plant species (Traveset and Riera 2005;
89	Rodríguez-Pérez and Traveset 2010), and seedlings and adults are highly associated
90	with the abundance of shrub cover (Rodríguez-Pérez et al., in press). This suggests that
91	the suitability of appropriate sites for plant recruitment is key for the stability of
92	population dynamics. Compared to all other Menorcan populations, there is a greater
93	abundance and regeneration in plants inhabiting Colom Islet, a small islet near Menorca
94	(Traveset and Riera 2005). There is no evidence of historical disturbance in these
95	populations (P. Fraga pers. com.) and, thus, Traveset and Riera (2005) suggest that the
96	main cause of population decline in Menorca is the loss of its only seed disperser, the
97	frugivorous lizard Podarcis lilfordi. Most fruits remain undispersed below mother

98 plants in Menorcan populations (disrupted populations, hereafter), whereas in Colom 99 Islet (undisrupted population, hereafter) P. lilfordi defecates seeds in habitats 100 appropriate for seedling establishment (Traveset and Riera 2005; Rodríguez-Pérez and 101 Traveset 2010; Rodríguez-Pérez et al., in press). In the disrupted populations, seedlings 102 recruit under mother plants creating high density clumps which ultimately have a low 103 survival probability (Rodríguez-Pérez et al., submitted). In the close vicinity of adult 104 plants, plant aggregation is much higher in disrupted than in undisrupted populations 105 (Rodríguez-Pérez et al., submitted). We can thus hypothesize that if co-specific 106 competition is strong, the lack of seed dispersers could have a greater negative effect for 107 the population dynamics of the disrupted populations. Much empirical evidence exists 108 in relation to the influence of the failure of mutualists on the early regeneration stages of 109 their plant partners, but there is still little evidence on whether these disruptions cascade 110 into the plant population dynamics (but see, for instance, Christian 2001; Sharam et al., 111 2009; Anderson et al., 2011). Here, we studied the demography of D. rodriguezii over a 112 period of eight to ten years in a total of five populations. The data obtained were used to 113 construct matrix population models in order to answer two questions: a) Do undisrupted 114 and disrupted populations differ in their past, present and future population dynamics? 115 and b) How do life-cycle transitions (stasis, growth, retrogression, fecundity) vary, 116 interact and influence population growth rate?

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118 Material and methods

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120 Study species

Daphne rodriguezii Teixidor (Thymelaeaceae) is an evergreen shrub that reaches
up to 1.5 m in height and lives no longer than 50 years (Rodríguez-Pérez et al. unpubl.

124 data). It inhabits coastal shrublands and grows mostly under shrubs of *Phillyrea latifolia* 125 subsp. media, Pistacia lentiscus and Erica multiflora (Appendix 1). The distribution of 126 D. rodriguezii is highly fragmented and mainly scattered throughout the north-western 127 part of Menorca Island. It is also present in Colom Islet, a small islet (c. 60 ha) very 128 close to Menorca (c. 250 m offshore) and with similar climatic conditions (temperature 129 and precipitation) than Menorca (see Appendix 1). From January to March, new shoots 130 develop, just before flowering. Fruit set is consistently low (<10% of flowers turn into 131 fruits) and selfing is infrequent (Traveset and Riera 2005; Rodríguez-Pérez and Traveset 132 2011). D. rodriguezii reproduces only through seeds. Orange-red drupes develop from 133 May through June and, at Colom Islet, are consumed by the endemic lizard Podarcis 134 *lilfordi*. This frugivore is its only seed disperser and it only coexists with D. rodriguezii 135 there (Traveset and Riera 2005; Rodríguez-Pérez and Traveset 2010). The passage of 136 seeds through the digestive tract of lizards does not affect their germination capacity 137 (Santamaría et al. 2007; Rodríguez-Pérez and Traveset 2010). The proportion of post-138 dispersal seed predation varies between populations (Traveset and Riera 2005), but it is 139 consistently very high (i.e. it is independent of the presence of the seed disperser) and 140 mostly caused by rodents (Traveset and Riera 2005; Appendix 1). All of the unpredated 141 seeds from Menorcan populations are de-pulped by ants and beetles (pers. obs.). 142 However, seeds are left intact, have high chances of germination (pulp removal de-143 inhibits seed germination capacity; Santamaría et al 2007), and are usually left in the 144 close vicinity of mother plants, where survival probabilities are proportionally lower 145 (Rodríguez-Pérez et al., submitted).

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147 Field data collection

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149 We considered five plots within each population which are representative of a) 150 the whole range of the species distribution (Appendix 2), b) the range of population 151 sizes and c) the disperser presence (disrupted vs. undisrupted): four disrupted in 152 Menorca (Porter, Mesquida, Pudent and Favàritx ranging from 50 and 300 individuals) 153 and the only undisrupted population in Colom Islet (with up to 18,000 individuals). For 154 more biological information, see Appendix 1. In 1999, we initially monitored three 155 populations (Porter, Mesquida and Colom Islet). In the first two, we marked (with 156 permanent metallic tags) all the plants that were detected during intensive surveys 157 throughout each plot. Subsequent surveys confirmed we had marked all plants in the 158 population. Due to the high plant density in Colom Islet (about ten-times higher than in 159 disrupted plots, see Appendix 1), we selected a smaller plot of c. 80 m² and we tagged 160 and monitored all individuals in it. This plot appeared representative of a large (and 161 comparable) area in the undisrupted population (see Appendix 1). In subsequent years, 162 we included plots from two other disrupted populations (Pudent in 2000 and Favàritx in 163 2001) where the proportion of monitored plants ranged from 60-75% of the entire 164 population (pers. obs.).

For each plant, we measured height, maximum crown diameter, and fruit crop.
Plant measures were recorded in yearly field surveys before the summer (in early May).
At this time, most fruits begin to ripen (and are quickly removed by lizards in Colom
Islet) whereas new seedlings emerge early in the year (between November and March).
We tagged and measured new recruits in each field survey. A total of 1,073 plants were
marked and monitored during the entire study period. The study concluded with the
2008 survey.

172 **Plant size categorization**

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174	Before building projection matrices, plants were grouped into size-classes. The
175	volume of the cylinder enclosing each plant was used as a biological criterion to
176	categorize plants. This volume was log transformed and grouped into five size-classes,
177	using the k-means clustering method in STATISTICA 6.0 (Statsoft Inc. 2005). We also
178	checked if the classes to which each plant had been assigned were heterogeneous
179	enough by using independent criteria (such as plant survival and mean flower
180	production), and found that they were indeed different [Generalized Linear Models:
181	χ^2 >186.5, d.f = 4, <i>p</i> <0.0001, for both analyses; GENMOD procedure, SAS statistical
182	software]. In the entire study, we used the following size-class boundaries: (1) seedlings
183	(Sdl, no reproduction, smaller than 30 cm ³), (2) juveniles (Juv, <10% reproducing
184	yearly, smaller than $2*10^3$ cm ³) (3) small reproductives (S Rep, <25% reproducing
185	yearly, smaller than $1*10^4$ cm ³), (4) medium reproductives (M Rep, c. 75% reproducing
186	yearly, smaller than $6*10^4$ cm ³) and (5) large reproductives (L Rep, c. 90% reproducing
187	yearly, larger than $6*10^4$ cm ³). <i>D. rodriguezii</i> does not have a seed bank (Traveset and
188	Riera 2005), and this stage was thus not considered as a new size-class.

189 Deterministic demographic analysis – current and past demographic 190 trends and influence on vital rates

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192 For each pair of consecutive years, we constructed a total of 42 projection 193 matrices (seven to nine for each plot). The annual fecundity of each size-class was 194 calculated as the fruit number multiplied by the probability of surviving post-dispersal 195 seed predation and seedling emergence (see Appendix 3). The deterministic population 196 growth rate (λ), or dominant eigenvalues, and their confidence intervals (95%) were 197 obtained following Alvarez-Buylla and Slatkin (1993). For each plot, the stable-class 198 distribution (or dominant right eigenvector) was compared with the observed 199 distribution using a G-test. These estimates were averaged by years. To combine life-

200 cycle and size-class transition, we added matrix elements following the methodology 201 described in Caswell (2001, p. 206-256). The deterministic λ , sensitivity (i.e. dominant 202 eigenvalue to small changes in each of the elements of the projection matrix), elasticity 203 (i.e. sensitivity of the dominant eigenvalue to proportional changes), reproductive value 204 (i.e. left eigenvector of the dominant eigenvalue) and stable-class distribution were 205 calculated with the software PopTools 3.0.3 (Hood 2008). For more information about 206 general methodology applied to matrix population models, see Caswell (2001).

207 Spatial and temporal variability on plant fate: log-linear models

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209 We performed log-linear models to identify the effect of size-class, plot and year 210 on the current plant fate (transition from t to t+1). We built a four-way contingency 211 table of counts of observed transitions, including the following variables: size-class in t 212 (C), plot (P), year (Y) and fate (i.e., size-class in t+1; F). In addition to the five size-213 classes described above, death was included as another class in the fate effect. We 214 considered only hierarchical models, in which the presence of a given interaction 215 implies the presence of all lower-order interactions; for instance, a CYP interaction 216 includes the effects of CY, CP, YP, besides the separate effects of C, Y and P. The 217 model goodness-of-fit (G^2) was calculated by comparing each model with the saturated 218 one (CYP, which includes all the effects). In order to test the within-class variability, we 219 also analysed the effects of plot and year on the fate of each size-class. 220 Contrasts between specific factors were calculated as the difference between G^2 for the two models (ΔG^2). ΔG^2 values closer to zero mean that a given reduced model 221

222 fits as well as the saturated model (i.e. reduced models with smaller ΔG^2 values

223 explained proportionally better the observed plant fate). Because we dealt with a two-

- 224 way factor design (year and plot), each effect was tested in two different ways,
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depending on whether models included only the tested effect (mutual independence,
suggesting that *P* was independent of *Y* and *Y* was independent of *P*) or whether this
was subsequently added to a model that previously did not include it (conditional
independence, which denotes the partial association between *PY* and *YP*). For more
information about this methodology applied to demographic data, see Caswell (2001, p.
326-335).

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231 Stochastic simulations: future trends

233 To estimate the future viability of *D. rodriguezii*, simulations were performed 234 multiplying the distribution of size-classes using a matrix drawn at random each time. 235 We chose this methodology as it preserves the correlation structure among vital rates 236 (Menges 2000). This is a fair assumption, considering that we detected demographic 237 trade-offs between vital rates in our study species (see Results). In order to check the 238 independence of the environmental conditions in the different years measured, we 239 analysed the correlation in rainfall among years. For instance, high-rainfall years could 240 be followed by low-rainfall years (negative autocorrelation), whereas high-rainfall years 241 might occur in strings (positive autocorrelation). In our case, we detected a positive autocorrelation in rainfall but this trend was weak (i.e. Spearman $R^2 = 0.11$, value 242 243 averaged by plots); hence, we consider that environmental conditions of a given year are 244 independent from those of the previous one.

Furthermore, we wanted to separate the effect of population size and its past population dynamics and thus we simulated each plot (500 replicates) under the same initial size (*n*=100) and the size-class distributions (averaged across plots and years). The stochastic λ was then calculated as log λ s=(log N_t-log N₀)/t (Caswell 2001, p. 387-401), N₀ and N_t being the initial (observed) and the after-simulation plot sizes,

250	respectively. The proportion of iterations that led to population extinction (i.e. all plants
251	of a given plot died) was used as a proxy to calculate its vulnerability to future
252	extinctions. Simulations were performed using a specified 'visual program'
253	implemented in Microsoft Excel (K. Lehtilä, X. Picó). Since this procedure does not
254	include demographic stochasticity and density-dependent survival (see Brook 2000 for
255	discrepancies and concordances between observed and estimated projections) we
256	projected plots up to a short-time scale (up to the maximum age of older individuals, see
257	Study species). Unless otherwise indicated, average values are reported as mean \pm
258	standard error (± 1 SE).

259 Life table response experiments (LTREs): decomposing spatial and temporal variability of population dynamics 260

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262 Life table response experiments (LTREs) allow the decomposition of the 263 observed variance of λ into life-history transitions (Caswell 2001, p. 258-

264 278). This μετηοδ αδδρεσσεσ τηε εξτεντ το ωηιχη εαχη ματριξ εντρψ ισ ρεσπον

265 σιβλε φορ χηανγεσ ιν τηε λ value of a given matrix, as compared to the λ value of the

266 average matrix. In the present study, 'plot' and 'year' were considered and analysed as

267 fixed effects due to the different intrinsic traits of each plot (presence of seed disperser,

268 population size) and year (specific abiotic conditions, see Appendix 1). Effects of main

factors and their interactions were decomposed into 'contributions' (i.e. deviance of the 269

270 corresponding matrix element to the mean matrix; Caswell 2001, p. 261). For each life-

271 history trait, we first assessed whether the overall contribution forof each matrix (42)

- 272 was correlated with its \Box . In our case, we also calculated the relative λ of each matrix
- 273 $(\Delta \lambda)$ based on: a) differences among plots within the same year (i.e. the λ value and the

274 yearly averaged λ), and b) differences among years within the same plot (i.e., between 275 each \Box value and the \Box averaged \Box by plots).

276 Given that rainfall is one of the most variable environmental factors in arid and 277 semiarid environments and may influence the population dynamics of long-lived 278 species (Marrero-Gómez et al., 2007), we further explored if the temporal variability in 279 'contributions' was associated to annual rainfall. Hence, we correlated the yearly-280 decomposed contributions with the averaged annual rainfall of the five studied plots; 281 rainfall data came from the closest meteorological station of each plot (i.e. not farther 282 than 5 km away; data from the Balearic Meteorological Institute). 283 In addition to computing the overall effect, 'contributions' were summed across 284 positive and negative values (net effects) separately for life-history transitions and plot. 285 This dissecting procedure provides a clearer biological picture of vital rates, since it 286 explores how positive and negative contributions compensate each other, and highlights 287 hidden trade-offs between life-history elements (Jongejans and de Kroon 2005). 288

289 Results

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291 **Population dynamics: past, present and future trends**

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In all plots, the most abundant individuals were those in the reproductive size-classes (>50%; Fig. 1). The plots with the highest proportions of reproductives (63%, on average) were all disrupted, whereas the undisrupted and one of the disrupted plots (Porter) had a high proportion of seedlings and juveniles (37%; Fig. 1). Plots from the two largest populations (Pudent and Colom Islet) had similar observed and stable-class distributions (Fig. 1). In the other plots, there was discordance between observed and

stable-class distributions, suggesting that these plots were currently in transientdynamics.

301 Pooling all size-classes, plant fate was found to be rather variable and such variation was better explained by the plot effect (i.e. ΔG^2 values were smaller; Table 1). 302 303 A plot effect was found to be more important in seedlings' and small reproductives' fate, 304 whereas the year effect was the most important for juveniles. For small reproductives 305 and juveniles, there was no year effect (Table 1). In the case of medium and large 306 reproductives, conditional dependence explained plant fate better (i.e. conditional ΔG^2 307 values were smaller than marginal values; Table 1). 308 During the study period, deterministic λ either decreased or was stable in each 309 plot, regardless of the absence of the disperser (Fig. 2). Mesquida and Pudent had, on 310 average, relatively high values of λ due to their high stasis of reproductive adults 311 (Appendix 3). By contrast, the other plots had low values of λ , due to the low survival 312 and high retrogression of large reproductives (Porter and Colom Islet), low survival of 313 small reproductives and juveniles (Porter) and low fecundity (Favàritx; Appendix 3). 314 The lowest λ in Colom Islet was found in 2004 (see Fig. 2), due to an unusual low stasis 315 (and survival) of medium reproductives (c. 38% lower than average); the λ for this year 316 was likely extreme for the population dynamics of this population due to non-different 317 observed and stable-class distributions (Fig. 1).

318 Stochastic simulations showed that the vulnerability for the next 50 years was 319 moderate on Colom Islet (21%). When we excluded the projection matrix of 2004 in the 320 simulations, the vulnerability decreased up to 5.8% in the latter plot. Vulnerability was 321 also moderate in two out of four disrupted plots (Porter: 38.0% and Favàritx 18.6%)

322 whereas it was almost null (<0.1%) in the others (Pudent and Mesquida).

324 Spatial-temporal variability, interaction among vital rates and their 325 influence on λ

327 As in many long-lived species, stasis accounted for most of the elasticity $(90.9\% \pm 3.0)$; 328 Appendix 3) and the lowest variability of life-cycle transitions (Appendix 4). By 329 contrast, fecundity had high variability but low elasticity. Reproductive plants were the 330 size-classes with the highest elasticity ($88.6\% \pm 2.23$), specifically, medium ($37.4\% \pm 5.7$) 331 and large reproductives (29.5%±6.2). Both size-classes accounted for the lowest 332 variability in stasis (1.31 and 1.36, respectively; Appendix 4). 333 Overall, LTREs showed that the observed variability in deterministic λ mostly 334 depended on the 'plot x year' interaction (44.8%), followed by year (33.2%) and plot 335 (21.9%). The difference between observed and modelled λ (i.e. calculated from the 336 decomposed contributions of each life-history transition) was rather low $(1.54\% \pm 0.16)$, 337 indicating that the two-way decomposition of the variance fitted well with the observed 338 λ . When we decomposed the relative contribution of each year and correlated it to 339 annual rainfall, we did not observe any trend (Spearman R=0.233, p=0.552). However, 340 when we removed the year 1999 (there was an inverse trend between contribution for 341 that year and rainfall; see Fig. 4a) and we repeated the analysis, we detected a 342 significant association between both variables (Spearman R=0.762, p=0.0368) 343 suggesting that an important fraction of the contribution variability depended on annual 344 rainfall. According to the variability in life-cycle transitions, only the $\Delta\lambda$ regarding the differences among plots was significantly associated with the relative contribution of 345 346 each transition (Table 2). Specifically, the $\Delta\lambda$ was positively correlated to the 347 contributions of stasis and growth (Fig. 3). Such variation did not depend on rainfall 348 variability since it was not correlated to contributions for either life-cycle transition (see

Table 2). We also detected trade-offs between contributions: fecundity was correlatednegatively and positively, respectively, to stasis and retrogression (Appendix 4).

351 Within plots, the magnitude of contributions by life-cycle transitions was highly 352 variable. Stasis was the most important element in two disrupted plots (Mesquida and 353 Favàritx); in the former, this life-cycle transition was responsible for its relatively high 354 value in λ (Fig. 4c) whereas in the latter, its positive contribution was unbalanced by the 355 negative effect of the rest of life-cycle transitions (Fig. 4e). In Colom Islet, by contrast, 356 contributions were negative in most life-cycle transitions (especially stasis and growth), 357 which finally resulted in an overall relatively low contribution (Fig. 4f). The positive 358 contributions of stasis and growth on Pudent were responsible for its overall positive 359 effect (Fig. 4d). Despite the fact that fecundity contribution was strongly positive on 360 Porter, the relatively low plot effect was due to its low stasis and growth (Fig. 4b).

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362 Discussion

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How did the variability and interaction among vital rates affect the population dynamics of *D. rodriguezii*?

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367 Persistence of plant populations greatly depends on the fluctuations of environmental 368 conditions (Menges 2000; Doak et al., 2005; Boyce et al., 2006; Salguero-Gómez and 369 de Kroon 2010; Buckley et al., 2010). In the case of D. rodriguezii, plant fate varied 370 much among plots for seedlings and medium reproductives (both pooled), but showed 371 to be independent of environmental variability (year or plot) for juveniles and small and 372 large reproductives. Despite most of the observed variability in deterministic population 373 growth rate (λ) was intrinsically "random" (i.e. 45%), an also important fraction (33%) 374 was explained by temporal variability. After decomposing the λ variability among 375 years, we detected that the relative contributions depended on rainfall; in other words,

376 lower and higher rainfall than average (c. 550 mm) cascaded into a negative and 377 positive λ , respectively, in *D. rodriguezii*. Our results agree with those of Marrero-378 Gómez et al. (2007) who found that the population dynamics of the narrow endemic 379 shrub Helianthemum juliae in the Canary Islands is also affected by rainfall. These 380 authors argue that rainfall (highly variable in arid and semiarid environments) could be 381 an environmental factor influencing the population dynamics of long-lived species. The 382 greater temporal variation in plant populations is partly scale-dependent due to similar 383 environmental factors in spatially close populations (Buckley et al., 2010). In our case, 384 study plots were located less than 40 km apart from each other, suggesting that 385 environmental conditions (in this case, rainfall) are equivalent and potentially 386 meaningful for the population dynamics of the five studied populations of D. 387 *rodriguezii*. Our findings further suggest that λ was also explained by an important 388 fraction of its variability decomposed by life-cycle transitions, specifically, its 389 dependence on stasis and growth. Our data did not allow us to detect correlations 390 between rainfall and λ variability decomposed by life-cycle transitions, suggesting that 391 the contribution of vital rates to λ is compensated within plots (Jongejans and de Kroon 392 2005).

Even though the rest of low-elasticity life-cycle transitions (growth, retrogression and fecundity) contribute relatively little to λ , we detected demographic trade-offs among them at the plot level. In Mesquida and Pudent, for instance, positive contributions in stasis and/or growth compensate the negative contributions for the rest of life-cycle transitions, and λ was thus higher than average in both plots. By contrast, the relatively low λ for the rest of the plots resulted from a negative contribution in most life-cycle transitions. Due to resource limitation, physiological allocation in one vital

400 rate may also be detrimental to the other vital rates (Lambers et al., 1989; Roff 2002). In 401 D. rodriguezii, positive contributions to fertility were weakly associated with negative 402 and positive contributions to stasis and retrogression, respectively. Fertility is the most 403 variable life-cycle transition in D. rodriguezii (as it occurs in the rest of long-lived 404 species; Silvertown et al., 1996) which suggests that high investment of resources in 405 fruit production during the reproductive season could strongly influence the rest of the 406 life-cycle transitions. Our hypothesis could thus be supported at the plot level. For 407 instance, Porter showed the highest contribution to fertility but the lowest λ of the five 408 studied plots, probably due to its high fruit production per plant (Appendix 1) and low 409 stasis of reproductives (Appendix 2). By contrast, Pudent and Mesquida have 410 comparatively low fruit production (Appendix 1), which probably translates into high 411 stasis and the most stable λ of the studied populations.

412 When individuals share the same environment, vital rates could be directly 413 correlated to the close environmental conditions (Doak et al., 2005). Lack of seed 414 dispersal leads to local aggregation of recruits and adults, which in turn leads to higher 415 competition for space and resources (Harper 1977). Here we detected negative 416 correlations in the variance of life-cycle transitions, interactions which could be 417 magnified when recruits compete for resources with adults in "good" years (when 418 rainfall is higher than average; Rodríguez-Pérez and Traveset 2010). This is what 419 probably occurs in our disrupted populations, since recruits achieve high local densities 420 and have proportionally lower probabilities of survival in the close proximity of 421 reproductives (Rodriguez-Perez et al. submitted). Thus, our findings point out to the 422 necessity of studying also the variability of ecological determinants that affect vital rates 423 when predicting its demographic consequences into plant species (Salguero-Gómez and 424 de Kroon 2010).

Is the loss of its disperser important for the population dynamics of *D. rodriguezzi*?

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428 Despite the broad evidence that seed dispersal is important for plant population 429 dynamics (Wang and Smith 2002; Levin et al., 2003; Howe and Miriti 2004), there are 430 few examples examining its effect within the entire plant's life cycle. In those cases, 431 seed dispersal 'matters' for population dynamics because the frugivores dispersing more 432 seeds increase their germination capacity and/or move them also to high quality sites for 433 seedling establishment (Godínez-Álvarez et al., 2002; Brodie et al., 2009a; Brodie et al., 434 2009b). In the present study, however, we found that λ of the undisrupted population is 435 similar to that in half of the disrupted ones. This is probably because short-term 436 regeneration processes such as emergence (i.e. ingestion does not affect seed 437 germination) and seedling survival (i.e. survival is consistently high under nurse plants) 438 do not depend on the presence or absence of the seed disperser (Traveset and Riera 439 2005; Rodríguez-Pérez and Traveset 2010). 440 In spatially heterogeneous environments, adult plants are likely to be located in 441 patches which are highly favourable for survival and, thus, local dispersal may be more 442 advantageous than longer dispersal (as habitats might be less suitable) for the 443 population dynamics (Cousens et al., 2008). Regardless of the presence of the disperser, 444 reproductives of D. rodriguezii are mostly concentrated under nurse shrubs (Rodríguez-445 Pérez et al., in press), which happen to be the best places for seedling survival (Traveset 446 and Riera 2005; Rodríguez-Pérez and Traveset 2010). We also observed that λ remains 447 approximately stable in two disrupted plots (Mesquida and Pudent) and that it is 448 independent of population size (Appendix 1). Since environmental variability is a 449 measure to predict species population dynamics in heterogeneous landscapes (Wiegand 450 et al., 1999; Chisholm and Wintle 2007), it is likely that the two latter disrupted

populations remain demographically stable because habitat features in them are
appropriate for the species persistence. Long-living forest plants can have remnant
populations not in equilibrium with the current degree of habitat fragmentation (Honnay *et al.*, 2005), which is attributed to the still optimal conditions in habitat fragments. In
our case, however, three of the four disrupted populations are in transient dynamics (see
Fig. 1), which suggests that their current population instability might be due to changes
in habitat quality.

458 Where lizards still persist, D. rodriguezii maintains high levels of plant 459 regeneration (i.e. seedlings and juveniles are proportionally more abundant than 460 reproductives) and a stable λ (if we exclude the unusual year, 2004). In Colom islet, 461 most plants are located in high-quality sites (Rodríguez-Pérez et al., in press) and they 462 are less aggregated at small distances than in the populations without lizards 463 (Rodríguez-Pérez et al., in prep). This allows the islet population to respond to any 464 environmental disturbance, because dispersers can move seeds into high-quality habitat 465 patches. A revealing study by Miriti et al. (2001) detected that changes in survival of 466 isolated adults of a perennial desert shrub (despite lower success during early stages of 467 recruitment) have greater influence on λ than changes in survival of adults close to 468 neighbours. Consequently, the importance of seed dispersal for the population dynamics 469 of long-lived plants may emerge in the long-term, and this is independent of whether its 470 effects on the population dynamics are positive (Godínez-Álvarez et al., 2002; Brodie et 471 al., 2009a; Brodie et al., 2009b), neutral (present work) or negative (Miriti et al., 2001). 472 The few studies examining whether the failure of mutualisms has implications 473 for plant population dynamics, point to strong cascading effects (Christian 2001;

474 Sharam et al., 2009; Anderson et al., 2011). In our case, such an effect does not appear

to be so strong, at least from a demographic viewpoint. D. rodriguezii is highly

476 associated with nurse plants (Traveset and Riera 2005; Rodríguez-Pérez et al., in press), 477 and positive plant-plant interactions seem thus one of the main forces governing its 478 population dynamics. This leads to assume that its populations could remain stable 479 within the current environmental conditions, independently of the seed-disperser 480 presence. Yearly rainfall (as a *proxy* of resource availability) includes demographic 481 "noise" in the population dynamics of D. rodriguezii: when rainfall is below the 482 average, λ declined (present work) but there is no competition of recruits beneath adults 483 (Rodríguez-Pérez & Traveset 2010). This suggests that plant-plant interactions buffer 484 population dynamics when resources are limiting, allowing a proportionally higher 485 survival of recruits when seed disperser is absent.

486 Since environmental factors strongly influence the population dynamics of 487 perennial plants (Buckley et al., 2010), it is thus likely that such demographic effects 488 influence D. rodriguezii by changes in the abundance of high-quality habitats. If this is 489 the case, seed dispersal would matter (Howe and Miriti 2004) because it would be a 490 demographic process helping to stabilize population dynamics in case of environmental 491 variability (Levin et al., 2003; Cousens et al., 2008). In disrupted populations, plants 492 currently live in high-quality 'islands' of habitat (i.e. habitat features fulfil the 493 ecological conditions to maintain stabilized populations) and, having lost the seed 494 dispersal capacity, they are not able to move between habitat patches (Rodríguez-Pérez 495 et al., in press). Consequently, plant species may face local extinctions under moderate-496 to-strong changes in habitat quality (Cousens et al., 2008).

In short, the demographic consequences on *D. rodriguezii* may not be so
dramatic after the loss of its seed disperser, provided that habitat quality remains stable.
This may create time-lags to population extinctions that can take up to thousands of
years to come into completion (Sax and Gaines 2008).

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503

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- 609 Zuidema, P. A. and Franco, M. 2001. Integrating vital rate variability into perturbation
- 610 analysis: an evaluation for matrix population models of six plant species. Journal of
- 611 Ecology 89: 995-1005.

- 612 Table 1 Log-linear analyses for the effect of plot (P) and year (Y) on the plant fate,
- 613 including the matrices decomposed by plant size-classes. Each effect over plant fate was
- 614 explored as the difference between the dependent-effect model and the saturated model
- 615 (see material & methods). The effects of including another factor under parenthesis test
- 616 for conditional independencey (see Caswell 2001, p. 326-335); i.e., P (Y) means that the
- 617 effect of plot (P) when added to a model already contained years (Y). In all analyses, the
- 618 P x Y interaction effect was never significant ($\Box G2 > 58.0$, p = 1.000, for all analyses).
- 619 *** p <0.001 ** p <0.01 *p <0.05 † p <0.1

	All Stages		Seedling		Juvenile		S Rep		M Rep		L Re	р
Effect	∆df	ΔG^2	∆df	ΔG^2 Δdf ΔG^2		∆df	ΔG^2	∆df	ΔG^2	∆df	ΔG^2	
Y	200	226.7†	40	70.0**	40	34.4 n.s.	40	38.0 n.s.	40	61.0*	40	23.7 n.s.
Р	100	196.3***	20	34.7*	20	47.2**	20	27.9 n.s.	20	34.4*	20	52.1***
P (Y)	100	204.9***	20	43.5**	20	52.7***	20	28.8†	20	31.6*	20	48.3***
Y (P)	200	235.4*	40	78.8***	40	39.5 n.s.	40	38.9 n.s.	40	58.2*	40	19.9 n.s.

622	Table 2 – Spearman-rank correlations between the contributions of each matrix ($n =$
623	42), in respect to (a) λ , (b,c) the relative λ and (d) annual rainfall. The relative λ of each
624	matrix ($\Delta\lambda$) was calculated following two criteria: b) in respect to differences between
625	each λ value and the averaged λ by plots within the same year, and c) to differences
626	between each λ value and the yearly-averaged λ within the same plot. Additionally, we
627	correlated the contributions for each matrix with the annual rainfall. For more
(20	information and M. (. 1 . 1 M. (. 1

628 information, see *Material and Methods*.

	Life-cycle transition	R	t (n-2)	р
a) λ	Stasis	0.210	1.36	0.181
	Fecundity	-0.108	-0.688	0.495
	Growth	0.259	1.70	0.098
	Retrogression	0.198	1.28	0.209
b) $\Delta\lambda$ by plots	Stasis	0.647	5.37	<0.0001
	Fecundity	-0.015	-0.094	0.926
c) Δλ by years	Growth Retrogression Stasis	$0.420 \\ 0.065 \\ 0.244$	2.92 0.411 1.59	0.006 0.683 0.119
	Fecundity Growth	0.038	0.242 0.586	0.810 0.561
d) Rainfall	Retrogression	0.246	1.61	0.116
	Stasis	0.036	0.229	0.820
	Fecundity	0.046	0.290	0.773
	Growth	-0.034	-0.217	0.829
	Retrogression	0.160	1.028	0.310

631 Figure captions

632	Figure 1 - Average distribution of individuals belonging to the observed (black bars)
633	and the stable-stage (white bars) size-classes. For each plot, we calculated the
634	significance between observed and stable-stage proportions. Year variability is
635	denoted in error bars (± 1 SE). Abbreviations: Sdl: seedlings, Juv: juveniles, S, M,
636	L Rep: small, medium and large reproductives.
637	
638	Figure 2 - Deterministic population growth rate (λ ; \pm 95% CI) across plots and years.
639	Confidence intervals were calculated following Alvarez-Buylla and Slatkin
640	(1993). Dashed line represents the λ at stability (i.e. $\lambda=1$). For each plot, we also
641	showed the average λ across years.
642	
643	Figure 3 - Contribution of each vital rate to the variance of the population growth rate
644	(X axes), in respect to the relative population growth rate ($\Delta\lambda$, Y axes). For each λ
645	value, we calculated its $\Delta\lambda$ in respect to the averaged λ by plots, since it explained
646	the contributions of two out of four vital rates (Table 2). Note that the values of
647	each plot are denoted with different symbols.
648	
649	Figure 4 - Contributions to the average matrix in λ (a) across years and (b-h) vital-rate
650	transitions. For years, we plotted contributions yearly-decomposed (bars and left
651	Y-axes); in other words, positive contributions (white bars) meant a proportionally
652	higher value for a given year than the average contribution. We also showed the
653	mean annual rainfall averaged by plots (grey line and right Y-axes); the average
654	value for the whole studied period was c. 550 mm. For vital-rates [i.e., fecundity
655	(F), stasis (S), growth (G) and retrogression (R)], the total contribution for each

- 656 plot (γ) was decomposed and grouped into positive and negative contributions
- 657 (white and black bars, respectively).

658 Appendix 1 – Differences between plots in some abiotic and biotic traits. Disrupted 659 plots (i.e. without the presence of the seed disperser: Porter, Mesquida, Pudent and 660 Favàritx) are located in Menorca Island whereas the undisrupted plot (i.e. the disperser 661 is present there; Colom Islet) was in an Islet nearby Menorca (c. 250 m offshore). 662 Climatic information (rainfall and temperature) came from the closest meteorological 663 station of each plot (no longer than 5 km; data from the Balearic Meteorological 664 Institute) and soil variables (conductivity, organic matter and total nitrogen [N: 665 Kjeldahl's method]) from 5 soils samples collected up to 5cm in depth; samples were 666 sampled and analysed in 2002 by the Agriculture Department of the Balearic Islands. We also estimated plant plot size, plant density (in each study site plot), percentage of 667 668 plant occurrence under shrubs and fruit crop and fruit set (averaged for each year); these 669 estimates were based on the number of reproductive plants. Shrub abundance was 670 calculated from a habitat map (1:1000 cartography) included in a GIS platform; this 671 map have the following categories: open habitat (i.e. bare soil and rocks) and shrubs 672 (i.e., Phillyrea latifolia, Pistacia lentiscus, Erica multifolia, Rosmarinus officinalis and 673 Santolina chamaecyparisus). In previous studies, we estimated post-dispersal seed 674 predation and seedling emergence by field experiments (Traveset & Riera 2005; 675 Rodríguez-Pérez & Traveset 2010); the average values of these experiments are shown 676 here.

Abiotic and biotic traits	Porter	Mesquida	Pudent	Favàritx	Colom Islet
Abiotic					
Plot sampling area (m ²)	7480.2	5741.0	4142.2	2959.3	80.0
Rainfall per year (mm)	541.8	548.4	579.4	440.8	589.5
Average temperature (Celsius)	16.5	16.9	16.8	16.9	16.8
Facing slope	S	W	W	W	W
Average distance to sea (m)	50	110	90	80	50
Soil conductivity (dS/m)	2.28	1.63	3.30	1.09	5.08
Organic matter (%)	8.66	8.06	2.70	2.44	3.50
Total Nitrogen in soil (mg					
N/L)	0.67	0.61	0.16	0.14	0.22
Biotic					
Plant plot size	48	47	298	168	c.18,000
Plant density (Plant/m ²)	0.033	0.041	0.045	0.051	0.447
Shrub abundance (%)	63.7	68.5	86.4	86.9	74.4
Plant occurrence under shrubs	57.2	80.5	81.8	73.5	81.2
(%)					
Fruit crop	$9.4{\pm}7.4$	4.3±4.8	4.9 ± 4.9	6.0 ± 7.3	6.0 ± 4.8
Post-dispersal seed predation					
(%)	78.4	60.7	91.5	99.5	97.3
Seedling emergence (%)	26.9	43.9	23.0	15.7	42.1
Presence of seed disperser	No	No	No	No	Yes

677

679 In a larger area (i.e. 3619.2 m^2) and a few meters apart from the monitoring plot

680 in Colom Islet, we also measured the abundance of plants by size-classes (see

681 Rodríguez-Pérez et al., in press). There, we found an also high plant density (this time,

682 six-times higher than disrupted plots) and that the size-class plant abundance (see Fig.

683 1) did not differ between those two plots ($\chi^2_{10} = 10.0$, p = 0.265). This indicates that our

684 monitored (but small) undisrupted plot is representative of the population dynamics of a

685 large area in our undisrupted population (Colom Islet).

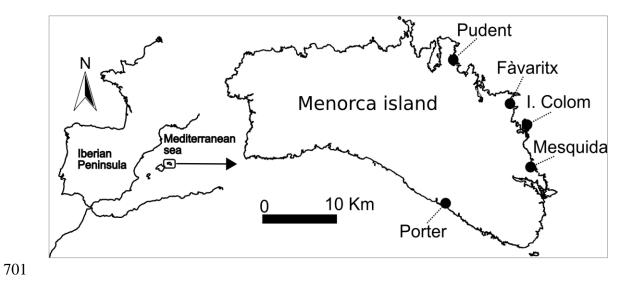
686

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690	its effects on a threatened endemic plant in the Balearic Islands Conservation Biology
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- 694 plant–lizard interaction and its consequences for plant regeneration after disperser loss.
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- 696
- 697 Rodríguez-Pérez, J., et al. (in press) Frugivore behaviour determines plant distribution:
- 698 a spatially-explicit analysis of a plant-disperser interaction. Ecography

- 699 Appendix 2 Location of study sites of *D. rodriguezii*. Black dots show names and
- 700 locations of study populations in the Menorca Island.



Appendix 3 - Average transition and elasticity elements per each year, plot, life-cycle and size-class of *D. rodriguezii*. For each matrix, each column (j) represents the transition p_{ij} of a given size-class at time t and the rows (i) at time t+1. p_{ii} is the transition of remaining in the same size-class (stasis), $p_{i>j}$ to regressing to a smaller sizeclass (retrogression) and $p_{i<j}$ of promoting to a higher size-class (growth).

708 Fecundity, or the averaged number of seedlings that each size-class contributed 709 to each matrix, was calculated after accounting for the fruit crop per each size-class, the 710 probability of post-dispersal seed predation and the seedling emergence (see Appendix 711 1). When this information was not available from a particular year, we used the average 712 of other years in that plot. In all plots, seed germination experiments were run from 713 2000 and 2005 (Traveset & Riera 2005; Rodriguez-Perez & Traveset 2011) whereas 714 post-dispersal seed predation experiments from 2002 to 2005 (Traveset & Riera 2005; 715 unpubl. data).

716 In some years, some transitions were not observed in a given plot. For instance, 717 reproductive individuals frequently did not grow and/or retrogress in a particular year; 718 in other words, all individuals remain alive or (if any) become dead. In consequence, we 719 modify matrices in order to include those unobserved events according to a conservative 720 approach. These transitions were estimated following $p_{ij} = 1/n$, in which *n* represents the 721 sample size of a given size-class in each matrix. The maximum transition value to be 722 modified was 0.1. In case that both growth and retrogression were not observed for a 723 given size-class, p_{ij}^* was evenly distributed across life-cycle transitions. We only 724 modified those transitions that grew and retrogressed only one stage (i.e. j+1, j-1, 725 respectively). Finally, p_{ii} was removed from stasis since each column have to sum 1.0.

	Stage at Transition probabilities									Elasticit	у	
	time t+1	Sdl	Juv	S Rep	M Rep	L Rep	_	Sdl	Juv	S Rep	M Rep	L Rep
Р	Sdl	0.3270	0.0165	0.1039	0.2690	1.3916		0.0284	0.0006	0.0011	0.0022	0.0066
o r	Juv	0.1315	0.7132	0.0513	0.0069	0.0000		0.0105	0.1139	0.0033	0.0003	0.0000
t	S Rep	0.0000	0.0569	0.7279	0.0450	0.0000		0.0000	0.0135	0.0869	0.0069	0.0000
e r	M Rep	0.0000	0.0000	0.0602	0.8317	0.1057		0.0000	0.0000	0.0160	0.4540	0.0151
1	L Rep	0.0000	0.0000	0.0000	0.0533	0.7900		0.0000	0.0000	0.0000	0.0217	0.2190
М	Sdl	0.5427	0.0000	0.0456	0.0562	0.1610		0.0054	0.0000	0.0015	0.0012	0.0025
e s	Juv	0.1647	0.7963	0.0594	0.0222	0.0175		0.0052	0.1491	0.0106	0.0022	0.0013
q	S Rep	0.0000	0.1167	0.8653	0.0556	0.0000		0.0000	0.0193	0.2828	0.0125	0.0000
u i	M Rep	0.0000	0.0000	0.0568	0.8248	0.0880		0.0000	0.0000	0.0197	0.2473	0.0151
d												
а	L Rep	0.0000	0.0000	0.0000	0.0638	0.8556		0.0000	0.0000	0.0000	0.0189	0.2053
Р	Sdl	0.3958	0.0000	0.0000	0.0631	0.1647		0.0000	0.0000	0.0000	0.0028	0.2055
u	Juv	0.3938	0.0022	0.0270	0.0031	0.1047		0.0003	0.0000	0.0011	0.0028	0.0004
d e	S Rep	0.1037	0.1384	0.0340	0.0042	0.0000		0.0102	0.0493	0.1303	0.0009	0.0000
n	M Rep	0.0000	0.1384	0.1342	0.0313	0.0000		0.0000	0.0000	0.1303	0.3616	0.0000
t	-	0.0000	0.0000	0.1342	0.8202	0.8820		0.0000	0.0000	0.0000	0.0270	0.0200
F	L Rep							-				
à	Sdl	0.5671	0.0002	0.0019	0.0100	0.0157		0.0020	0.0000	0.0000	0.0004	0.0007
v	Juv	0.1075	0.7435	0.0487	0.0000	0.0000		0.0011	0.0352	0.0072	0.0000	0.0000
a r	S Rep	0.0000	0.1210	0.7579	0.0517	0.0000		0.0000	0.0082	0.2413	0.0129	0.0000
i	M Rep	0.0000	0.0000	0.0925	0.8347	0.0428		0.0000	0.0000	0.0140	0.2575	0.0115
t x												
л	L Rep	0.0000	0.0000	0.0000	0.0440	0.9037		0.0000	0.0000	0.0000	0.0121	0.3959
С	Sdl	0.4272	0.0018	0.0165	0.0554	0.2804		0.0093	0.0001	0.0007	0.0049	0.0032
0 1	Juv	0.1595	0.8029	0.0336	0.0103	0.0058		0.0090	0.1768	0.0050	0.0013	0.0000
0	S Rep	0.0000	0.0738	0.7264	0.0411	0.0184		0.0000	0.0150	0.1505	0.0150	0.0009
m	M Rep	0.0000	0.0030	0.1464	0.8443	0.1400		0.0000	0.0002	0.0245	0.4250	0.0150
	L Rep	0.0000	0.0000	0.0048	0.0377	0.7491		0.0000	0.0000	0.0006	0.0185	0.1246
	Stage at time t								S	Stage at tin	ne t	

728

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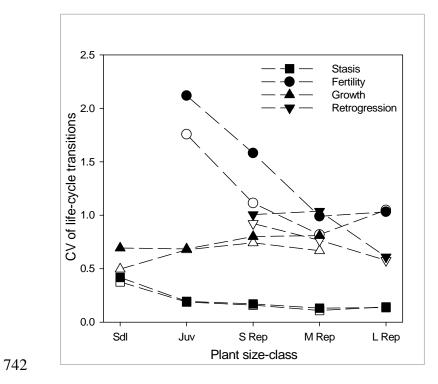
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- 737 Appendix 4 Coefficient of variation (CV) of life-cycle transitions for each plant size-
- class. CV was calculated by plots (black symbols) and by years (white symbols). Stasis,
- 739 growth and retrogression values came from transitions of matrix elements; fertility
- values from the number of seedlings (i.e. the fruit crop after including the probability of
- 741 post-dispersal seed predation and seedling emergence).



743 Appendix 5 – Correlations of contributions between life-cycle transitions. Each point
744 represents the contribution of each plot and year for a given matrix element. Note that
745 the value for each plot was depicted in different symbols.

