

The dependence of halo bias on age, concentration, and spin

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Accepted 2019 May 2. Received 2019 March 26; in original form 2018 October 4

ABSTRACT

Halo bias is the main link between the matter distribution and dark matter haloes. In its simplest form, halo bias is determined by halo mass, but there are known additional dependencies on other halo properties which are of consequence for accurate modelling of galaxy clustering. Here, we present the most precise measurement of these *secondary-bias* dependencies on halo age, concentration, and spin, for a wide range of halo masses spanning from $10^{10.7}$ to $10^{14.7} h^{-1} M_{\odot}$. At the high-mass end, we find no strong evidence of assembly bias for masses above $M_{\text{vir}} \sim 10^{14} h^{-1} M_{\odot}$. Secondary bias exists, however, for halo concentration and spin, up to cluster-size haloes, in agreement with previous findings. For halo spin, we report, for the first time, two different regimes: above $M_{\text{vir}} \sim 10^{11.5} h^{-1} M_{\odot}$, haloes with larger values of spin have larger bias, at fixed mass, with the effect reaching almost a factor 2. This trend reverses below this characteristic mass. In addition to these results, we test, for the first time, the performance of a multitracer method for the determination of the relative bias between different subsets of haloes. We show that this method increases significantly the signal to noise of the secondary-bias measurement as compared to a traditional approach. This analysis serves as the basis for follow-up applications of our multitracer method to real data.

Key words: methods: numerical – galaxies: formation – galaxies: haloes – dark matter – large-scale structure of Universe – cosmology: theory.

1 INTRODUCTION

The clustering of galaxies is the prime observable that can be used to trace the large-scale structure of the Universe (LSS). In the standard model of cosmology, dark matter clusters along density peaks that were generated during inflation and collapse to form dark matter haloes. In this scenario, galaxies form when gas falls into collapsing dark-matter haloes (e.g. White & Frenk 1991). Hence, the relationship between galaxies, haloes, and the underlying matter distribution is crucial to our ability to test cosmological and galaxy formation models against observations.

The bias of dark matter haloes can be broadly defined as the relation between the distribution of haloes and the underlying matter density field. In its simplest description, the linear halo bias depends only on halo mass, with more massive haloes being more strongly clustered than less massive haloes (Kaiser 1984). However, halo bias is a much more complex effect that is known to depend on a variety of secondary halo properties. Among these dependencies on secondary properties, the most studied is the dependence on assembly history, called *halo assembly bias*. Low-mass haloes (M

$< M_*$, where M_* is the characteristic mass scale) that assemble a significant portion of their mass early on were shown to be more tightly clustered than haloes that assemble at later times, *at fixed halo mass* (see e.g. Gao, Springel & White 2005; Wechsler et al. 2006; Li, Mo & Gao 2008; Han et al. 2018; Salcedo et al. 2018). For high-mass haloes ($M > M_*$), however, the picture is less clear, with most studies showing a small or absent assembly bias signal at $M \sim 10^{14} h^{-1} M_{\odot}$ (Gao & White 2007; Salcedo et al. 2018). For cluster-size haloes of $M \sim 10^{15} h^{-1} M_{\odot}$, Chue, Dalal & White (2018) report a significant detection of halo assembly bias, of the opposite sign to the one observed at small halo masses (i.e. haloes are more tightly clustered than their older counterparts), with the inversion occurring precisely at $M \sim 10^{14} h^{-1} M_{\odot}$. These conclusions appear to be in contradiction with the results from Mao, Zentner & Wechsler (2018), who claimed to have found no halo assembly bias for very massive haloes. The analysis of halo assembly bias is extended to higher orders in Angulo, Baugh & Lacey (2008). Li et al. (2008) and Chue et al. (2018) further discussed the dependence of the halo assembly bias signal on the particular definition of halo age/formation time.

Halo bias has been shown to depend on a number of other halo properties, including spin, concentration, and shape. Hence, the

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term *halo assembly bias* has progressively been replaced by the more general *secondary bias*.¹ Concentration, in particular, has been extensively used as a proxy for formation epoch (see Wechsler et al. 2002). Although the behaviour observed for concentration is similar to that reported for age, there are some qualitative differences. At halo masses of $M \sim 10^{13} h^{-1} M_{\odot}$, a change of regime is well established: halo bias is larger for more concentrated haloes below this mass, but the trend reverses for higher masses (see e.g. Wechsler et al. 2006; Gao & White 2007; Han et al. 2018; Salcedo et al. 2018). The other ‘secondary’ property that has drawn significant attention in recent years is spin, λ , which is proportional to the angular momentum of the halo. At fixed mass, haloes with larger values of λ are found to be more tightly clustered than those having smaller values across the entire mass range considered, which typically covers $M \gtrsim 10^{12} h^{-1} M_{\odot}$ (e.g. Bett et al. 2007; Gao & White 2007; Faltenbacher & White 2010; Lacerna & Padilla 2012). The effect, however, appears to increase at the high-mass end (Salcedo et al. 2018).

Despite the variety of measurements, a comprehensive physical model of secondary halo bias is yet to be established. Dalal et al. (2008) proposed that certain features of halo assembly bias for high-mass haloes can be understood through the statistics of primordial density peaks. At low masses, the authors argue that assembly bias arises from a subpopulation of low-mass haloes whose mass accretion has ceased. For other attempts, we refer the reader to Zentner (2007) and Sandvik et al. (2007), which are based on the implementation of the ellipsoidal collapse model in the framework of the excursion set formalism.

Secondary halo biases have important consequences for the modelling of galaxy clustering. Measurements of the two-point correlation function from surveys like the Sloan Digital Sky Server (York et al. 2000) or the 2dF Galaxy Redshift Survey (Colless et al. 2001) indicate that more massive, more luminous, and redder galaxies are, in general, more tightly clustered than their less massive, less luminous, and bluer counterparts (e.g. Zehavi et al. 2005; Guo et al. 2014). The establishment of secondary bias invalidates any simple conclusion exclusively based on halo mass, consequently forcing halo-galaxy connection models to adjust (see Hearin et al. 2014, 2016; Tojeiro et al. 2017; McEwen & Weinberg 2018). In this context, it has become an observational challenge to prove the existence of the so-called *galaxy assembly bias*, i.e. the dependence of galaxy clustering on secondary halo properties such as the accretion history of haloes (see e.g. Lin et al. 2016; Miyatake et al. 2016; Guo et al. 2017; Montero-Dorta et al. 2017; Niemiec et al. 2018).

In this paper, we provide state-of-the-art measurements of secondary bias for age, spin, and concentration over four orders of magnitude in mass, in the virial mass range $10.7 \leq \log_{10} (M_{\text{vir}}/h^{-1} M_{\odot}) \leq 14.7$. This large dynamical range is achieved by combining four different MultiDark N -body numerical simulations.² In addition, we test for the first time the application of a full multitracer approach to the measurement of secondary bias. This technique is based on the fact that different tracers of LSS (e.g. distinct types of haloes) occupying the same cosmological volume reflect the same underlying density field. Multitracer techniques are designed

Table 1. Numerical properties of MultiDark simulations. The columns correspond to the name of the simulation, the length of the box’s side, the number of particles, the force resolution, and the mass of each simulated particle.

Name	L_{box} (h^{-1} Gpc)	N_{p}	ϵ (h^{-1} kpc)	M_{p} ($h^{-1} M_{\odot}$)
SMDPL	0.4	3840^3	1.5	9.6×10^7
MDPL2	1	3840^3	5	1.5×10^9
BigMDPL	2.5	3840^3	10	2.4×10^{10}
HugeMDPL	4.0	4096^3	25	7.9×10^{10}

to minimize the statistical uncertainties associated with cosmic variance by combining the information from distinct biased tracers of the LSS (McDonald & Seljak 2009; Seljak 2009; Abramo & Leonard 2013; Abramo, Secco & Loureiro 2016) – however, our method does not rely on knowledge of the density field, hence it can also be applied to real data.

This paper is organized as follows. Section 2 provides a brief description of the MultiDark simulations. The standard method used to measure the relative bias from the correlation function is presented in Section 3. Our main measurements of secondary bias are shown in Section 4. These measurements are compared with those obtained from a multitracer technique in Section 5, where this approach is also briefly described (an extended description of the method can be found in the Appendix). Finally, we compare our results with previous literature and summarize the main conclusions of our analysis in Section 6. Throughout this work, we assume the standard Λ CDM cosmology (Planck Collaboration XVI 2014), with parameters $h = 0.677$, $\Omega_{\text{m}} = 0.307$, $\Omega_{\Lambda} = 0.693$, $n_{\text{s}} = 0.96$, and $\sigma_8 = 0.823$.

2 SIMULATIONS

In order to study the clustering of haloes with different secondary properties we use the publicly available suite of MultiDark cosmological N -body simulations (Klypin et al. 2016). In this work, we analyse four different simulation boxes: Small MultiDark Planck (SMDPL), MultiDark Planck 2 (MDPL2), Big MultiDark Planck (BigMDPL), and Huge MultiDark Planck (HugeMDPL). These boxes have $\sim 4000^3$ particles and side lengths of 0.4, 1, 2.5, and 4 Gpc h^{-1} . A summary of the numerical parameters of each simulation is shown in Table 1.

Haloes were identified using the ROCKSTAR software (Behroozi, Wechsler & Wu 2013) and we only use the halo catalogue at redshift $z = 0$. Furthermore, only distinct haloes were included in this analysis. A halo is said to be distinct if its centre does not lie within a larger halo. The halo mass function and the velocity function for all distinct haloes in each MultiDark simulation are displayed in Fig. 1. We emphasize in Fig. 1 the large dynamical range addressed in this analysis. Fig. 1 also illustrates the halo mass and velocity incompleteness of each box (see Comparat et al. 2017 for further details).

In this study, we focus on the following halo properties:

(i) Virial mass, M_{vir} , computed in ROCKSTAR using the virial threshold of Bryan & Norman (1998), see Behroozi et al. (2013) for more details.

(ii) Maximum circular velocity, V_{max} , defined as

$$V_{\text{max}}^2 = \max \left[\frac{GM(< r)}{r} \right]. \quad (1)$$

¹Throughout the text, the term ‘assembly bias’ will exclusively refer to secondary bias produced by age, i.e. formation epoch. The effects produced by concentration and spin will be dubbed spin and concentration bias, respectively.

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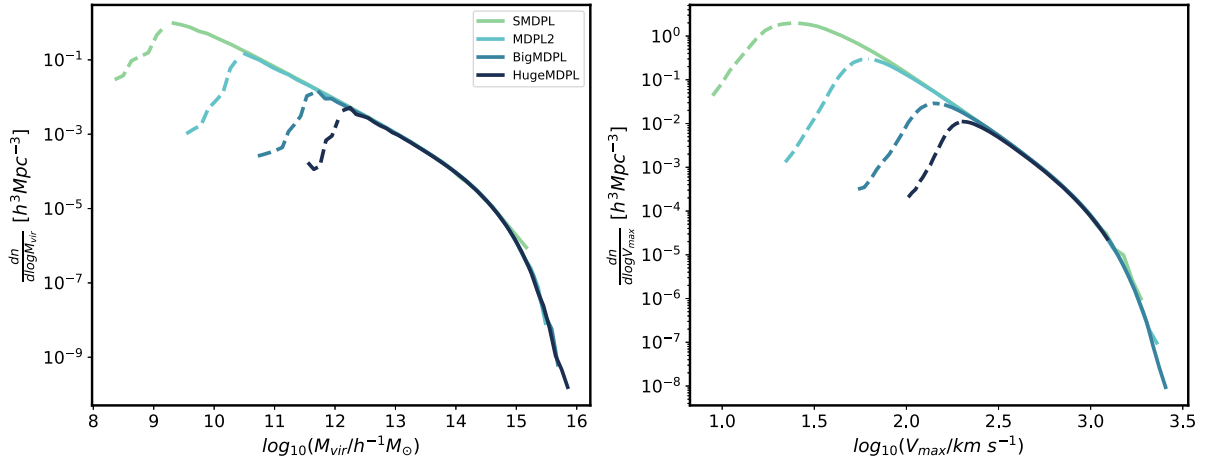


Figure 1. Halo mass function (left) and velocity function (right) for the four MultiDark simulation boxes at $z = 0$ employed in this work (see Table 1 for their numerical parameters). This figure illustrates the vast range of halo masses, velocities, and abundances analysed in this work. Dashed lines show the range of halo masses and velocities where completeness drops due to mass resolution effects. For the more conservative resolution limits employed in this work, see Table 2.

(iii) Age, $a_{1/2}$, defined as the scale factor at which half of the peak mass of the halo was accreted.

(iv) Spin, λ , defined as in Bullock et al. (2001), namely

$$\lambda = \frac{|J|}{\sqrt{2}M_{\text{vir}}V_{\text{vir}}R_{\text{vir}}}, \quad (2)$$

where J is the halo’s angular momentum and V_{vir} is its circular velocity at the virial radius R_{vir} .

(v) Concentration, c_{200} , defined as

$$c_{200} = \frac{R_{200}}{R_s}, \quad (3)$$

where R_{200} is computed from

$$M_{200} = \frac{4\pi}{3}200\rho_{\text{cr}}R_{200}^3, \quad (4)$$

and R_s is the Klypin scale radius (Klypin, Trujillo-Gomez & Primack 2011), which includes V_{max} and M_{vir} in its definition and assumes a NFW profile (Navarro, Frenk & White 1997).

We consider not only M_{vir} as the primary halo property, i.e. the main predictor of halo clustering, but also V_{max} , which characterizes the depth of the gravitational potential well. Note that V_{max} is defined unambiguously in N -body numerical simulations, in contrast to the virial mass, which depends on the particular density-contrast threshold adopted. In addition, adopting V_{max} has been shown to provide important advantages in the context of halo-galaxy connection frameworks such as halo abundance matching (Conroy, Wechsler & Kravtsov 2006; Trujillo-Gomez et al. 2011). To ensure a robust measurement of all relevant halo properties, we limit our analysis to haloes with more than 500 particles. An independent V_{max} cut was implemented, such that less than 1 per cent of haloes in each sample possesses less than 500 particles. For each box, this is equivalent to performing the mass and velocity cuts shown in Table 2.

To illustrate the general features of our data set, we show in Fig. 2 the distributions of the properties discussed above in the MDPL2 box. The age parameter, $a_{1/2}$, ranges from ~ 0.3 to ~ 0.5 , which correspond to redshifts of 2.3 and 1, respectively. Spins are typically in the range ~ 0.01 – 0.05 , and concentrations span values

Table 2. Cuts in virial mass performed in each simulation. These cuts were chosen so that only haloes with more than 500 particles are included in the analysis.

Name	Mass cut ($h^{-1} M_{\odot}$)	V_{max} cut (km s^{-1})
SMDPL	$>4.8 \times 10^{10}$	>80
MDPL2	$>7.5 \times 10^{11}$	>200
BigMDPL	$>1.2 \times 10^{13}$	>316
HugeMDPL	$>4.0 \times 10^{13}$	>631

of 5–15, approximately. As expected, very little correlation between secondary and primary halo properties is found.

3 RELATIVE BIAS MEASUREMENT

In this section, we provide a brief description of the standard procedure used to measure secondary bias from simulations, which is based on the computation of the 2-point correlation function.

To quantify the dependence of halo clustering on a secondary property S , we measure the relative bias between a subsample of haloes selected according to S and all haloes in the same primary bias property (B) range,

$$b^2(r, B, S) = \frac{\xi(r, B, S)}{\xi(r, B)}, \quad (5)$$

where we consider the primary bias parameters $B = M_{\text{vir}}$ and V_{max} , and the secondary parameters $S = a_{1/2}, \lambda, c_{200}$.

Each simulation box at $z = 0$ was divided in sub-boxes with $L_{\text{sub-box}} = L_{\text{box}}/4$. The resulting sub-catalogues were further divided in bins of width 0.15 in $\log_{10}(M_{\text{vir}})$ or 0.1 in $\log_{10}(V_{\text{max}})$. The haloes with the 25 per cent highest and lowest values of a particular secondary property were then selected in order to compute the expression shown in equation (5).

The 2-point correlation function was measured using CORFUNC (Sinha & Garrison 2017). In the top panels of Fig. 3, we show the correlation function for haloes selected according to age (left) and spin (right), in the $\log(M_{\text{vir}})$ ranges 12–12.15, and 13.35–13.5, respectively. Here, each point corresponds to the mean value of all sub-boxes and the error bars represent the standard deviation, computed from the entire set of sub-boxes. As found by

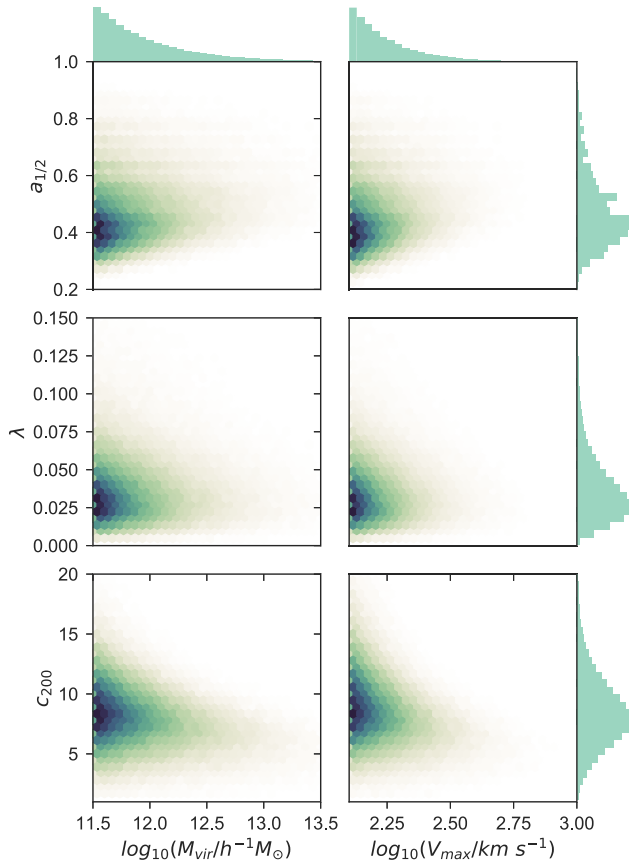


Figure 2. Joint histograms of M_{vir} , V_{max} , $a_{1/2}$, λ , and c_{200} and their marginal distributions. Each panel shows the correlation between a primary property and a secondary property in the MDPL2 simulation.

several previous studies, at fixed halo mass, haloes that assemble a significant portion of their mass at earlier times are more tightly clustered than those that assemble at later times. Additionally, haloes with a higher spin value are more strongly clustered than low-spin haloes. Note that the magnitude of these effects strongly depends on the particular halo mass range selected. In the lower panels of Fig. 3, the power spectrum for the same subsets of haloes as measured using the multitracer approach is also provided (see Section 5).

To fit the linear bias parameter we include both auto- and cross-correlations between different primary property bins, for distances ranging from 5 to 10 $\text{Mpc } h^{-1}$. We choose this range of scales due to the higher signal to noise in the assembly bias detection and to facilitate the comparison with previous literature. We then take the ratios of the correlation functions $\xi_{B,S}(r)/\xi_B(r) \rightarrow b_{B,S}^2/b_B^2$ and perform a minimum- χ^2 estimation for the relative bias.

4 THE EFFECT OF SECONDARY BIAS

The relative biases $b(B, S)$ for all combinations of primary properties $B = M_{\text{vir}}$, V_{max} and secondary properties $S = a_{1/2}$, λ , c_{200} are presented in Fig. 4. In each plot, the relative bias for haloes with the 25 per cent highest and lowest values of each secondary property, with respect to the entire population in the corresponding primary-property bin, is shown across all available MultiDark simulations. Error bars represent the standard deviation computed from all sub-boxes, where cross-correlations between primary-property bins

have been taken into account. Note that the error on the mean is, in most cases, smaller than the size of the markers.

The shaded region in each plot corresponds to the maximum finite width error, which can potentially be produced by the size of the mass bins. This effect could arise if a secondary property is correlated with halo mass. Thus, when selecting the top and bottom quartiles of such secondary property, one would also be selecting haloes based on mass. In order to quantify this effect, we select the 25 per cent highest and lowest mass values in each mass bin and compute the relative bias.

As shown in the two upper panels of Fig. 4, the assembly bias signal at the low-mass end ($M \sim 10^{10.7} h^{-1} M_{\odot}$) corresponds to an effect of ~ 45 per cent for old haloes and of ~ 30 per cent for young haloes. The assembly bias detection is consistent with zero at $M \sim 10^{14} h^{-1} M_{\odot}$ and beyond, in agreement with recent findings from Mao et al. (2018). The same behaviour is seen in the right-hand column, where the primary property is V_{max} and the effect vanishes at $\log_{10} V_{\text{max}} \gtrsim 2.7$.

In the third panel of Fig. 4, we present results on concentration bias, characterized by the secondary property c_{200} . In agreement with previous literature, a significant secondary bias signal is found for c_{200} in the same mass range (see e.g. Wechsler et al. 2006; Gao & White 2007; Jing, Suto & Mo 2007; Han et al. 2018; Salcedo et al. 2018). Here, an inversion occurs at masses $\sim 10^{13} h^{-1} M_{\odot}$. At the high-mass end, the difference in bias between the two quartiles is of a factor of 1.75. We note, however, that the inversion of the top and bottom quartiles when the haloes are binned by V_{max} for very massive haloes. In massive haloes, the bias decreases as V_{max} increases (Xu & Zheng 2018), which contributes to the trend seen in Fig. 3.

Another feature of interest can be seen in the secondary bias signal for spin, λ , shown in the second row of Fig. 4. By extending our analysis to very low halo masses, using the SMDPL box, we are able to detect, for the first time, an inversion similar to that found for concentration, with the top and bottom quartiles this time crossing over at masses of $\sim 10^{11.5} h^{-1} M_{\odot}$ (or, equivalently, $V_{\text{max}} \sim 10^{2.2} \text{ km s}^{-1}$). The large scale of the spin bias effect at the high-mass end, reaching a factor 2 at $\sim 10^{14.5} h^{-1} M_{\odot}$, is also noteworthy. It is interesting to note that the values of the spin parameter for the quartiles are remarkably similar across the entire mass range that we analysed, at around $\lambda \approx 0.015$ for the lowest quartile, and $\lambda \approx 0.08$ for the highest quartile.

Due to the resolution gap between SMDPL and MDPL2, we further investigated the inversion at the low-mass end using the BolshoiP simulation (Klypin et al. 2011) and found results consistent with those shown in Fig. 4. We additionally compared the different spin definitions in Peebles (1969) and Bullock et al. (2001) and found the results to be consistent with each other.

Fig. 4 confirms previous results regarding the asymmetric nature of secondary bias for age and concentration (see e.g. Salcedo et al. 2018 for a recent work). The effect appears significantly less pronounced for spin, for which it is only really noticeable at the very high mass end. A thorough discussion of our results in the context of previous literature is presented in Section 6.

5 THE MULTITRACER APPROACH

Multitracer techniques are designed to minimize the statistical uncertainties associated with cosmic variance by combining the information from distinct biased tracers of the LSS, including haloes (McDonald & Seljak 2009; Seljak 2009; Abramo & Leonard 2013;

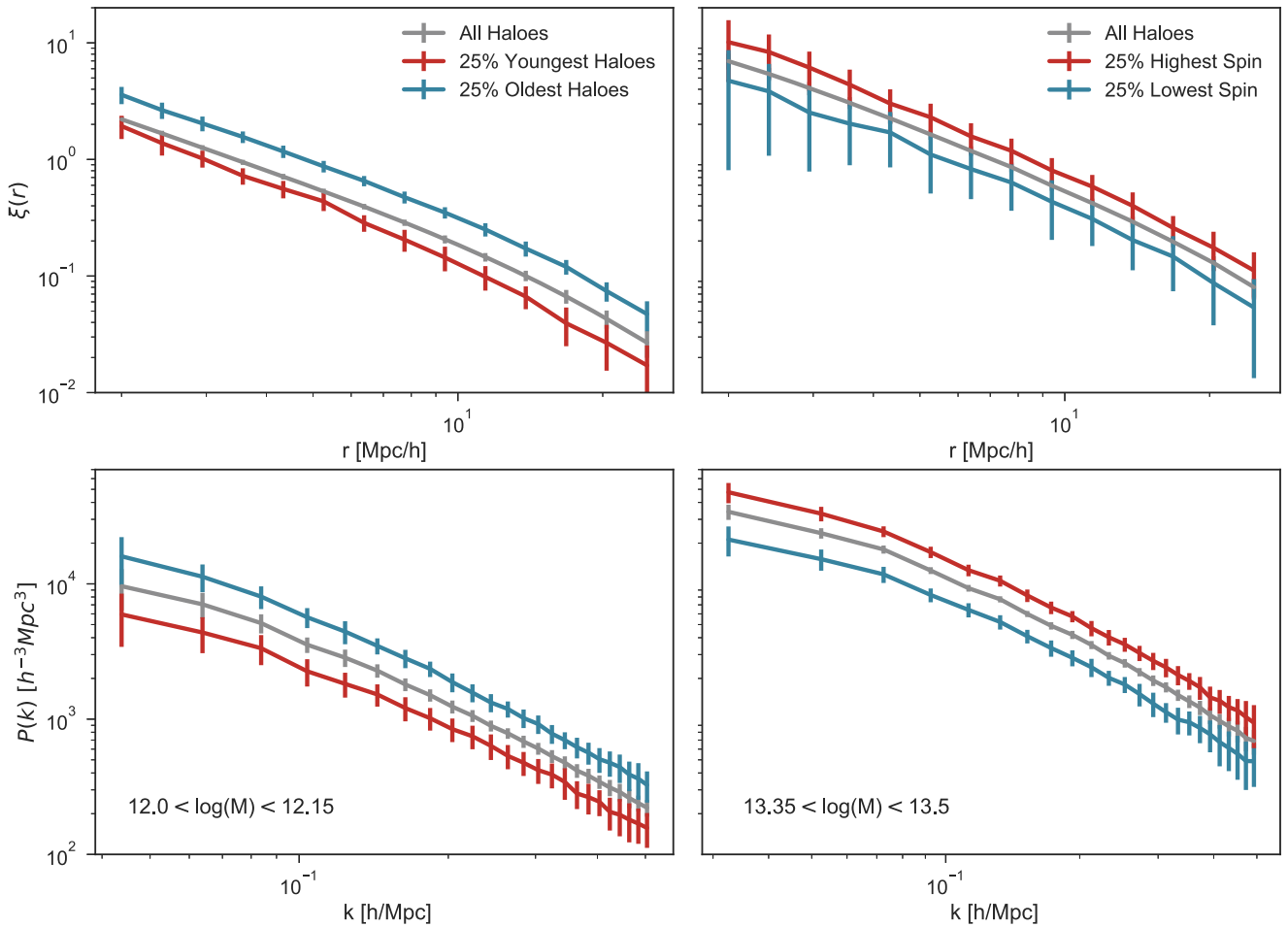


Figure 3. 2-point correlation function (upper panels) and power spectra (lower panels) illustrating the effect secondary bias. The left-hand panels show the effect of age bias for haloes of masses between 10^{12} and $10^{12.15} h^{-1} M_{\odot}$ in the MDPL2 simulation. The right-hand panels show the effect of spin bias for haloes of masses ranging between $10^{13.35}$ and $10^{13.5} h^{-1} M_{\odot}$ in the BigMDPL simulation.

Abramo et al. 2016). They have been employed in the analysis of real data (Blake et al. 2013; Ross et al. 2014; Marín et al. 2016), as well as in forecasts for future surveys, where they are expected to be especially beneficial due to the broad nature of their galaxy selections (Ferramacho et al. 2014; Alonso & Ferreira 2015; Fonseca et al. 2015; Abramo & Bertacca 2017; Witzemann et al. 2018). In this work, we have applied multitracer techniques to the measurement of secondary bias from the MultiDark simulations, and have regarded each subset of haloes, defined through primary and secondary properties, as a different LSS tracer. A brief description of the method can be found in the Appendix (for more information, see Abramo et al. 2016).

In order to compute the effect of the secondary-bias parameters using power spectra we use the same MultiDark boxes and sub-boxes that were described in Section 3, with the same primary bias (M_{vir} or V_{max}) bins and same criteria for splitting the samples into secondary-bias classes. We then compute the multitracer power spectra for a grand total of 32 different species of tracers – 8 bins in halo mass, times 4 bins in the secondary bias parameters.

The linear bias parameter is finally computed in terms of the ratios of spectra $P_{B,S}(k)/P_B(k)$. The variance of those ratios is determined from the sample of 27 sub-boxes in the case of the $400 h^{-1}$ Mpc box ($L_{\text{sub-box}} = L_{\text{box}}/3$), and 64 sub-boxes ($L_{\text{sub-box}} = L_{\text{box}}/4$) in the

cases of the 1, 2.5, and $4 h^{-1}$ Gpc boxes. We fit a constant (linear) bias to the ratios of spectra as a function of bandpower, using all available scales up to $k \leq \text{Min}(0.3 h \text{ Mpc}^{-1}, k_{\text{Shot}})$ where k_{Shot} is the scale at which the shot noise of the tracer species (B, S) dominates over the power spectrum of that tracer at the level of 90 per cent, i.e. $P_{B,S}(k_{\text{Shot}}) = 0.1 P_{B,S}^{\text{shot}} = 0.1 \bar{n}_{B,S}^{-1}$.

In the lower panels of Fig. 3, we show the power spectra computed using the multitracer method for haloes selected according to age (left) and spin (right), in the same halo mass bins used for the correlation function measurements presented in the top panels. As expected, the effects of assembly and spin bias are also significant in the power spectrum.

Fig. 5 shows the same secondary-bias measurements of Fig. 4, but determined through the multitracer method. We immediately notice that the multitracer method yields significantly smaller statistical errors across all MultiDark simulations when measuring secondary biases. Our derived uncertainties in the secondary bias range from $\lesssim 1$ per cent for the least massive (and most abundant) haloes of the SMDPL box, to ~ 5 –10 per cent for the massive haloes of the HugeMDPL box. When compared to the standard method, we find that the decrease in uncertainty depends on secondary property, quartile, and mass. For the MDPL2 simulation, as a reference, we report a decrease of roughly 25–40 per cent in the uncertainty for

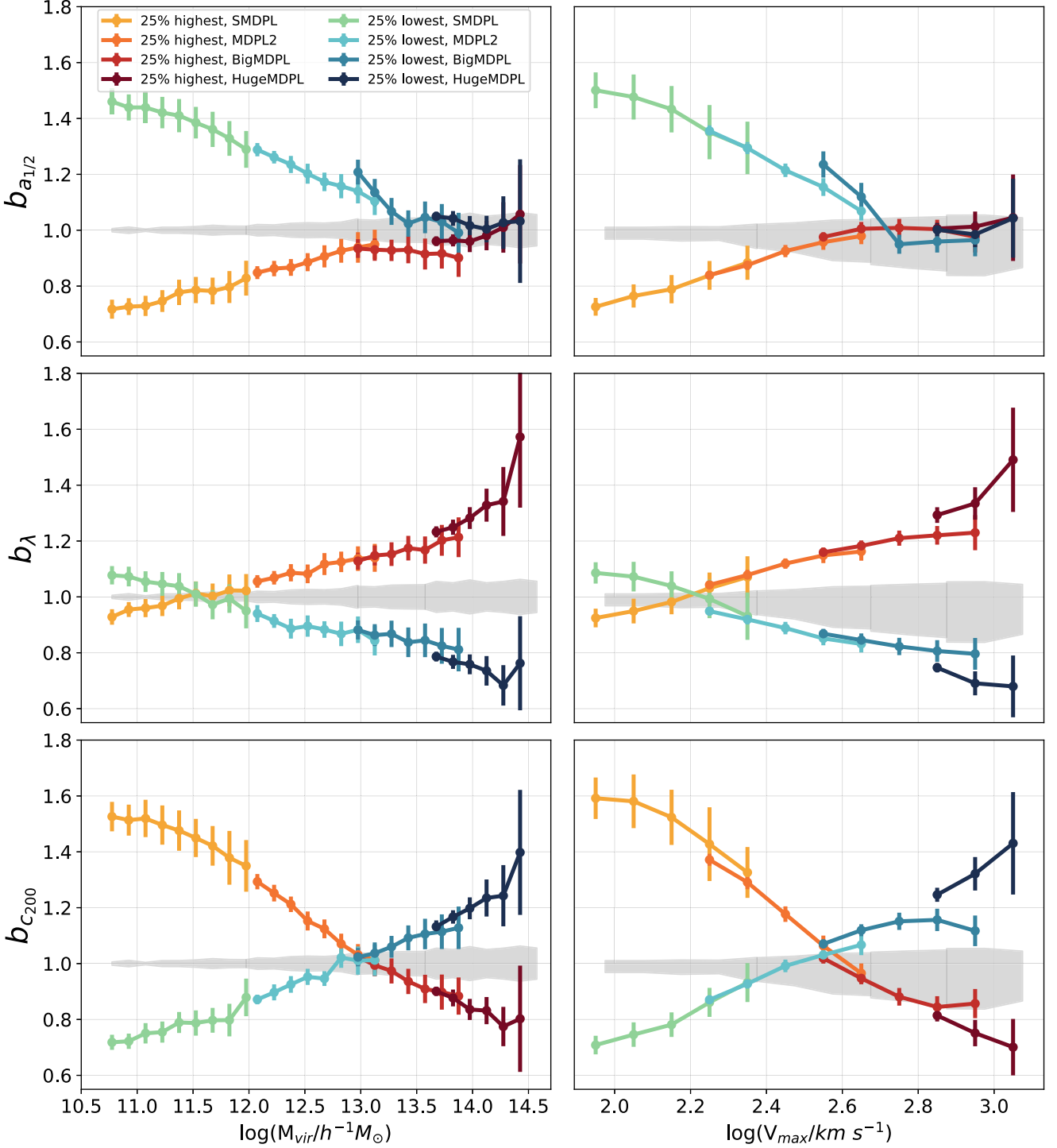


Figure 4. Secondary biases $b(S, P)$ as defined in equation (5), estimated using the standard correlation-function method, for primary properties M_{vir} (left-hand column) and V_{max} (right-hand column), and secondary properties $a_{1/2}$, λ , and c_{200} . Each point corresponds to the average among all sub-boxes of a particular simulation and the errors are the standard deviation computed from the entire set of sub-boxes. Note that cross-correlations between primary-property bins have been taken into account in the determination of error bars. All plots show the biases of top and bottom quartiles of a particular secondary property, for all available MultiDark simulations (see the text).

age as the secondary property. For spin the decrease is typically around 10–35 per cent and for concentration we find a decrease of up to 30 per cent.

We must emphasize that the results presented in Figs 4 and 5 are not directly comparable, since they probe different distance scales.

In particular, the bias estimated through the correlation function is measured within a narrow distance range of $5\text{--}10 h^{-1}$ Mpc, whereas the bias in terms of the power spectrum is measured on scales $k \sim 0.05\text{--}0.3 h \text{Mpc}^{-1}$, depending on the box size and level of shot noise. In particular, this means that, if there is a significant

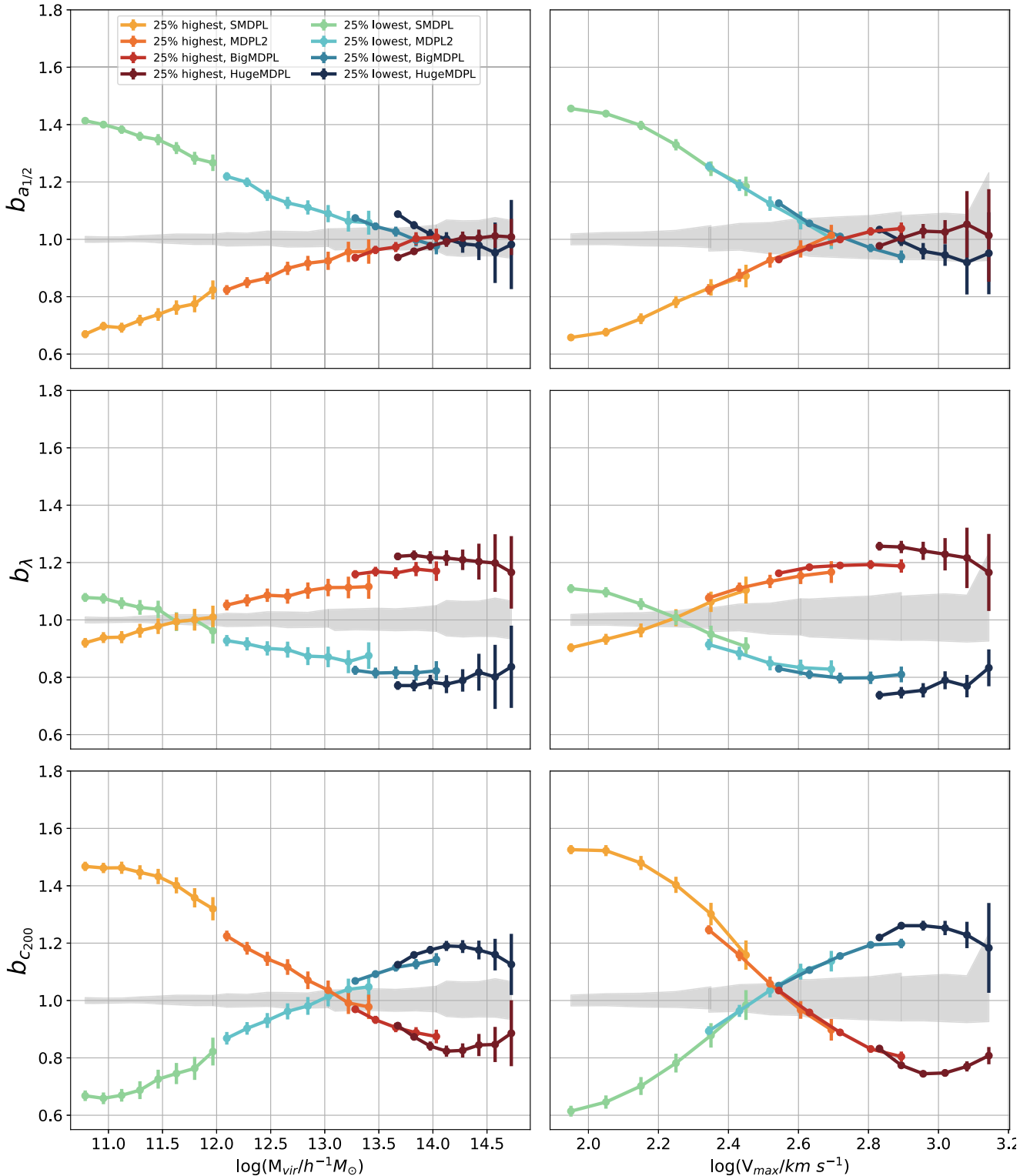


Figure 5. Secondary bias measured through the multitracer method, as described in Section 5. Similarly to Fig. 4, the left-hand column corresponds to the primary property M_{vir} , the right-hand column to the primary property V_{max} , and each row corresponds to the secondary properties $a_{1/2}$, λ , and c_{200} . Once again, each point show the average among all sub-boxes and the errors are the standard deviation from the entire set of sub-boxes.

scale dependence of the secondary bias, the measurements using correlation function and power spectra could differ. In that case our results using the power spectra may differ from tracer to tracer, and from box to box – on that regard, see the Appendix.

With the aforementioned caveats, both methods provide qualitatively similar results for assembly, concentration and spin bias. Some small differences are, however, worth noting. In particular, the multitracer measurements appear flatter at the high-mass end,

especially for HugeMDPL, as compared to the standard measurements. In addition, the top-right panel of Fig. 5 seems to show hints of a cross-over between young and old haloes for high- V_{\max} haloes. Further investigation will be required in order to clarify the origin of these differences.

6 DISCUSSION AND CONCLUSIONS

By combining all available MultiDark simulations, we provide the most precise measurement of secondary halo bias across a wide virial mass range, spanning from $10^{10.7}$ to $10^{14.7} h^{-1} M_{\odot}$. Furthermore, we present the first multitracer measurement of secondary bias across the same comprehensive dynamical range.

For masses below $10^{14} h^{-1} M_{\odot}$ we find, in agreement with, e.g. Gao et al. (2005), Wechsler et al. (2006), Li et al. (2008), Salcedo et al. (2018), Mao et al. (2018), and Chue et al. (2018), that older haloes are more strongly clustered than younger haloes, for the definition of age given in Section 2. Above $10^{14} h^{-1} M_{\odot}$ we detect no solid evidence of halo assembly bias or of a cross-over between the top and bottom age quartiles, in agreement with Mao et al. (2018), who also used the MDPL2 box. Interestingly, some hints of a cross-over are found when V_{\max} is used as the primary halo property and the simulation data is analysed using the multitracer technique. This is not observed when a traditional approach is employed, nor when halo mass is used as the primary halo property in the multitracer measurement. Follow-up work will be devoted to clarify the origin of this feature in V_{\max} . We highlight that, using the BigMDPL simulation, Chue et al. (2018) reports a significant assembly bias signal at $10^{15} h^{-1} M_{\odot}$. Through the methodology adopted in this paper, however, we were unable to obtain statistically significant measurements in the same mass range.

We present a novel feature in the secondary bias produced by spin: a cross-over between top and bottom quartiles at $M \sim 10^{11.5} h^{-1} M_{\odot}$ or, equivalently, at $V_{\max} \sim 10^{2.2} \text{ km s}^{-1}$. We are able to probe such small halo masses due to the high-mass resolution of the SMDPL simulation. Although the presence of a cross-over in age is still uncertain, the detection of this phenomenon in spin, in combination with previous detections for concentration and several other properties, might be an indication that the inversion of the secondary bias signal is a universal property. In this context, Salcedo et al. (2018) argues that the behaviour of spin bias is different from that of other secondary biases, which may suggest a distinct physical origin. Some of the evidence we present here, however, might cast serious doubt on such claim. For example, Salcedo et al. (2018) states that spin bias differs from other secondary biases due to its weak mass dependence and the fact that it increases with mass, whereas most other secondary biases decrease with mass. However, in the wider mass range probed in this work, we find that the relative bias for spin varies as much as for age (~ 0.5). Moreover, the detection of a cross-over places spin bias in the same footing as other secondary biases, since it shows that the fact that it increases with mass was an artefact of the mass scales that were being probed.

Our results place updated constraints to models addressing the connection between galaxies and haloes. The wide dynamical range that we probe encompasses haloes hosting from emission-line galaxies at the low-mass end to massive quiescent galaxies and clusters at the high-mass end. In the context of the halo-galaxy connection, it has become an observational challenge to prove the existence of *galaxy assembly bias*, a hypothesis that states that the clustering properties of galaxies, at fixed halo mass, depend on secondary halo properties such as the accretion history (or age) of haloes (see e.g. Lin et al. 2016; Miyatake et al. 2016;

Montero-Dorta et al. 2017; Niemiec et al. 2018). For luminous red galaxies (LRGs), indications of the existence of galaxy assembly bias have been reported by Montero-Dorta et al. (2017) and Niemiec et al. (2018), from a sample extracted from the Baryon Oscillation Spectroscopic Survey (BOSS; Dawson et al. 2013). The typical mass of the haloes inhabited by these galaxies is estimated in $\sim 10^{12.7-13} h^{-1} M_{\odot}$ (Niemiec et al. 2018), but a large scatter is expected, so that the most massive LRGs could inhabit haloes above $\sim 10^{14} h^{-1} M_{\odot}$. These results are, in principle, compatible with both the halo assembly bias and the concentration bias signal reported in Fig. 4. Our measurement of halo assembly bias at the high-mass end can be properly connected to the results presented in Montero-Dorta et al. (2017) and Niemiec et al. (2018) through the technique of ‘age matching’ (Hearin et al. 2016).

In addition, the large scale of the spin bias signal for cluster-size haloes, of almost a factor 2 at $\sim 10^{14.5} h^{-1} M_{\odot}$, suggests an alternative route towards an observational proof of secondary bias. Probing the effect of spin bias requires measuring rotation, or a proxy for it, in a large sample of galaxy clusters. These measurements, which are still extremely challenging, might be possible in the near future thanks to techniques such as kinetic Sunyaev–Zel’dovich effect (see e.g. Baldi et al. 2018).

We have applied, for the first time, a fully multitracer approach to the measurement of secondary bias. Here, different subsets of haloes are viewed as different LSS tracers. We show that this method is capable of reproducing all features observed from the standard measurement. In addition, we have obtained statistical uncertainties which are comparable to other bias estimates from N -body simulations that minimize cosmic variance by relying on direct knowledge of the density field, e.g. Gao & White (2007); Chue et al. (2018). However, in contrast to those methods, which in effect compute δ_a/δ_m at each point (in configuration or Fourier space), the technique used in this work can be applied both to simulations as well as to actual data, in real or redshift space.

The multitracer approach is a promising technique for future surveys like *Euclid*³ or *DESI*⁴, where cosmological measurements will be performed from data sets containing multiple galaxy populations. The effort presented in this paper will serve as the basis for future developments aimed at the application of the method to real data.

The main conclusions of this work can be summarized as follows:

- (i) No statistically significant halo assembly bias signal (secondary bias on $a_{1/2}$) is detected for haloes above $M \sim 10^{14} h^{-1} M_{\odot}$.
- (ii) A cross-over is detected, for the first time, in the spin bias signal. Below $M \sim 10^{11.5} h^{-1} M_{\odot}$, lower spin haloes are more tightly cluster than higher spin haloes, at fixed halo mass. The signal reverses above this characteristic mass. The effect of spin bias increases significantly at the high-mass end, reaching almost a factor of 2 at $M \sim 10^{14.5} h^{-1} M_{\odot}$.
- (iii) We test, for the first time, the performance of a fully multitracer approach for the measurement of secondary bias. These techniques are designed to minimize the statistical uncertainties associated with cosmic variance by combining information from distinct biased tracers of the LSS, in this case, haloes.
- (iv) The multitracer approach is capable of reproducing all secondary-bias features observed from the standard measurement, with the advantage that the signal to noise improves significantly. We find that the decrease in uncertainty depends on secondary property,

³<https://www.euclid-ec.org>

⁴<https://www.desi.lbl.gov>

quartile, and mass. For the MDPL2 simulation, as a reference, we report a decrease of roughly 25–40 per cent in the uncertainty for age as the secondary property. For spin the decrease is typically around 10–35 per cent and for concentration we find a decrease of up to 30 per cent.

(v) Our halo assembly bias measurements are consistent with results suggesting that galaxy assembly bias can be detected from massive galaxies alone, and, in particular, from LRGs.

ACKNOWLEDGEMENTS

GSP and ADMD thank FAPESP for financial support. LRA thank both FAPESP and CNPq for financial support. FP acknowledges support from the Spanish MINECO grant AYA2010-21231-CO2-01.

We thank New Mexico State University (USA) and Instituto de Astrofísica de Andalucía CSIC (Spain) for hosting the Skies & Universes site for cosmological simulation products.

We also thank Peter Behroozi for running Rockstar Consistent Tree on the MultiDark simulations.

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APPENDIX A: THE MULTITRACER TECHNIQUE

The optimal combination of the density fields of a number of biased tracers, which results in minimum-variance estimators for the power spectra of those tracers, was derived in Abramo et al. (2016). The technique consists in a generalization of the weights obtained by Feldman, Kaiser & Peacock (1994) for a single tracer – the so-called FKP weighting. The multitracer weighted density field of a tracer species α is defined as $f_\alpha(\vec{x}) = \sum_\beta w_{\alpha\beta}(\vec{x})\delta_\beta(\vec{x})$, where the weights are defined by

$$w_{\alpha\beta} = \bar{n}_\alpha b_\alpha \delta_{\alpha\beta} - \frac{\bar{n}_\alpha b_\alpha^2 P_m}{1 + \mathcal{P}_{\text{eff}}} \bar{n}_\beta b_\beta. \quad (\text{A1})$$

In the expression above, $\bar{n}_\alpha(\vec{x})$ is the measured number density of the tracer α , b_α is its fiducial bias, $P_m = P_m(k_p)$ is the fiducial matter power spectrum at the pivot scale $k_p = 0.1 h \text{ Mpc}^{-1}$, and $\mathcal{P}_{\text{eff}} = \sum_i \bar{n}_i b_i^2 P_m$. We stress the fact that the multitracer weights are not equivalent to the FKP weights: in particular, notice that the weighted fields f_α are admixtures of the density fields of all tracers. Only when there is a single tracer the multitracer weight reduces to the FKP weight.

After computing the Fourier transform of the weighted fields we take their (quadratic) amplitudes and combine them into bandpowers to get the spectra of the weighted fields. In the FKP method, the estimator of the power spectrum for the tracers is basically the power spectra of the weighted fields, with an appropriate normalization and its respective shot noise subtraction. In the multitracer method,

the minimum-variance estimators for the autopower spectra of the individual tracers are computed in terms of a linear combination of the spectra of the weighted fields, with the subtraction of a shot noise term – see Abramo et al. (2016) for details.

Since we have 8 mass bins, each one having 4 bins corresponding to the secondary bias parameters, there are 32 individual tracers. We compute the power spectra of each individual tracer, and average the spectra of tracers in the same mass bin in order to obtain the power spectra for the entire mass bin (we checked that this procedure yields the same result as computing the power spectra for the combination of tracers in the same mass bins directly). Once we have the power spectra of the individual tracers $[P_{B,S}(k)]$ as well as the spectra corresponding to the tracers in the corresponding mass bin $[P_B(k)]$, we determine the ratios of halo biases through the ratios of their spectra, $P_{B,S}(k)/P_B(k) \rightarrow b_{B,S}^2(k)/b_B^2(k)$.

A key aspect of multitracer optimal weighting is the fact that the weighted fields f_α are linear combinations of the density fields of each tracer, hence the multitracer power spectrum estimators automatically include all the auto- and cross-spectra of all tracers, in a minimum-variance combination. In particular, this means that the information from cross-spectra has already been taken into account by the multitracer autospectra estimators, so in principle we do not

need to compute them separately. This is not the case for the FKP power spectra, where the information in the cross-spectra needs to be added separately.

In Fig. A1, we show the power spectrum (left-hand panel) for a particular age quartile (youngest haloes) in the mass bin $\log_{10} M[h^{-1} M_\odot] = 13.825 \pm 0.075$, extracted from the 4 Gpc HMDPL box, as well as the ratio of this power spectra to the power spectrum of the entire mass bin (right-hand panel). The dashed lines correspond to the power spectra computed using the multitracer optimal weighting, whereas the solid lines show the traditional (FKP) weighting. It is clear that the variances of the amplitudes of the power spectra, as well as the statistical fluctuations of the ratios of spectra (the relative bias), are smaller when we employ the multitracer weighting. In the case of this mass bin and box, the spectra are dominated by shot noise, so the multitracer technique, which uses the information from the cross-correlations, is naturally a less noisy estimator of the power spectrum. When shot noise is less relevant (as happens for low-mass bins), the statistical fluctuations of the power spectra estimated with the two methods can be similar, but the covariance of the ratios of the power spectra are typically much smaller when we employ the multitracer method (Abramo et al. 2016).

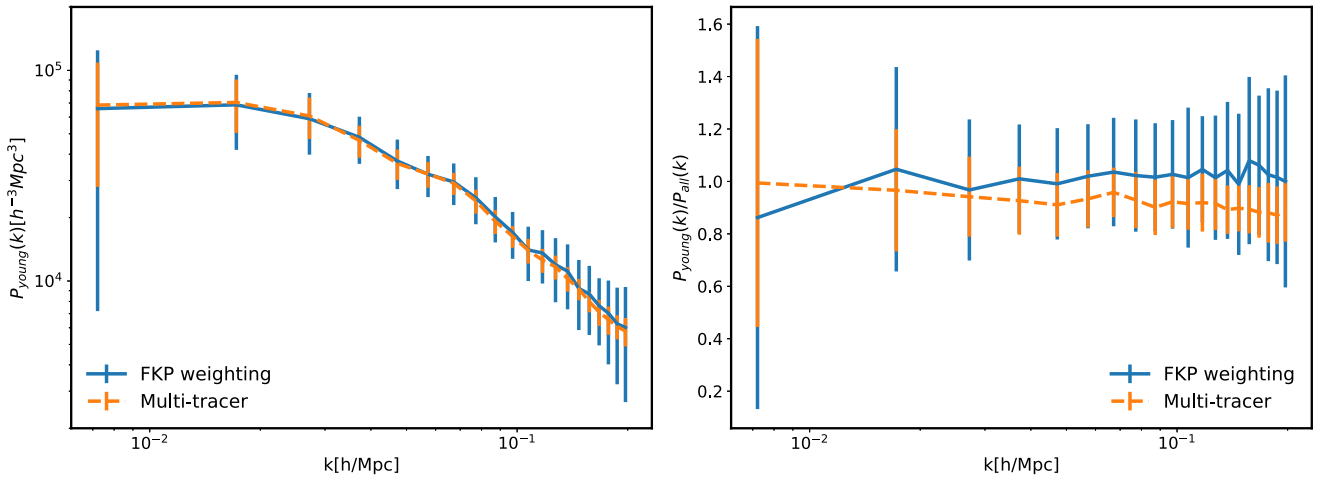


Figure A1. Left: Power spectrum in the entire wavenumber range considered for a subset containing the youngest haloes in the virial mass bin $\log_{10} M[h^{-1} M_\odot] = 13.825 \pm 0.075$, extracted from the 4 Gpc HMDPL box. Dashed lines correspond to the power spectra computed using the multitracer optimal weighting, whereas solid lines show the traditional (FKP) weighting. Right: Same as before for the ratio of the power spectrum of the youngest haloes and that of the entire population in the corresponding bin.

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